

## Motivation

- Cholesky multivariate stochastic volatility (CMSV) model commonly used to specify dynamic covariance matrices of a  $n$ -dimensional vector  $y_t$ :

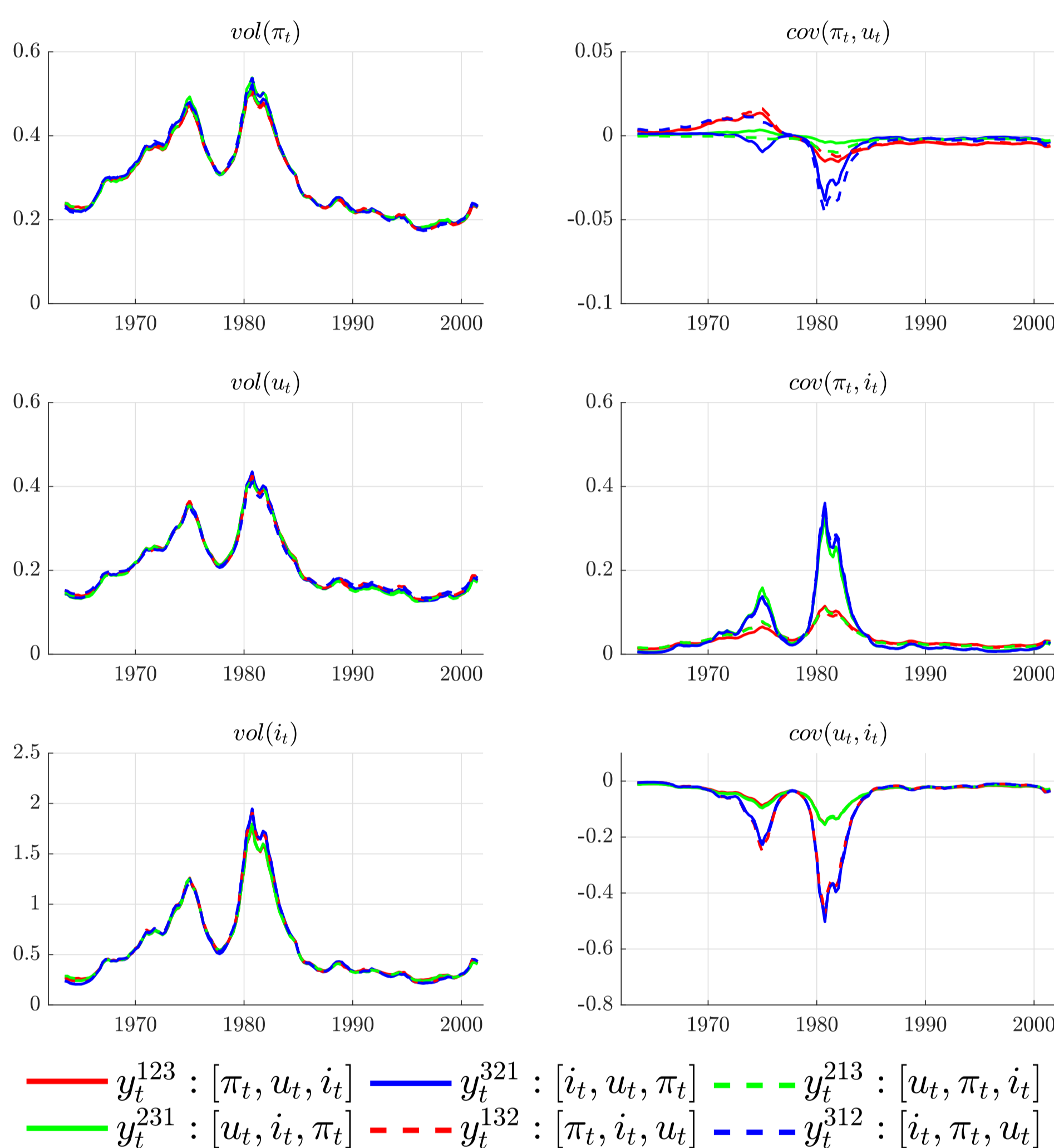
$$y_t \sim N(0, \Sigma_t), \quad \text{assume } \Sigma_t = A_t^{-1} D_t D_t' A_t^{-1}$$

$$\Rightarrow y_t = A_t^{-1} D_t \epsilon_t, \quad \epsilon_t \sim N(0, I_n)$$

i.e.  $\Sigma_t$  implicitly modelled by specifying  $A_t^{-1}$  and  $D_t$

- But:** estimates of  $\Sigma_t$  may be sensitive to the ordering variables in  $y_t$ , see e.g., Primiceri (2005)  
 $\Rightarrow$  inference may hinge on a chosen variable ordering  
 $\Rightarrow$  majority of applied literature ignores this property

- Illustration:** Effect of alternative variable orderings on dynamics of  $\Sigma_t$  in Primiceri's (2005) application  
 $\Rightarrow$  volatilities similar, covariances differ in stagflation



## Research questions

- Role of variable ordering on the dynamics of  $\Sigma_t$ ?
- Variable ordering important for conclusions?
- How to mitigate the ordering sensitivity?

## Contributions and Results

- Ordering sensitivity not negligible in CMSV model!  
 $\Rightarrow$  volatility pattern impose alternative restrictions
- Propose a robust modelling alternative  
 $\Rightarrow$  dynamic correlation Cholesky MSV (DC-CMSV)
- Monte Carlo simulation to fit  $\Sigma_t$   
 $\Rightarrow$  Estimated correlations almost ordering insensitive when there is no volatility (CMSV & DC-CMSV)  
 $\Rightarrow$  Estimated covariances of CMSV model more distinct when there are stronger idiosyncratic volatility clusters, while covariances hardly affected by alternative volatility pattern under DC-CMSV
- Inference may hinge on a ordering for estimating  $\Sigma_t$ !  
(1) US monetary policy during stagflation  
 $\Rightarrow$  unchanged or more aggressive response?  
(2) Predictability of US inflation-gap  
 $\Rightarrow$  gradual or abrupt improvement in predictability?

## On the Cholesky MSV model

Let  $y_t$  be a 2-dimensional vector (tractability)

How does the CMSV structure affect dynamics of  $\Sigma_t$ ?

- CMSV model:  $\Sigma_t = A_t^{-1} D_t D_t' A_t^{-1}$   
-  $a_t$  off-diag. of  $A_t$  and  $g_t$  log-vol. process of  $D_t$   
-  $a_t$  and  $g_t$  are Gaussian random walk (RW)

### Properties of $\Sigma_t$ under CMSV

- the ratio of volatilities  $\frac{\sigma_{22,t}}{\sigma_{11,t}}$  is time-varying

- the correlation  $\rho_t$  evolves nonlinearly

$$\rho_t = \rho_{t-1} \frac{\exp(\eta_{1,t}^g)}{\exp(\eta_{2,t}^{g**})} + \epsilon_t^a \frac{\sigma_{11,t}}{\sigma_{22,t}}$$

- the contemporaneous relation  $a_t$  evolves linearly
- the dynamic structure of  $\Sigma_t$  cannot be generated by an analogously setup CMSV model for  $\tilde{y}_t$
- dynamic restrictions increase in the variability of idiosyncratic volatility patterns

Comparison to separated volatilities and correlations?

- DC-MSV model:  $\Sigma_t = D_t R_t D_t'$  (Yu and Meyer, 2006)  
-  $h_t$  log-vol. process of  $D_t$  and  $\rho_t$  correlation of  $R_t$   
-  $\rho_t(m_t) = \frac{\exp(m_t) - 1}{\exp(m_t) + 1}$ ,  $m_t$  and  $h_t$  are Gaussian RW  
- applicable only to  $n \leq 3$  (psd of  $R_t$  not guaranteed)

### Properties of $\Sigma_t$ under DC-MSV

- the ratio of volatilities time-varying or constant
- the correlation  $\rho_t$  evolves approximately linearly
- the contemp. relation  $a_t$  evolves nonlinearly

$$a_t = a_{t-1} \frac{\exp(\eta_{2,t}^h)}{\exp(\eta_{1,t}^h)} + \eta_t^a \frac{\exp(h_{2,t})}{\exp(h_{1,t})}$$

Fit  $\Sigma_t$  with CMSV, when  $y_t$  generated by DC-MSV?

- nonlinear transformation of  $a_t$  as volatilities switch position ( $a_t = \rho_t \frac{\exp(h_{2,t})}{\exp(h_{1,t})}$ ,  $\tilde{a}_t = \rho_t \frac{\exp(h_{1,t})}{\exp(h_{2,t})}$ )
- systematically different paths of the covariance ( $a_t$  underestimated in one ordering, while mechanically overestimated in reverse ordering)

Special cases:

- $\rho_t = \rho, \forall t$ : effect more severe (no offsetting by  $\eta_t^a$ )
- $h_t = h, \forall t$ :  $a_t$  is almost ordering insensitive

## The DC-Cholesky MSV model

- Let  $y_t$  be a  $n$ -dimensional vector with  $y_t \sim N(0, \Sigma_t)$

- DC-CMSV model:  $\Sigma_t = D_t R_t D_t'$

$$\Rightarrow y_t = D_t \epsilon_t, \quad \epsilon_t \sim N(0, R_t)$$

$\Rightarrow$  estimate auxiliary matrix  $Q_t = A_t^{*-1} D_t^* D_t'^* A_t^{*-1}$

using the CMSV model on stand. data  $\epsilon_t = D_t^{-1} y_t$   
 $\Rightarrow$  estimate correlations via Engle's (2002) formulas

$$R_t = Q_t^{*-1/2} Q_t Q_t^{*-1/2}, \quad Q_t^* = \text{diag}[\text{vecd}(Q_t)]$$

where  $\text{vecd}(Q_t)$  selects the diagonal of  $Q_t$ .

### Further assumptions

- State dynamics: RW, stationary or combination
- Independent innovations of volatility and correlation

## Monte Carlo Simulation

DGP: Correlations from Engle (2002) w/o SV

(1) Fitting correlations with CMSV model (wo SV)

	$\rho_t$	$a_t - \tilde{a}_t$	$a_t$	$\tilde{a}_t$
const	<b>0.008</b>	0.084	0.018	0.019
sine	<b>0.022</b>	0.086	0.035	0.034
fsine	<b>0.016</b>	0.070	0.020	0.020
step	<b>0.010</b>	0.076	0.018	0.018
ramp	<b>0.023</b>	0.087	0.037	0.037

Table: Mean absolute distance (MAD) without stochastic volatility

- MAD lowest for  $\rho_t$  ( $R_t = \Sigma_t^{*-1/2} \Sigma_t \Sigma_t^{*-1/2}$ )

$\Rightarrow$  implied  $\rho_t$  almost ordering insensitive

(2) Fitting covariances (with SV)

	High Vol		Low Vol	
	CMSV	DC-CMSV	CMSV	DC-CMSV
const	0.207	<b>0.037</b>	0.153	<b>0.026</b>
sine	0.179	<b>0.021</b>	0.081	<b>0.023</b>
fsine	0.210	<b>0.012</b>	0.049	<b>0.015</b>
step	0.169	<b>0.021</b>	0.089	<b>0.019</b>
ramp	0.183	<b>0.023</b>	0.085	<b>0.025</b>

Table: Mean absolute distance (MAD) with stochastic volatility

- CMSV: MAD of  $\sigma_{12,t}$  increases for high vol. DGP

$\Rightarrow$  DC-CMSV: almost insensitive to alt. DGPs

## Empirical Application

(1) Evolution of US monetary policy (Primiceri, 2005)

- unchanged or more aggressive response?

$\Rightarrow$  ambiguous with CMSV model

$\Rightarrow$  DC-CMSV model suggest that the Fed counteracted  $\pi$  and  $UR$  more aggressively!

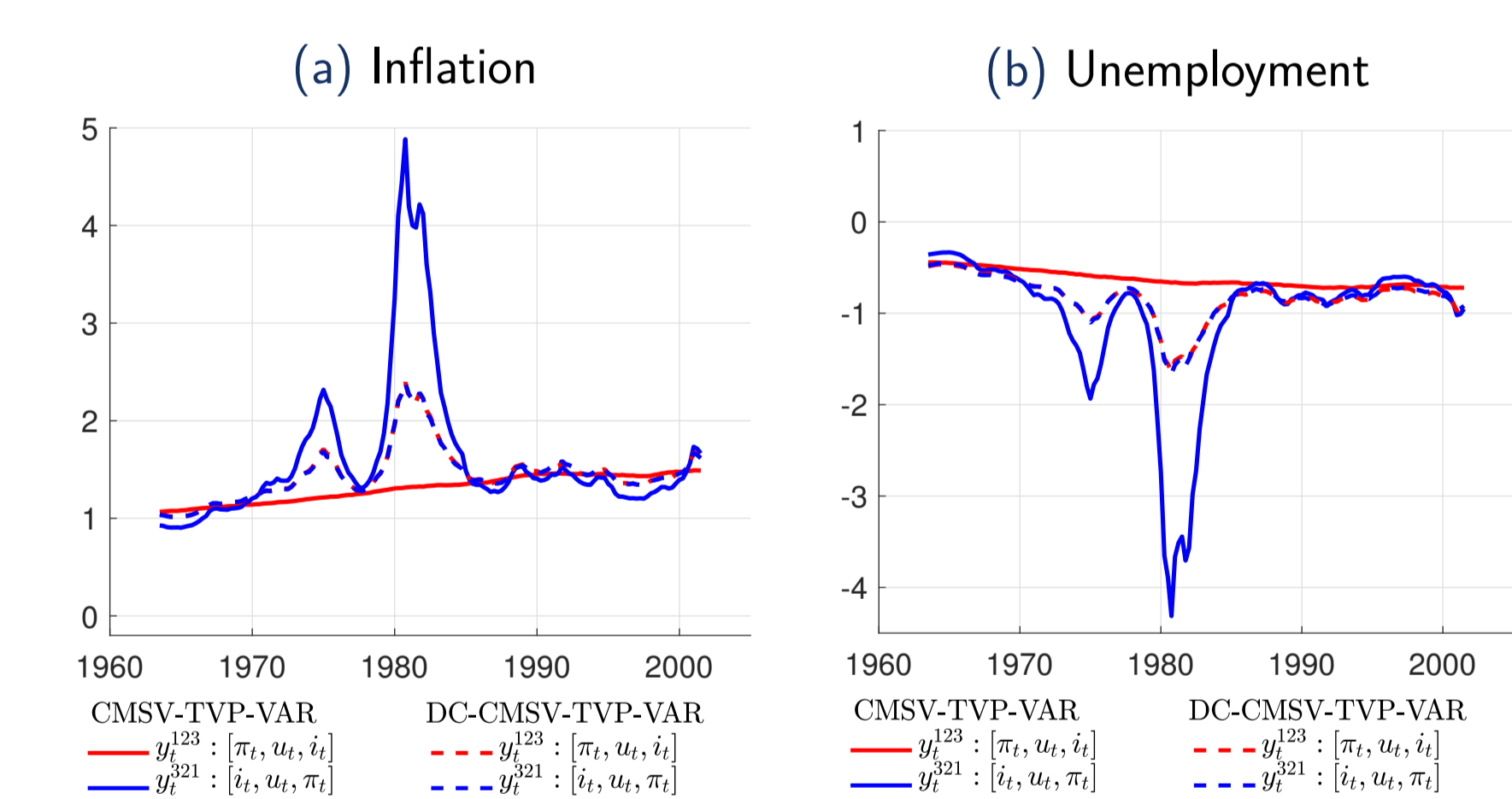


Figure: Estimated long-run US systematic interest rate response

(2) Properties of US inflation-gap persistence

(Cogley, Primiceri, and Sargent, 2010)

- decline after great inflation or unchanged?

$\Rightarrow$  ambiguous with CPS-TVPSV-VAR model; driven by CMSV heteroskedasticity in TVP

- without CMSV in TVP, unambiguous conclusion!

$\Rightarrow$  persistence declined after 1980s

## Conclusion

- Variable ordering in CMSV model important!  
 $\Rightarrow$  volatility pattern imposes restrictions  
 $\Rightarrow$  ambiguous conclusions in applications  
 $\Rightarrow$  idiosyncratic volatility pattern not uncommon
- DC-CMSV model as robust alternative  
 $\Rightarrow$  estimates almost ordering invariant  
 $\Rightarrow$  nonlinear contemporaneous relations