Online Appendix

 \mathbf{to}

"Can you move to opportunity? Evidence from the Great Migration"

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Appendix A Historical and contemporary measures of upward mobility

This appendix describes the construction of the historical measures of upward mobility used in the analysis as well as details on contemporary measures from Chetty and Hendren (2018c) and Chetty et al. (2020b).

Educational upward mobility in 1940 I follow Card et al. (2018) and define educational upward mobility as the fraction of 14-18 year-old boys and 14-16 year-old girls in each commuting zone with 9 or more years of schooling from households where the most educated parent has between 5 and 8 years of schooling, or the median for adults in the US at the time.³²

Alternative historical measures of upward mobility Prior to 1940, the Census did not record years of schooling attained. For the years 1900-1940, I measure educational upward mobility as the school attendance rates of teenagers with low income fathers using the Integrated Public Use Microdata Series (IPUMS) version of the complete count US censuses (Ruggles et al., 2021). For 1900, I do so by creating a binary indicator for having attended school defined as reporting a positive number of months of school attendance during the year ending June 1, 1900, the Census day that year. In 1910, Census enumerators asked whether respondents of school age had attended school between September 1, 1909 and April 15, 1910 (the Census day in 1910), a period of 9 months. In 1920, by contrast, the analogous question asked about school attendance between September 1, 1919 and January 1, 1920 (the Census day that year), a period of 4 months. Mean school attendance for all teens (reported at the bottom of Table 3) is higher for 1910 (75%) than for 1920 (65%), likely due to the above discrepancy.

 $^{^{32}}$ Card et al. (2018) show that up to age 18 for boys and up to age 16 for girls, there is little selection on observable characteristics into living with one or more parent (Card et al., 2018, p. 14).

Income upward mobility for 1980s birth cohorts For contemporary measures of upward mobility in commuting zones, I use data made available by Chetty and Hendren (2018b) and Chetty et al. (2020a). Based on the universe of federal income tax records from 1996-2012, the data contain measures of income upward mobility by childhood commuting zone for individuals born between 1980 and 1986. Parents and children in these records were linked via dependent claiming. The key measure of upward mobility is estimated mean individual or household income rank, conditional on parent household income rank. Household income measures for parents and children are drawn from Adjusted Gross Income on 1040 tax returns, and individual income rank is measured using income reported on W-2 forms, unemployment or disability insurance benefits, or half of household self-employment income where relevant.

Income for individuals in this sample is income at age 26, during the years 2006-2012, and income rank is rank in the national income distribution for individuals from the same birth cohort. Parent income is measured using returns filed when individuals were between the ages of 14 and 20, and parent income rank is rank in the national parent income distribution by child birth cohort. Separate upward mobility estimates are available for individuals from the 25th and 75th percentile of the parent income distribution. Estimates are also available separately by gender.

Childhood exposure effects of commuting zones I use an alternative measure of upward mobility in the 2000s from Chetty and Hendren (2018c): the childhood exposure effects of commuting zones. Starting from the universe of tax filers described above, the authors restricted the sample to individuals whose parents moved once across commuting zones during their childhood. They then compare the outcomes of children exposed for more or less time to a given commuting zone based on children's ages at the time their families moved. I use these outcomes estimated solely off of these mover families. The estimates reflect the causal effect of one additional year of childhood in a given commuting zone relative to an average commuting zone, for an arbitrary child. The outcome of interest is adult income rank at age 26. The estimates and assumptions behind them are discussed in greater detail in Section D.4.

Race-specific measures of upward mobility Race-specific measures of upward mobility come from Chetty et al. (2020b). These data are based on the same universe of federal income tax records as the measure described above; however, they cover a slightly different set of birth cohorts: 1978-1983. Individual federal income tax records were linked to the US Census in order to retrieve information on race as well as additional outcomes measured by the Census. The data contain the estimated mean individual or household income rank, conditional on parent household income rank, of Black and white men and women at the 25th and the 75th percentiles of the parent income distribution by childhood commuting zone. In this dataset, outcomes are measured in 2015 when individuals were between the ages of 32 and 37.

Appendix B City demographics (1940-1970), CZ sample construction and descriptive statistics, and geographic crosswalks

The following section provides additional details on the construction of the analysis sample of cities and commuting zones described in Section 3.2. I applied two selection criteria for CZs in the sample. First, following prior literature on the Great Migration, I restrict to locations in states that were net receivers of Black migrants. These include states in the northeastern, midwestern, and western census regions plus Maryland, Delaware, and Washington, D.C., which are located in the Southern census region, but on net received Black migrants between 1940 and 1970 (see Boustan (2016a)).

Second, I restrict to CZs for which I'm able to obtain data on their urban Black population between 1940 and 1970. I draw on two main sources of data to construct historical Black population measures for cities in northern commuting zones in 1940 and 1970: the complete count 1940 US census and the County and City Data Books 1944-1977 series (CCDB) available from the Inter-university Consortium for Political and Social Research (ICPSR), which contains information on cities with a population of 25,000 or more in the survey year (U.S. Census Bureau, Department of Commerce, 2008, 2012).

I obtain measures of the 1940 Black population from census, as the CCDB only report information on the number of white and non-white individuals in cities that year. Information on the Black population in cities in 1970 is obtained from the CCDB. My sample is therefore restricted to cities in the CCDB that can be matched to the 1940 census and to those with non-missing Black population data in 1970, a total of 294 cities. Approximately 78% of the Black population in the commuting zones in the sample resided in one of these 296 cities while 56% of the overall population did so.

The following cities from the CCDB could not be identified in the 1940 census: Boise City, ID; East Providence, RI; Huntington Park, CA; West Haven, CT; and Warwick, RI. I drop these cities from the analysis due to missing data. I also supplement my sample with two cities for which I manually recorded the Black population data using the 1970 Census: Butte, MT and Amsterdam, NY. Both cities received Black southern migrants between 1935 and 1940, but data on their Black population in 1970 was not available in the CCDB. Including these two cities brings the total number of cities to 296 from 294 and the total number of commuting zones in the sample to 130 from 128. Finally, the city of New Albany, IN is in the Louisville, KY commuting zone, which is included in the sample. Results are robust to excluding this commuting zone (see Figure D7).

B.1 Sample commuting zones and their demographic characteristics

Table B1 lists the commuting zones in the sample.

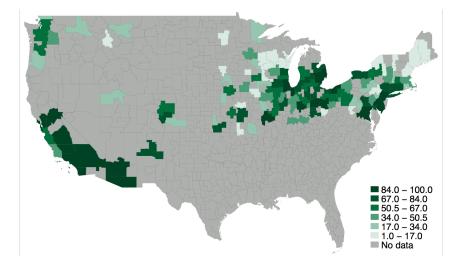
TABLE B1:	Commuting	ZONES	IN	SAMPLE	
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Phoenix, AZ	Rockford, IL	Joplin, MO	Youngstown, OH
Tucson, AZ	Springfield, IL	Kansas City, MO	Zanesville, OH
Bakersfield, CA	Center, IN	Springfield, MO	Eugene, OR
Fresno, CA	Concord, IN	St. Joseph, MO	Portland, OR
Los Angeles, CA	Evansville, IN	St. Louis, MO	Allentown, PA
Sacramento, CA	Fort Wayne, IN	Butte-Silver Bow, MT	Altoona, PA
San Diego, CA	Gary, IN	Great Falls, MT	Erie, PA
San Francisco, CA	Indianapolis, IN	Fargo, ND	Hagerstown, PA
San Jose, CA	Lafayette, IN	Lincoln, NE	Harrisburg, PA
Santa Barbara, CA	Muncie, IN	Omaha, NE	Philadelphia, PA
Colorado Springs, CO	South Bend, IN	Manchester, NH	Pittsburgh, PA
Denver, CO	Terre Haute, IN	Newark, NJ	Reading, PA
Pueblo, CO	Wayne, IN	Albuquerque, NM	Scranton, PA
Bridgeport, CT	Hutchinson, KS	Albany, NY	Williamsport, PA
Washington, DC	Topeka, KS	Amsterdam, NY	Providence, RI
Wilmington, DE	Wichita, KS	Buffalo, NY	Sioux Falls, SD
Burlington, IA	Louisville, KY	Elmira, NY	Salt Lake City, UT
Cedar Rapids, IA	Boston, MA	New York, NY	Burlington, VT
Clinton, IA	Pittsfield, MA	Poughkeepsie, NY	Bellingham, WA
Des Moines, IA	Springfield, MA	Syracuse, NY	Seattle, WA
Dubuque, IA	Baltimore, MD	Union, NY	Spokane, WA
Mason City, IA	Cumberland, MD	Watertown, NY	Yakima, WA
Ottumwa, IA	Bangor, ME	Canton, OH	Eau Claire, WI
Sioux City, IA	Portland, ME	Cincinnati, OH	Green Bay, WI
Waterloo, IA	Detroit, MI	Cleveland, OH	Kenosha, WI
Bloomington, IL	Grand Rapids, MI	Columbus, OH	La Crosse, WI
Chicago, IL	Jackson, MI	Dayton, OH	Madison, WI
Davenport, IL	Kalamazoo, MI	Lima, OH	Milwaukee, WI
Decatur, IL	Lansing, MI	Lorain, OH	Oshkosh, WI
Edwardsville, IL	Saginaw, MI	Mansfield, OH	Sheboygan, WI
Galesburg, IL	Duluth, MN	Scioto, OH	Wausau, WI
Peoria, IL	Minneapolis, MN	Steubenville, OH	
Quincy, IL	Rochester, MN	Toledo, OH	

Notes: Name refers to largest city in the commuting zone.

Figure C4 shows the distribution of Black population change throughout the commuting zones in the sample. As indicated by the figure, there is within region variation in the intensity of the Migration. Table B2 shows the CZ Black population share as well as the urban Black population share in the sample of cities within the sample CZs. The CZ Black population share more than doubles from 1940 to 2000, from 2.38% to 5.85% while the urban Black population share increases fivefold over the period, from 3.14% to 15.88%.

FIGURE B1: MAP OF 1940-70 CHANGE IN THE BLACK POPULATION



Notes: This map depicts Great Migration commuting zones and each CZ's percentile change in the Black population between 1940 and 1970. *Data sources*: CCDB; IPUMS complete count 1940 US census; Boustan (2016a).

TABLE B2: BLACK SHARE IN SAMPLE

	1940	1970	2000
CZ Black population share	2.382	4.160	5.854
	(8.11)	(10.57)	(11.89)
Observations	130	130	130
City Black population share	3.135	10.03	15.88
	(12.80)	(13.45)	(14.81)
Observations	296	296	296

Notes: This table reports the mean percentage (and standard deviation in parentheses) of the Black population in the analysis sample of 130 Great Migration commuting zones and the 296 cities for which the instrument for historical Black migration is constructed. Section 3 and 4 explain the criteria for selection of the sample and the construction of the instrument, respectively. *Data sources*: IPUMS complete count 1940 US census; CCDB; Boustan (2016a).

B.2 Geographic crosswalks

B.2.1 Historical county to 1990 CZ crosswalks

To construct the geographic crosswalks used in the analysis, polygon shapefiles for US geographic areas were downloaded from IPUMS *National Historical Geographic Information Systems* (NHGIS) (Manson et al., 2021) and merged based on spatial location using ArcGIS software. Listed below are the raw files and the website where they can be downloaded.

Raw data files from NHGIS and Census (U.S. Census Bureau, 2001, 1998)

- 1. US_county_1940.shp
- 2. US_county_1990.shp
- 3. US_msacmsa_1990.shp
- 4. US_necma_1990.shp

The following procedure was used to crosswalk between historical county boundaries (1920-1940) and 1990 commuting zones. Using ArcGIS, polygon shapefiles were converted to points representing the centroid of the polygon and then merged to the commuting zone polygon containing the centroid.³³

Because CZs are aggregations of 1990 counties, historical counties are matched to the CZ in which the geographic centroid of their 1940 borders falls. This procedure allows me to rapidly assign many historical county-level datasets to 1990 commuting zones. However, this procedure may result in assignment errors if county borders change substantially over time.

The vast majority of the counties in the paper's sample did not experience boundary changes over the timeframe of analysis. The ICPSR code assigned to each county has a numeric flag for counties that were dissolved and/or merged before 1970 (a final digit of 5); no counties in the sample fall into this category.

Data on changes in county boundaries can be obtained from the Atlas of Historical County Boundaries at the Newberry Library.³⁴ Of the 776 counties in the sample, only 32 had mapped boundary changes, representing 4.12 percent. The majority of these boundary changes are referred to by the Newberry Library researchers as "small," many too small to map. While most counties in the sample changed boundaries rarely if at all, one notable outlier is the Denver, Colorado metro area, where Denver, Jefferson, Adams, and Arapahoe Counties swapped patches of land at a sustained pace between 1940 and 1970. However, these counties fall within the Denver-Boulder-Longmont, CO commuting zone, therefore their border changes do not affect the accuracy of the matching procedure.

 $^{^{33}}$ The commuting zone polygon was created by dissolving borders between counties in the commuting zone using the crosswalk between 1990 counties and commuting zones from Autor and Dorn (2013).

 $^{^{34}{\}rm See}$ the information available at the following webpage: https://publications.newberry.org/ahcbp/.

B.2.2 City name standardizing

Names of cities in all city-level data digitized or collected for this paper were first standardized to be consistent with those in the 2010 U.S. place point shapefile from NHGIS. Places in the 2010 US place point file were matched to the county or CZ they fell within, allowing for the matching of city-level datasets to counties and commuting zones. In instances where a city did not appear in the 2010 US place point file, the city was assigned to the same CZ as places geographically close to the city in the 2010 US place point file.

Appendix C Great Migration shift-share instrument

This appendix details the construction of the shift-share instrument for the Great Migration, beginning with the construction of the shares from pre-1940 migrant location choices and following with the prediction of migration from southern counties using a machine learning approach.

C.1 Pre-1940 Black southern migrant shares

I measure Black southern migrant shares using the IPUMS version of the complete count 1940 census (Ruggles et al., 2021). The 1940 census was the first census in which enumerators asked individuals to report their place of residence in 1935. There are several advantages to this approach of measuring pre-1940 Black migration patterns. The first is that I am able to observe the universe of enumerated recent Black southern migrants, generating a nearly complete picture of recent migration flows into northern cities. The second is that the census microdata allow me to observe fine geographies for individuals' 1935 place of residence, including city and county. I define a recent Black southern migrant as a Black individual who reported a southern county of residence in 1935, but was enumerated in a different county (whether southern or not) in 1940. There are over 340,000 such individuals.

Using this population of recent Black southern migrants, I construct the share of migrants from each 1935 southern county j who settled in a northern city c by 1940:

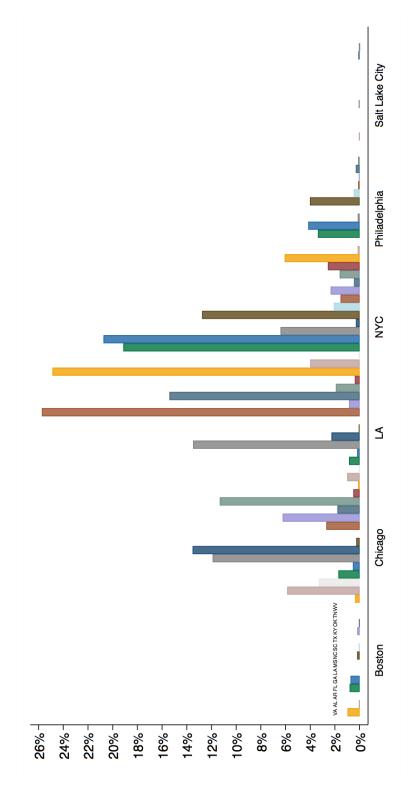
$$\omega_{jc}^{1935-1940} = \frac{b_{cj}}{b_j} \tag{12}$$

where b_j is the number of Black individuals who listed j as their county of residence in 1935, and b_{cj} is the number of Black individuals who were enumerated

in city c.

Figure C1 depicts $\omega_{jc}^{1935-1940}$ for a select group of cities and southern counties. Depicted is the share of 1935-1940 Black migrants from the largest sending county for each southern state who settled in the following cities: Boston, Chicago, Los Angeles, New York, Philadelphia, and Salt Lake City. The figure captures the immense heterogeneity in settlement patterns across and volume of migration into the cities in question.

Descriptive evidence on migrant characteristics Figure C2 shows the educational distribution for 1935-1940 Black southern migrants aged 25 plus. The median Black southern migrant moving between 1935 and 1940 had at least 5-8 years of schooling, comparable to the national median. This is striking given low levels of education among Black children in the South at this time (Margo, 1990). Finally, Figure C3 explores selection of migrants relative to northern incumbent Black families in the North. If anything, Black children from low socioeconomic status families whose parents were southern born had better educational outcomes than those whose parents were northern born.





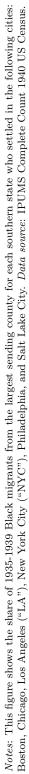
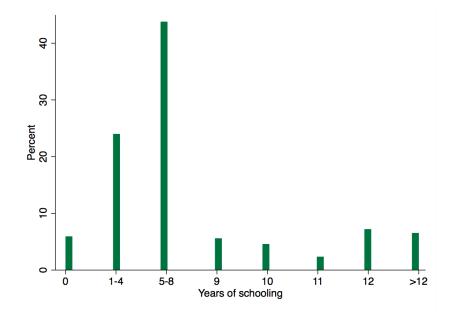
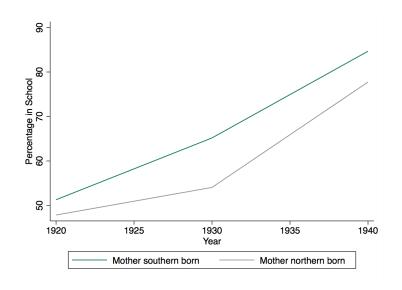


FIGURE C2: 1935-1940 Black southern migrant educational attainment

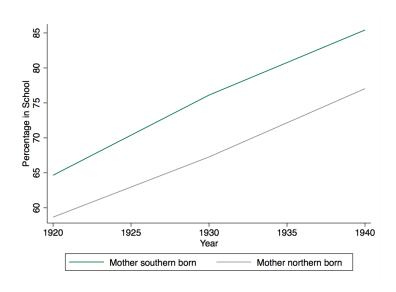


Notes: Histogram of years of schooling for 1935-1940 Black southern migrants aged 25 plus. *Data source*: IPUMS Complete Count 1940 US Census.

FIGURE C3: SCHOOL ATTENDANCE FOR BLACK TEENS IN NORTH WITH SOUTHERN- VS. NORTHERN-BORN MOTHERS



(a) Black teens with illiterate mothers



(b) Black teens with low-occ-score fathers

Notes: 1920-1940 school attendance rates (in percentage points) for Black 14-17 year-old boys and girls by mother birth region. Data sources: IPUMS Complete Count 1920-1940 US Censuses.

C.2 Post-LASSO prediction of southern county net migration

Under the assumption that county-level variation in southern economic indicators from 1940-1970 is uncorrelated with northern destination city characteristics for migrants from those counties, I view estimating southern county net migration rates as a pure prediction problem. Belloni et al. (2011) propose a machine-learning-based estimation of the first stage in an instrumental variables context where the number of instruments is large relative to the number of observations. In my case, I use this approach to select predictors in the "zero" stage prediction of net migration from southern counties using southern push factors³⁵:

$$\operatorname{mig rate}_{it} = \beta_0 + Z'_{it-10} \boldsymbol{\beta} + \varepsilon_{jt}, \qquad (13)$$

for $t \in \{1950, 1960, 1970\}$ where m_{jt} is net migration for southern county j from the decade of t - 10 to t and \mathbf{Z}'_{jt-10} is the set of predictors measured in t - 10. Using LASSO, I shrink the set of predictors to an optimal subset. Excluding a predictor from the subset corresponds to setting the respective element of $\boldsymbol{\beta}$ to zero. More explicitly, LASSO solves the following problem:

$$\min_{\beta_0,\beta} \left\{ \frac{1}{N} \sum_{j=1,\dots,1223} \left(\text{mig rate}_{jt} - \beta_0 - Z'_{jt-10}\beta + \varepsilon_{jt} \right) \right\} \quad \text{subject to} \quad \sum_{k=1}^9 |\beta_k| \le p_{jt}$$

In a "zero stage," I predict where p is the tuning parameter and β_k are the coefficients on each of the nine predictors in Z'_{jt-10} as suggested by Boustan (2010): the percent acreage in cotton; percent tenant farms; share of the labor force in agriculture; indicator for being in a tobacco-growing state and the interaction between tobacco growing state and share in agriculture; WWII spending per capita; share of the labor force in mining, an indicator for being in a mining state (OK and TX), and the interaction between the two.

 $^{^{35}}$ Southern county net-migration rates are taken from Boustan (2016b).

For each decade, I use five-fold cross-validation to choose the tuning parameter p that minimizes the expected prediction error.

In my case, LASSO selects the following for each year:

Variables selected in 1940:

- Percent tenant farms
- Share of the labor force in agriculture
- WWII spending per capita
- Percent acreage in cotton
- Share of the labor force in agriculture × Tobacco growing state
- Indicator for mining state
- Indicator for mining state \times Share of the labor force in mining

Variables selected in 1950:

- Percent tenant farms
- Share of the labor force in agriculture
- WWII spending per capita
- Percent acreage in cotton
- Percent acreage in tobacco
- Indicator for mining state
- Indicator for mining state × Share of the labor force in mining
- Share of the labor force in mining

Variables selected in 1960:

- Percent tenant farms
- Share of the labor force in agriculture

- Indicator for tobacco growing state
- Share of the labor force in agriculture × Tobacco growing state
- Percent acreage in cotton
- Indicator for mining state
- Indicator for mining state × Share of the labor force in mining
- Share of the labor force in mining

Using LASSO-selected variables improves the F-statistic for county outmigration prediction from 1940-1950 from 11.56 to 14.78. The F-statistics in the models for county outmigration prediction from 1950-1960 and 1960-1970 are identical using the original set of variables in Boustan (2010) and the LASSO-selected set.³⁶

Given this choice of included predictors, I estimate Equation (13) using OLS to predict net migration from county j, \hat{m}_{jt} , for each decade $t \in$ {1950, 1960, 1970}.³⁷ Next, I generate predicted migration into northern city c, \hat{m}_{ct} , by multiplying the share of pre-1940 migrants from each county by the predicted number of migrants leaving that county between 1940 and 1970:³⁸

$$\hat{m}_{ct} = \sum_{j=1,\dots,1223} (\omega_{cj}^{1935-40} \cdot \hat{m}_{jt})$$

where $\omega_{cj}^{1935-40}$ is the share of Black migrants from southern county j living in city c. The estimated total Black in-migration is calculated as $\hat{m}_c =$

 $^{^{36}}$ Chernozhukov et al. (2018) discuss inference adjustment in empirical settings where machine learning is used; they show that in a variety of empirical examples, qualitative conclusions of results remain unchanged after inference adjustment.

³⁷Direct measures of county-level in-migration and out-migration are not available for this time period, so I use net migration estimates produced by Boustan (2010) and made available in Boustan (2016a).

³⁸Because the available figures are net migration figures, and some southern counties experienced positive net migration (in-migration) as opposed to negative (out-migration), this procedure may result in predicted *decreases* in the Black population. This is the case for a small share of the commuting zones in the sample, particularly those in western states that are more likely to be connected to counties in Oklahoma or Texas, for example, some of which experienced net in-migration between 1940 and 1970.

 $\sum_{t \in \{1950, 1960, 1970\}} \hat{m}_{ct}$. Finally, I update the estimated share of Black residents in city $c, \hat{b}_{c,t}$, as

$$\hat{b}_{c,t} = \hat{b}_{c,t-10} + \hat{m}_{c,t} \tag{14}$$

where $\hat{b}_{c,1940} = b_{c,1940}$ as observed in the data.

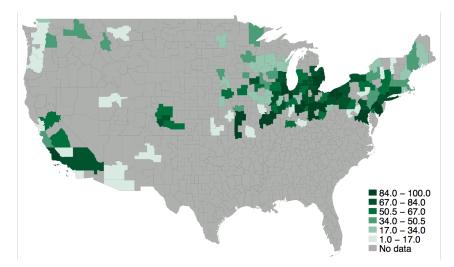


FIGURE C4: MAP OF GREAT MIGRATION INSTRUMENT

Notes: This map depicts Great Migration commuting zones and each CZ's predicted percentile change in the Black population between 1940 and 1970, predicted using the methods described in Appendix C. *Data sources*: CCDB; IPUMS complete count 1940 US census; Boustan (2016a).

Appendix D Additional upward mobility results and robustness

This appendix provides additional results on upward mobility as well as further robustness checks on the main findings. I begin with descriptive analysis of the change in the geography of upward mobility between 1940 and 2015. I then provide additional results and supporting evidence on the impact of the Great Migration on upward mobility for recent cohorts.

D.1 Long run change, 1940-2015

In this section, I provide correlations between measures of educational upward mobility in 1940 with income upward mobility in 2015. Following a method similar to Card et al. (2018), I measure the fraction of teenagers from house-holds in which the household head has 5-8 years of schooling³⁹ who obtain at least 9 years of education. The measure of income upward mobility in 2015 consists of estimated average adult income rank at the commuting zone level, for children from different parent income percentiles, where adult income is measured between the ages of 32 and 37.⁴⁰ Section 3.1 describes these data in much greater detail.

In Table D.1, I report the correlation coefficients between historical and contemporary upward mobility measures separately by race and gender. For white men and women, historical educational upward mobility is positively correlated with income upward mobility across commuting zones today. However, for Black men and women, these measures are virtually uncorrelated. This racial difference is particularly pronounced among men. Figures D1 shows the correlation between the historical measure and the contemporary measure for Black men in the top panel and for white men in the bottom panel.

TABLE D1: CORRELATION BETWEEN HISTORICAL AND CON-TEMPORARY UPWARD MOBILITY MEASURES, BY RACE AND GENDER

	Men	Women
Black	-0.18	0.28
White	0.43	0.54

Notes: Correlation coefficients between 1940 and 2015 measures of upward mobility, by race and gender. The sample in each column is the set of CZs within each gender for which both Black and white upward mobility measures can be computed.

³⁹Approximately the median of adult education in 1940.

 $^{^{40}}$ The children come from 1980s birth cohorts (1978-1983).

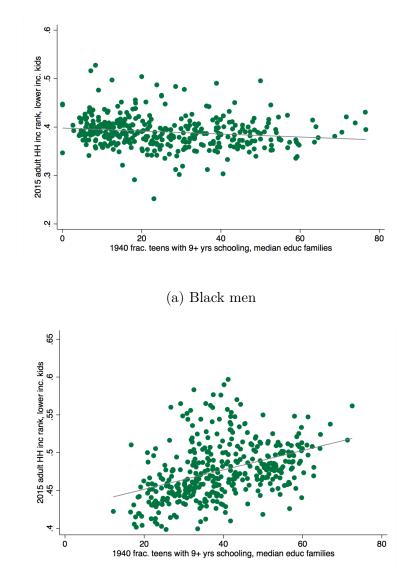


FIGURE D1: CORRELATION 1940 & 2015 UPWARD MOBILITY

(b) White men

Notes: This figure depicts scatter plots of the relationship between historical upward mobility and contemporary upward mobility for Black and white men. In panel (a), the right hand side ("RHS") is 1940 educational upward mobility defined as fraction of 14-18 year old Black boys who have at least 9 years of schooling, from families where the household head has 5-8 years of education. The left hand side ("LHS") is expected average individual adult income rank based on 2014-2015 IRS tax returns of Black men from 1978-1983 birth cohorts who come from families at the 25th percentile of the parent income distribution. Panel (b) shows the same relationship as in panel (a) for white men. In order to compare the same set of commuting zones and to minimize the influence of CZs with small numbers of Black children, I restrict the sample of CZs in both panels to those with at least 10 14-17 year old Black boys in 1940 and at least 10 Black men in the IRS sample. The correlations between historical and contemporary upward mobility are reported for Black and white women in Appendix Table D1. *Data sources*: IPUMS for 1940 measure; and Chetty, Hendren, Jones, and Porter (2018) for 2015 measures.

D.2 Supporting evidence and additional results

This section provides supporting evidence for the upward mobility results in the paper as well as results on additional outcomes or subgroups.

D.2.1 Coefficients on baseline controls

Table D2 provides the coefficients on the controls in the baseline specification of the reduced form effect of the Great Migration instrument on upward mobility. Two controls are worth noting. First, the share of the labor force in manufacturing is negatively correlated with average upward mobility for low income families and upward mobility for Black men and women from low income families, but uncorrelated with childhood exposure effects. This suggests sorting of families who tend not to produce better outcomes for their children in places with historically high manufacturing employment (today's Rust belt locations). The other control that is negatively correlated with all four outcomes is the share of the urban population in 1940 made up of Black southern migrants from any southern state. Including this control is important for restricting to the idiosyncratic variation between Great Migration destination cities and southern origin counties and subsequent shocks to the latter as drivers of in-migration as opposed to destinations that were markedly different at baseline.

	Average	Expos. Effects	Black, p25	Black, p75
\hat{GM}	-0.0370	-0.00232	-0.0264	-0.0386
	(0.00974)	(0.000631)	(0.0114)	(0.0169)
Edu. Upward Mobility 1940	0.0163	-0.000572	0.00554	-0.0348
	(0.0391)	(0.00212)	(0.0457)	(0.0679)
Share of LF employed in manufacturing, 1940	-0.152	-0.00323	-0.0835	-0.00524
	(0.0271)	(0.00192)	(0.0317)	(0.0471)
Black Southern Mig 1935-1940	-4.312	-0.0820	-0.383	-2.014
	(1.446)	(0.0671)	(1.698)	(2.523)
Midwest	-0.536	0.0981	-1.449	-0.870
	(0.603)	(0.0365)	(0.705)	(1.048)
South	-2.004	0.167	-0.294	1.430
	(1.306)	(0.0758)	(1.527)	(2.270)
West	-2.682	-0.100	-1.575	-1.691
	(0.872)	(0.0459)	(1.028)	(1.528)
R-squared	0.481	0.249	0.185	0.110

TABLE D2: UPWARD MOBILITY RESULTS WITH COEFFICIENTS ON BASELINE CONTROLS

Notes: This table reports coefficients on the baseline controls from the key analyses in the paper. In column 1, the outcome is average upward mobility for low income families; in column 2, the outcome is childhood exposure effects; in column 3, the outcome is average upward mobility for Black men from low income families; in column 4, the outcome is average upward mobility for Black men from high income families. The unit of observation is a CZ. Independent variable is predicted change in Black population share between 1940 and 1970. Baseline controls include share of CZ population made up of 1935-1939 Black southern migrants from any southern county, median education levels in 1940, share of employment in manufacturing in 1940, and census region fixed effects. The omitted region is the Northeast. *Data sources*: CCDB; IPUMS complete count 1940 US census; Boustan (2016a); Chetty and Hendren (2018b); Chetty et al. (2020a).

D.2.2 Alternative measures of father economic status for historical measures of upward mobility

Table D3 reproduces the results in Table 3 using alternative measures of father socioeconomic status. Columns 1-5 use the Lasso Industry Demographic and Occupation (LIDO) Score from Saavedra and Twinam (2020a). The LIDO score improves on the occupational income score by using additional variables,

namely including industry and demographic characteristics, and a machine learning approach to predict earnings. Column 6 uses wage income from the 1940 Census as the measure of father's socioeconomic status. As in Table 3, the results show no relationship between the instrument for Black migration and historical measures of educational upward mobility.

	Percentage teens with low SES fathers attending school						
		Ι	LIDO scor	e		Wage Income	
	1900	1910	1920	1930	1940	1940	
\hat{GM}	0.009	0.033	0.030	0.023	0.012	0.007	
	(0.030)	(0.027)	(0.027)	(0.026)	(0.016)	(0.014)	
Baseline mean	53.652	74.532	62.824	73.222	80.854	81.441	
Std Dev	9.239	8.221	8.219	8.209	5.793	5.296	
Observations	130	130	130	130	130	130	
Baseline Controls	Υ	Υ	Υ	Υ	Υ	Υ	

TABLE D3: PLACEBO TEST OF IDENTIFICATION STRATEGY USING PRE-1940 UPWARD MOBILITY WITH ALTERNATIVE MEASURES OF FATHER SES

Notes: This table reports the effect of the Great Migration on pre-1940 educational upward mobility and attainment. In columns 1 through 3, the dependent variable is the school attendance rate (in percentage points) of 14-17 year-old boys and girls with below-median Lasso Industry Demographic and Occupation (LIDO) score fathers in 1920, 1930, and 1940, respectively. The LIDO score, developed by Saavedra and Twinam (2020b), uses industry and demographic information, in addition to occupation, from the 1950 Census to predict earnings. In column 6, father's wage income is used as a measure of father's socioeconomic status. Independent variable is the instrument for black population increase from 1940 to 1970: the percentile of predicted black population increase, defined as the interaction between pre-1940 black southern migration patterns and post-1940 outflows of migrants as predicted by southern economic factors alone. Baseline 1940 controls include share of urban population made up of 1935-1940 black southern migrants, educational upward mobility, share of labor force in manufacturing, and census region fixed effects. Standard errors are in parentheses. *Data sources*: IPUMS complete count 1940 US census; Boustan (2016a).

D.2.3 Relationship between Great Migration instrument and 1940 CZ characteristics

Table D4 examines the relationship between the instrument for the Great Migration and socioeconomic characteristics of the destination CZs in 1940, including marriage rates, log mean occupation score, and log mean wage income. The results show no systematic relationship between the instrument for Black migration and these socioeconomic characteristics of the destination CZs.

	Married	Married, Spouse Present	Mean Occscore	Mean wage inc.
\hat{GM}	-0.001	-0.005	0.000	0.004
	(0.008)	(0.009)	(0.000)	(0.002)
Baseline mean	74.388	71.185	2.580	11.426
Std Dev	2.326	2.900	0.085	0.640
Observations	130	130	130	130
Baseline Controls	Υ	Y	Υ	Υ

TABLE D4: Great Migration instrument association with destination CZ characteristics in $1940\,$

Notes: This table reports the relationship between the instrument for the Great Migration and baseline 1940 socioeconomic characteristics of commuting zones. CZ characteristics are measured on adults aged 25 and older. The dependent variable in column 1 is the marriage rate; in column 2, the share of adults who are married with a spouse present in their household; in column 3, the log mean OCCSCORE; and in column 4, log mean wage income. Independent variable is the instrument for black population increase from 1940 to 1970: the percentile of predicted black population increase, defined as the interaction between pre-1940 black southern migration patterns and post-1940 outflows of migrants as predicted by southern economic factors alone. Baseline 1940 controls include share of urban population made up of 1935-1940 black southern migrants, educational upward mobility, share of labor force in manufacturing, and census region fixed effects. Standard errors are in parentheses. *Data sources*: IPUMS complete count 1940 US census; Boustan (2016a).

D.2.4 Great Migration effect on high income families

In Table D5 below, I report estimates of the effect of the Great Migration on average upward mobility for individuals with parents from the 75th percentile of the parent income distribution ("high income" families). The results show more modest impacts on this group relative to individuals from low income families (see Section 5.1 in the main text). A 1-percentile increase in the historical Black population lowered household income rank by 0.054 percentile points (s.e. = 0.023). The effect varies across gender groups and measures of income. Both men's and women's household income is lower in places that experienced greater Black in-migration historically, but only men's earnings (individual income in columns 4-6) are affected.

I next examine the effect of the Migration on CZ childhood exposure effects for individuals from high income families. Consistent with the patterns described above, the Great Migration's impact on childhood exposure effects of commuting zones is more muted for high income compared to low income families, with effects on household income of about one half the size while the effects on individual earnings are more comparable.

			First Stag	ge on GM		
\hat{GM}	0.297 (0.0759)	0.297 (0.0759)	0.297 (0.0759)	0.297 (0.0759)	0.297 (0.0759)	0.297 (0.0759)
F-Stat	15.34	15.34	15.34	15.34	15.34	15.34

TABLE D5: GREAT MIGRATION IMPACT ON AVERAGE UPWARD MOBILITY OF HIGH INCOME FAMILIES IN $2000\mathrm{s}$

	Household Income Rank		Individual Income Rank			
	Pooled	Women	Men	Pooled	Women	Men
			Ordinary L	east Squares	3	
GM	-0.0413	-0.0373	-0.0453	-0.0169	-0.00149	-0.0316
	(0.00793)	(0.00789)	(0.00828)	(0.00796)	(0.0102)	(0.00809)
R-squared	0.529	0.521	0.530	0.503	0.467	0.470
			Reduce	ed Form		
\hat{GM}	-0.0161	-0.0137	-0.0184	-0.00839	-0.000228	-0.0165
	(0.00766)	(0.00752)	(0.00804)	(0.00717)	(0.00911)	(0.00748)
R-squared	0.445	0.448	0.438	0.490	0.467	0.426
			Two-stage	least squares	3	
GM	-0.0541	-0.0462	-0.0618	-0.0283	-0.000768	-0.0556
	(0.0232)	(0.0230)	(0.0244)	(0.0233)	(0.0297)	(0.0243)
none						
Ν	130	130	130	130	130	130
Mean Rank	58.82	60.40	57.28	57.95	55.39	60.44
SD Rank	2.570	2.533	2.684	2.510	3.118	2.470
$SD \ GM$	28.98	28.98	28.98	28.98	28.98	28.98

Notes: This table reports the estimated impact of the Great Migration on average upward mobility in the 2000s for men and women with high income parents. The unit of observation is a CZ. Dependent variable is mean income rank for individuals with parents at the 75th percentile of the parent income distribution. Income is measured from IRS tax returns for cohorts and parents of cohorts born between 1980 and 1986. Pooled income refers to pooling across men and women. Independent variable is the percentile of Black population increase during the Great Migration. The instrument is the percentile of predicted Black population increase, defined as the interaction between pre-1940 Black southern migration patterns and post-1940 outflows of migrants as predicted by southern economic factors alone. Baseline 1940 controls include share of urban population made up of 1935-1940 Black southern migrants, educational upward mobility, share of labor force in manufacturing, and census region fixed effects. Standard errors are in parentheses. *Data sources*: CCDB; IPUMS complete count 1940 US census; Boustan (2016a); Chetty and Hendren (2018b).

\hat{GM}	$First\ Stage\ on\ GM$							
	0.274 (0.0618)	0.273 (0.0617)	0.274 (0.0619)	0.273 (0.0618)	0.274 (0.0619)	0.274 (0.0618)		
F-Stat	19.68	19.53	19.65	19.56	19.63	19.71		

TABLE D6: GREAT MIGRATION IMPACT ON CHILDHOOD EXPOSURE EFFECTS FOR HIGH INCOME FAMILIES

	Household Income Rank			Individual Income Rank				
	Pooled	Women	Men	Pooled	Women	Men		
		Ordinary Least Squares						
GM	-0.00119	-0.0000847	-0.00267	-0.000736	0.00134	-0.00278		
	(0.000829)	(0.00115)	(0.00111)	(0.000804)	(0.00121)	(0.00113)		
R-squared	0.305	0.233	0.154	0.472	0.388	0.204		
	Reduced Form							
\hat{GM}	-0.00131	-0.00104	-0.00153	-0.00203	-0.00171	-0.00241		
	(0.000604)	(0.000839)	(0.000827)	(0.000564)	(0.000880)	(0.000823)		
R-squared	0.320	0.243	0.138	0.520	0.400	0.219		
	Two-stage least squares							
GM	-0.00479	-0.00382	-0.00557	-0.00742	-0.00622	-0.00877		
	(0.00232)	(0.00313)	(0.00297)	(0.00262)	(0.00361)	(0.00325)		
none								
Ν	130	130	130	130	130	130		
Precision Wt	Υ	Υ	Υ	Y	Υ	Υ		
Mean Expos FX	-0.00323	-0.0253	-0.0162	0.0305	0.0182	-0.00525		
SD Expos FX	0.175	0.228	0.212	0.195	0.270	0.222		
SD GM	24.40	24.08	24.29	24.52	24.33	24.38		

Notes: This table reports the estimated impact of the Great Migration on commuting zone childhood exposure effects. The unit of observation is a commuting zone. Dependent variable is the estimated causal impact of one additional year of childhood in the commuting zone on adult household income rank for men and women with parents at the 75th percentile of the parent income distribution. Income is measured from IRS tax returns for cohorts and parents of cohorts born between 1980 and 1986. The unit of observation is a commuting zone. Pooled income refers to household income, pooling across men and women. Independent variable is the percentile of Black population increase during the Great Migration. The instrument for Black population increase is the percentile of predicted Black population increase, defined as the interaction between pre-1940 Black southern migration patterns and post-1940 outflows of migrants as predicted by southern economic factors alone. Baseline 1940 controls include share of urban population made up of 1935-1940 Black southern migrants, educational upward mobility, share of labor force in manufacturing, and census region fixed effects. Standard errors are in parentheses. *Data sources*: CCDB; IPUMS complete count 1940 US census; Boustan (2016a); Chetty and Hendren (2018b).

D.2.5 Great Migration effect on labor market outcomes

Table D7 reports the effects of the Migration on measures of labor market outcomes, using data from Chetty et al. (2020b).

With the exception of the indicator for having positive W-2 tax form earnings, all other labor market measures come from the ACS⁴¹ and are not available for Black men and women in over a third of the commuting zones in the sample due to the limited sample.

The results suggest that Black men from low income families growing up in major Great Migration destinations today work fewer hours per week and are less likely to have positive W-2 earnings. I also find that white men from low income families are less likely to have positive W-2 earnings, though the effect is smaller than on Black men from similar parent income. I do not see significant effects for white men from low income families on hours worked although the coefficient is negative. Despite this negative effect on the probability of working for white men from low income families, I do not observe significant reductions in their adult income rank—the results reported in the main paper in Table 7. This may be due to other non-W-2 sources of income, for example self-employment or other kinds of earnings that are not reported on the W-2 form. There are no significant effects on other subgroups.

Of interest is the lack of an effect on the percent with positive W-2 earnings or on hours worked for Black women from low income families growing up in major Great Migration commuting zones. These results are consistent with Figure 8, which show that income rank based on individual earnings is not lower for Black women from low income families in Great Migration commuting zones. There is a small, positive effect of the Great Migration on the individual income rank of Black women from high income families, also reported in Figure 8, although this effect is not significantly different from zero. The results on labor supply suggest this is not a mechanism for their higher

⁴¹See Chetty et al. (2018), specifically Online Appendix A on data and variable construction, for details.

individual earnings. One potential explanation is that they have higher non-W-2 income through self-employment, for example, and this increased income is not captured through labor supply responses measured through positive W-2 earnings or hours worked.

	Low Parent Income				High Parent Income			
	Black Men	Black Women	White Men	White Women	Black Men	Black Women	White Men	White Women
	Dep var: Fraction with positive W-2 earnings at age 32							
GM (IV)	-0.158***	0.00303	-0.0932**	-0.0670	-0.171**	-0.0143	-0.0250	-0.00361
	(0.0443)	(0.0492)	(0.0302)	(0.0415)	(0.0605)	(0.0591)	(0.0164)	(0.0314)
Dep var mean	69.27	78.91	76.95	72.92	80.63	84.33	87.63	81.34
Ν	129	129	130	130	129	129	130	130
F-Stat	17.49	17.49	15.34	15.34	17.49	17.49	15.34	15.34
R-squared	0.387	0.246	0.343	0.377	0.0539	0.160	0.589	0.567
	Dep var: Mean weekly hours worked in past year							
GM (IV)	-0.229**	-0.0677	-0.0354	0.00114	-0.0714	0.0177	-0.0113	0.00341
	(0.0717)	(0.0450)	(0.0223)	(0.0262)	(0.0848)	(0.0600)	(0.0143)	(0.0215)
Dep var mean	23.57	25.88	33.56	25.22	30.43	29.89	38.79	29.91
Ν	87	89	130	130	87	89	130	130
F-Stat	8.963	12.17	15.34	15.34	8.963	12.17	15.34	15.34
R-squared	0.0154	-0.0454	0.461	0.259	0.284	0.0853	0.533	0.427
	Dep var: Fraction with positive hours in past year							
GM (IV)	-0.209	-0.0587	-0.0429	-0.0152	0.0860	-0.0421	-0.0131	0.0240
	(0.176)	(0.110)	(0.0331)	(0.0551)	(0.204)	(0.184)	(0.0172)	(0.0386)
Dep var mean	67.92	78.65	86.92	77.28	81.21	86.74	94.62	85.85
Ν	87	89	130	130	87	89	130	130
F-Stat	8.963	12.17	15.34	15.34	8.963	12.17	15.34	15.34
R-squared	0.0317	0.0590	0.429	0.279	-0.0262	0.0450	0.437	0.370
	Dep var: Hourly wage rank							
GM (IV)	0.122	0.0669	0.0316	0.0230	0.266	-0.164	0.00210	0.0547
	(0.145)	(0.122)	(0.0377)	(0.0457)	(0.205)	(0.177)	(0.0282)	(0.0365)
Dep var mean	37.76	37.70	44.92	39.93	46.35	47.32	56.36	50.98
Ν	83	84	130	130	83	84	130	130
F-Stat	7.143	7.420	15.34	15.34	7.143	7.420	15.34	15.34
R-squared	0.187	0.370	0.480	0.493	-0.00976	0.124	0.450	0.475
SD GM	20.88	21.37	28.98	28.98	20.88	21.37	28.98	28.98

TABLE D7: EFFECTS OF THE GREAT MIGRATION ON CHILDREN'S LABOR MARKET OUTCOMES

Notes: This table reports the estimated impact of the Great Migration on average educational upward mobility in the 2000s for men and women with low income parents. The unit of observation is a commuting zone. Dependent variables are listed at the top of each panel. Independent variable is the percentile of Black population increase during the Great Migration. The instrument for Black population increase is the percentile of predicted Black population increase, defined as the interaction between pre-1940 Black southern migration patterns and post-1940 outflows of migrants as predicted by southern economic factors alone. Baseline 1940 controls include share of urban population made up of 1935-1940 Black southern migrants, educational upward mobility, share of labor force in manufacturing, and census region fixed effects. Standard errors are in parentheses. *Data sources*: CCDB; IPUMS complete count 1940 US census; Boustan (2016a); Chetty et al. (2020a).

D.2.6 Great Migration effect on incarceration and family outcomes

Table D8 reports the effects of the Migration on measures of incarceration and family structure, using data from the Opportunity Insights website (www. opportunityinsights.org).

Information on incarceration comes from the 2010 Census and indicates whether an individual is incarcerated as of April 1, 2010. Note that this measure is likely a noisy measure of interaction with the criminal justice system as it only measures incarceration for those in jail or a correctional institution on the day of the Census enumeration. Other types of interaction with the criminal justice system or those incarcerated on other days of the year but not April 1, 2010 will not be picked up by this measure.⁴²

The results indicate that growing up in a Great Migration CZ is associated with reduced father presence for both Black women and men from low income families—and to a reduced extent, for Black men from higher income families. I also find large effects on the probability of being married both for Black men and women from lower and higher parent income backgrounds. Finally, I find marginally significantly higher probability of giving birth as a teen for Black women from both higher and lower income families growing up in Great Migration commuting zones.

⁴²See Chetty et al. (2018), specifically Online Appendix A on data and variable construction, for details.

		L. D.	ent Income			U. I. D.	ent Income		
	Black Men			White Women	Black Men	0		White Women	
	Dep var: Fraction incarcerated on April 1st, 2010								
GM (IV)	0.0326	-0.0280*	0.0166	0.00511	0.0762	0.00887	0.00129	-0.0000359	
	(0.0549)	(0.0114)	(0.0117)	(0.00274)	(0.0484)	(0.0115)	(0.00270)	(0.000806)	
Dep var mean	13.59	0.976	3.074	0.412	5.197	0.358	0.690	0.101	
Ν	129	129	130	130	129	129	130	130	
F-Stat	17.49	17.49	15.34	15.34	17.49	17.49	15.34	15.34	
R-squared	0.0974	0.0118	0.165	-0.0432	-0.0736	0.0642	0.272	0.0767	
		Dep var: Father likely present							
GM (IV)	-0.269***	-0.299***	0.00758	-0.00542	-0.155^{*}	-0.167	-0.0206	-0.0271	
	(0.0716)	(0.0793)	(0.0657)	(0.0669)	(0.0752)	(0.0900)	(0.0165)	(0.0175)	
Dep var mean	40.21	39.58	64.90	62.37	92.49	91.90	97.67	97.35	
Ν	129	129	130	130	129	129	130	130	
F-Stat	17.49	17.49	15.34	15.34	17.49	17.49	15.34	15.34	
R-squared	0.575	0.644	0.201	0.220	0.382	0.343	0.582	0.586	
			Ĺ	Dep var: Fraction	n married in	2015			
GM (IV)	-0.0937*	-0.173***	-0.00605	-0.00800	-0.256***	-0.250*	-0.0148	-0.00234	
	(0.0408)	(0.0514)	(0.0378)	(0.0422)	(0.0744)	(0.101)	(0.0362)	(0.0334)	
Dep var mean	18.71	18.00	41.36	47.72	32.20	31.77	56.44	63.50	
Ν	129	129	130	130	129	129	130	130	
F-Stat	17.49	17.49	15.34	15.34	17.49	17.49	15.34	15.34	
R-squared	0.587	0.551	0.435	0.360	0.365	0.390	0.539	0.513	
	Dep var: Fraction gave birth as a teen								
GM (IV)		0.169^{*}		-0.0241		0.164^{*}		-0.0349	
		(0.0716)		(0.0561)		(0.0789)		(0.0260)	
Dep var mean		42.44		22.40		21.13		8.837	
Ν		129		130		129		130	
F-Stat		17.49		15.34		17.49		15.34	
R-squared		0.537		0.363		0.447		0.288	
SD GM		28.80		28.98		28.80		28.98	

TABLE D8: EFFECTS OF THE GREAT MIGRATION ON CHILDREN'S INCARCERATION AND FAMILY STRUCTURE

Notes: This table reports the estimated impact of the Great Migration on incarceration and family structure in the 2000s for men and women with low income parents. The unit of observation is a commuting zone. Dependent variables are listed at the top of each panel. Independent variable is the percentile of Black population increase during the Great Migration. The instrument for Black population increase is the percentile of predicted Black population increase, defined as the interaction between pre-1940 Black southern migration patterns and post-1940 outflows of migrants as predicted by southern economic factors alone. Baseline 1940 controls include share of urban population made up of 1935-1940 Black southern migrants, educational upward mobility, share of labor force in manufacturing, and census region fixed effects. Standard errors are in parentheses. *Data sources*: CCDB; IPUMS complete count 1940 US census; Boustan (2016a); Chetty et al. (2020a).

D.2.7 Great Migration effect on educational upward mobility

Table D9 reports the effect of GM on educational upward mobility for different racial and gender groups, using data from the Opportunity Insights website (www.opportunityinsights.org). Educational upward mobility is defined as the fraction of individuals with a high school diploma or GED, a community college degree, some college, or a college degree conditional on parent income rank.⁴³ Estimates for some groups are imprecise as data on educational outcomes are only available for individuals who received the ACS or Census long form in 2000 and CZ-level outcomes are only available in commuting zones with a sufficient number of the race and gender group in question. These restrictions decrease the number of commuting zones with outcome data for Black men and women in particular, to 98 and 102 commuting zones out of 130, respectively.

Still, the results are strongly consistent with the results on income upward mobility by race and gender group. Growing up in a Great Migration destination CZ lowers educational attainment for Black men from low income families, and the effects are significant at the 5%- or 10%-level, except for the college graduation outcome. The magnitudes suggest that a 1-percentile increase in historical Black migration leads to a decrease in high school graduation rates of 0.261 pp, in community college graduation rates of 0.246 pp, in the fraction obtaining some college education of 0.319 pp, and in college graduation rates of 0.108 pp (though not statistically different from zero). Effects on Black women from low income families tend to have the opposite sign but are less precise. There are no precisely estimated effects on Black women or men from high income families though the impacts on Black men are consistently negative while those on Black women are not. Across the board, the Great Migration has no effect on the educational outcomes of white men or women irrespective of parental income. Overall, these results are consistent with stronger effects on Black men and, in the case of educational attainment, appear to be an

 $^{^{43}\}mathrm{See}$ Chetty et al. (2018), specifically Online Appendix A on data and variable construction, for details.

important driver for the effects on Black men from low income families but with the exception of the college margin—are less likely to be important for explaining the effects on Black men from high income families.

	Low Parent Income High Parent Income							
	Black Men	Black Women	White Men	White Women	Black Men	Black Women	White Men	White Women
	Dep var: Fraction graduated from high school							
GM (IV)	-0.261*	0.101	0.00164	0.0281	-0.0623	-0.114	0.0285	0.0136
	(0.111)	(0.0999)	(0.0469)	(0.0468)	(0.110)	(0.112)	(0.0190)	(0.0146)
Dep var mean	73.16	81.05	78.54	84.32	85.64	91.31	90.96	94.04
Ν	114	113	130	130	114	113	130	130
F-Stat	11.18	8.961	15.34	15.34	11.18	8.961	15.34	15.34
R-squared	-0.00901	0.0138	0.196	0.141	0.0643	-0.108	0.00835	0.117
			Dep var:	Fraction graduat	ted from com	nunity college		
GM (IV)	-0.246**	0.0581	0.00706	0.0462	-0.121	0.123	-0.0136	0.0243
	(0.0908)	(0.125)	(0.0597)	(0.0801)	(0.147)	(0.185)	(0.0520)	(0.0585)
Dep var mean	17.63	29.63	24.57	36.29	37.32	51.59	51.73	64.80
Ν	98	102	130	130	98	102	130	130
F-Stat	12.54	10.98	15.34	15.34	12.54	10.98	15.34	15.34
R-squared	0.0124	0.199	0.307	0.221	0.220	0.0947	0.349	0.312
			I	Dep var: Fraction	with some c	ollege		
GM (IV)	-0.319*	0.199	-0.0213	0.00981	-0.0285	0.116	0.0132	0.0248
	(0.145)	(0.120)	(0.0684)	(0.0766)	(0.195)	(0.142)	(0.0414)	(0.0366)
Dep var mean	44.50	61.99	47.55	62.75	65.76	78.84	75.98	85.38
Ν	98	102	130	130	98	102	130	130
F-Stat	12.54	10.98	15.34	15.34	12.54	10.98	15.34	15.34
R-squared	0.0109	-0.0469	0.419	0.232	0.0686	0.00604	0.218	0.0711
			Dep	o var: Fraction g	raduated from	n college		
GM (IV)	-0.108	0.126	0.0540	0.0608	-0.125	-0.0313	0.0437	0.0706
	(0.0703)	(0.103)	(0.0498)	(0.0680)	(0.164)	(0.174)	(0.0512)	(0.0581)
Dep var mean	11.23	18.85	15.79	23.81	28.79	41.27	39.98	51.66
Ν	98	102	130	130	98	102	130	130
F-Stat	12.54	10.98	15.34	15.34	12.54	10.98	15.34	15.34
R-squared	0.123	0.0204	0.312	0.310	0.204	0.144	0.222	0.281
SD GM	24.57	25.30	28.98	28.98	24.57	25.30	28.98	28.98

TABLE D9: EFFECTS OF THE GREAT MIGRATION ON EDUCATIONAL MOBILITY

Notes: This table reports the estimated impact of the Great Migration on average educational upward mobility in the 2000s for men and women with low income parents. The unit of observation is a commuting zone. Dependent variables are the fraction of individuals with parents at the 25th percentile of the parent income distribution who graduated from high school, from community college, with some four-year college, and from a four-year college. Independent variable is the percentile of Black population increase during the Great Migration. The instrument for Black population increase is the percentile of predicted Black population increase, defined as the interaction between pre-1940 Black southern migration patterns and post-1940 outflows of migrants as predicted by southern economic factors alone. Baseline 1940 controls include share of urban population made up of 1935-1940 Black southern migrants, educational upward mobility, share of labor force in manufacturing, and census region fixed effects. Standard errors are in parentheses. *Data sources*: CCDB; IPUMS complete count 1940 US census; Boustan (2016a); Chetty et al. (2020a).

D.3 Heterogeneity by race, gender, and geography

This section explores further heterogeneity in the effects of the Great Migration by race, gender, and geography.

D.3.1 Regression results weighted by number of individuals underlying mobility estimates

Below I report the effect of the Migration on different race, gender, and parent income subgroups where the regressions are weighted by the number of individuals off of which upward mobility statistics are based on to address the fact that some CZ-level estimates are based off of relatively small numbers of individuals. Tables D10 and D11 report the OLS, reduced form, and 2SLS effects of a 1-percentile increase in the historical Black population on both household and individual income rank of Black and white men and women, respectively. Figure D2 reports the 2SLS effect of the Migration on individual income rank of different subgroups where the coefficient on the Great Migration has been scaled to represent a 1-standard-deviation increase. Overall, results are similar to unweighted estimates reported in Tables 6 and 7 and Figure 8 in the main text, with slightly larger negative effects on white men with low income parents and a smaller positive effect on Black women with high income parents.

	First Stage on GM							
\hat{GM}	0.272	0.273	0.270	0.272	0.273	0.270		
	(0.0542)	(0.0541)	(0.0542)	(0.0542)	(0.0541)	(0.0542)		
F-Stat	25.16	25.45	24.86	25.16	25.45	24.86		

TABLE D10: GREAT MIGRATION IMPACT ON BLACK FAMILIES, USING WEIGHTS

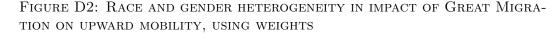
	Low Income			High Income					
	Pooled	Women	Men	Pooled	Women	Men			
			Ordinary L	east Squares					
GM	-0.0567	-0.0108	-0.0591	-0.0651	0.00550	-0.0586			
	(0.00992)	(0.0106)	(0.0113)	(0.00979)	(0.0128)	(0.0111)			
R-squared	0.797	0.802	0.699	0.797	0.777	0.637			
		Reduced Form							
\hat{GM}	-0.0264	-0.0117	-0.0271	-0.0190	0.00437	-0.0205			
	(0.00691)	(0.00688)	(0.00783)	(0.00729)	(0.00840)	(0.00785)			
R-squared	0.770	0.805	0.665	0.737	0.777	0.578			
	Two-stage least squares								
GM	-0.0973	-0.0428	-0.100	-0.0699	0.0160	-0.0758			
	(0.0247)	(0.0255)	(0.0280)	(0.0229)	(0.0299)	(0.0263)			
none									
Ν	129	129	129	129	129	129			
Mean Rank	0.332	0.403	0.389	0.453	0.493	0.515			
SD Rank	0.0275	0.0276	0.0315	0.0398	0.0504	0.0448			
SD GM	28.80	28.80	28.80	28.80	28.80	28.80			

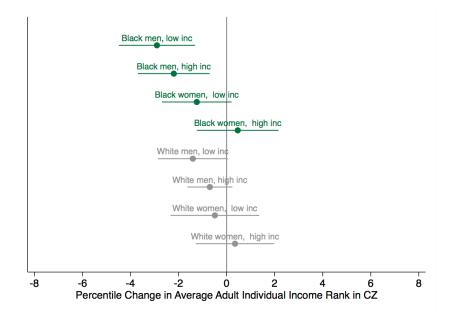
Notes: This table reports the estimated impact of the Great Migration on average upward mobility in the 2000s for Black men and women born between 1978 and 1983. The unit of observation is a CZ. The dependent variable is expected mean income rank for those with parents at the 25th and 75th percentile of parent income. Income is measured from IRS tax returns. The independent variable is the percentile of black population increase during the Great Migration. Regressions are weighted by the number of individuals per subgroup whose tax records where used to construct outcome measures. The instrument is the percentile of predicted black population increase, defined as the interaction between pre-1940 black southern migration patterns and post-1940 migrant outflows predicted by southern economic variables. Baseline 1940 controls include share of urban population made up of 1935-1940 black southern migrants, educational upward mobility, share of labor force in manufacturing, and census region fixed effects. Standard errors are in parentheses. *Data sources*: CCDB; IPUMS complete count 1940 US census; Boustan (2016a); Chetty et al. (2020a).

	First Stage on GM							
\hat{GM}	0.326	0.326	0.327	0.326	0.326	0.327		
	(0.0653)	(0.0653)	(0.0652)	(0.0653)	(0.0653)	(0.0652)		
F-Stat	25.02	25.00	25.05	25.02	25.00	25.05		
		Low Income	e	High Income				
	Pooled	Women	Men	Pooled	Women	Men		
			Ordinary I	Least Squares	3			
GM	-0.0120	0.0129	-0.0194	-0.0143	0.0193	-0.0198		
	(0.0123)	(0.0136)	(0.0106)	(0.00769)	(0.0122)	(0.00701)		
R-squared	0.531	0.531	0.516	0.624	0.573	0.590		
	Reduced Form							
\hat{GM}	-0.0101	-0.00550	-0.0158	-0.00385	0.00394	-0.00784		
	(0.00970)	(0.0108)	(0.00839)	(0.00616)	(0.00975)	(0.00568)		
R-squared	0.531	0.528	0.517	0.614	0.565	0.570		
	Two-stage least squares							
GM	-0.0309	-0.0168	-0.0485	-0.0118	0.0121	-0.0240		
	(0.0291)	(0.0325)	(0.0257)	(0.0181)	(0.0287)	(0.0165)		
none								
Ν	130	130	130	130	130	130		
Mean Rank	0.452	0.405	0.490	0.606	0.517	0.630		
SD Rank	0.0316	0.0326	0.0267	0.0223	0.0295	0.0187		
$SD \ GM$	28.98	28.98	28.98	28.98	28.98	28.98		

TABLE D11: GREAT MIGRATION IMPACT ON WHITE FAMILIES, US-ING WEIGHTS

Notes: This table reports the estimated impact of the Great Migration on average upward mobility in the 2000s for white men and women born between 1978 and 1983. The unit of observation is a CZ. The dependent variable is expected mean income rank for those with parents at the 25th and 75th percentile of parent income. Income is measured from IRS tax returns. The independent variable is the percentile of black population increase during the Great Migration. Regressions are weighted by the number of individuals per subgroup whose tax records where used to construct outcome measures. The instrument is the percentile of predicted black population increase, defined as the interaction between pre-1940 black southern migration patterns and post-1940 migrant outflows predicted by southern economic variables. Baseline 1940 controls include share of urban population made up of 1935-1940 black southern migrants, educational upward mobility, share of labor force in manufacturing, and census region fixed effects. Standard errors are in parentheses. *Data sources*: CCDB; IPUMS complete count 1940 US census; Boustan (2016a); Chetty et al. (2020a).





Notes: This figure plots coefficients from regressions of average upward mobility in the 2000s for white and Black men and women from low and high income parents on the instrument for Black population increases during the Great Migration, in approximately one standard deviation units. The unit of observation is a commuting zone. Upward mobility is defined as expected mean individual income rank where income is measured from IRS tax returns for cohorts born between 1978 and 1983. Regressions are weighted by the number of individuals per subgroup whose tax records where used to construct outcome measures. The instrument is the percentile of predicted Black population increase, defined as the interaction between pre-1940 Black southern migration patterns and post-1940 outflows of migrants as predicted by southern economic factors alone. A one standard deviation increase is approximately 30 percentiles. Baseline 1940 controls include share of urban population made up of 1935-1940 Black southern migrants, educational upward mobility, share of labor force in manufacturing, and census region fixed effects. *Data sources*: IPUMS complete count 1940 US census; Boustan (2016a); Chetty et al. (2020a).

D.3.2 Race-specific versus pooled upward mobility results

In this Appendix section, I show that the effect of the Migration on average upward mobility can be decomposed into an effect on Black and white individuals separately plus a "composition effect"—the effect of increasing the Black share of low income families in a CZ. Because Black children have lower upward mobility than their white counterparts across the US—potentially driven by factors such as systemic racism—this may lead to a divergence between the Great Migration's impact on race-specific average upward mobility versus average upward mobility pooled across racial groups. I formalize this issue into a decomposition between the effect of the Migration on the upward mobility of each group and the effect of the Migration on the Black share of low or high income families. In what follows, for simplicity of notation, I suppress place and parent income subscripts c and p, respectively.

To see this decomposition, first assume the Migration has no effect on the racial composition of low-income families by the 1990s when parent income for the cohorts in the mobility data is measured (in other words that the historical shock of the Great Migration to a local area's racial composition has faded by the 1990s). I construct the "average effect" of the Migration (β_{avg}) under this assumption by taking the weighted average of the estimated effect on Black and white individuals where the weights are the average Black and white shares of low income parents, for simplicity denoted as $\mathbb{E}[s]$ and $1 - \mathbb{E}[s]$ below,⁴⁴ across the CZs in my sample:

$$\beta_{\text{avg}} = \frac{\text{Cov}(y_{\text{avg}}, \hat{GM})}{\text{Var}(\hat{GM})} = \mathbb{E}[s] \cdot \frac{\text{Cov}(y_b, \hat{GM})}{\text{Var}(\hat{GM})} + (1 - \mathbb{E}[s]) \cdot \frac{\text{Cov}(y_w, \hat{GM})}{\text{Var}(\hat{GM})}$$

I then compare this to the impact of the Great Migration on the locally population-weighted average of Black and white outcomes in a CZ (β_{pooled}) where population weights equal the local Black share and white share of parents of national income rank p:

$$\beta_{\text{pooled}} = \frac{\text{Cov}(y_{\text{pooled}}, \hat{GM})}{\text{Var}(\hat{GM})} = \frac{\text{Cov}(s \cdot y_b + (1 - s) \cdot y_w, \hat{GM})}{\text{Var}(\hat{GM})}$$

These two effects are equal only if the following holds:

$$\frac{\operatorname{Cov}\left((s - \mathbb{E}[s]) \cdot y_b + (1 - s - (1 - \mathbb{E})) \cdot y_w, \hat{GM}\right)}{\operatorname{Var}(\hat{GM})} = 0$$

 $^{^{44}}$ When I estimate the pooled effect and the effect assuming no composition effect, I use the actual white share, not 1 minus the Black share. These shares do not sum to 1 due to the presence other racial groups in the same commuting zone. However, results are similar when I set the white share equal to 1 minus the Black share.

This can be seen by subtracting and adding y_{avg} from y_{pooled} :

$$y_{\text{pooled}} = y_{\text{avg}} + s \cdot y_b + (1 - s) \cdot y_w - \mathbb{E}[s] \cdot y_b + (1 - \mathbb{E}[s]) \cdot y_w,$$

which implies

$$\beta_{\text{pooled}} = \frac{\text{Cov}\left(y_{\text{pooled}}, \hat{GM}\right)}{\text{Var}(\hat{GM})} = \frac{\text{Cov}\left(y_{\text{avg}}, \hat{GM}\right)}{\text{Var}(\hat{GM})} + \frac{\text{Cov}\left((s - \mathbb{E}[s]) \cdot y_b + (1 - s - (1 - \mathbb{E})) \cdot y_w, \hat{GM}\right)}{\text{Var}(\hat{GM})}$$
$$= \beta_{\text{avg}} + \frac{\text{Cov}\left((s - \mathbb{E}[s]) \cdot y_b, \hat{GM}\right)}{\text{Var}(\hat{GM})} + \frac{\text{Cov}\left(1 - s - (1 - \mathbb{E})\right) \cdot y_w, \hat{GM}\right)}{\text{Var}(\hat{GM})} \quad (15)$$

These last two terms make up the "composition effect": the impact of the Migration driven by changes in the composition of families. The Great Migration increased the Black share of low income parents, which lowers average upward mobility irrespective of locations. Note, this can be attributed to a "systemic racism" component. Black children from families with similar income to white families still face barriers throughout the US, and these barriers can affect outcomes cumulatively through the life cycle. They include, for example, lower starting wealth or discrimination faced by their parents in the housing market, both of which may affect long-run investments in children despite similar income levels.

To quantify this composition effect empirically, I use county-level Census aggregates on the income distribution by race in 2000 from NHGIS (Manson et al., 2021) to construct the CZ-level Black (white) share of households in approximately the bottom quartile of the income distribution. I convert the bins to 2015\$ using the CPI-U-RS. I use the crosswalk from parent income ranks to 2015\$ levels available in Chetty et al. (2020a) to identify the bins in approximately the bottom quartile of the national parent income distribution in 2000. I can then estimate each component of equation 15.

Column 1 of Table D12 reports the estimated reduced form effect of the

Migration on Black individuals (from column 1 of Table 6) and column 2 on white individuals (from column 1 of Table 7). Column 3 reports the average of these estimates assuming no composition effect, β_{avg} . Column 4 reports this average effect plus the composition effects, $\frac{\text{Cov}\left((s - \mathbb{E}[s]) \cdot y_b, \hat{GM}\right)}{\text{Var}(\hat{GM})}$ and $\frac{\text{Cov}\left(1 - s - (1 - \mathbb{E})\right) \cdot y_w, \hat{GM}}{\text{Var}(\hat{GM})}$, reported further down in the same column as "Black Comp Effect" and "White Comp Effect." Column 5 reports the effect on locally weighted average outcomes, β_{pooled} , and column 6 the effect on the

pooled mobility estimate from Chetty et al. (2020a).⁴⁵

The results are consistent with the effect of the Migration on pooled mobility (columns 5 and 6), capturing both the effect on Black and white individuals separately, plus an effect of the change in the racial composition of low income parents in Great Migration CZs. Separately estimating the effect of the Migration by racial group purges the estimates of this racial composition effect. However, this procedure does not address other sources of unobserved heterogeneity across families that may affect children's long-run outcome. These additional sources of selection may be positive or negative. My comparison of the Migration's impact on average upward mobility versus the childhood exposure effects of CZs in Section 5.1 provides suggestive evidence that selection of families into Great Migration CZs is on net modest. However, given the noise in upward mobility estimates based on childhood exposure effects, I cannot rule out some negative selection.

⁴⁵Note that this pooled mobility measure also includes outcomes for Asian Americans, Native Americans, and Hispanic individuals (all other groups do not include those identifying as Hispanic) whereas I focus specifically on the effects on Black versus white families. Given the relatively smaller sizes of these other demographic groups, however, I expect and in fact show, that the results on pooled Black and white mobility look very similar to the results on pooled mobility from Chetty et al. (2020a).

					Pooled B & W:	
	Black	White	Weighted Avg No Comp FX	Weighted Avg + Comp FX	Weighted by Local Shares	Pooled CHJP (2019)
\hat{GM}	-0.0183 (0.00930)	-0.00761 (0.0108)	-0.00835 (0.00625)	-0.0253 (0.0140)	-0.0253 (0.0140)	-0.0257 (0.00961)
Lower Bound	-0.0366	-0.0288	-0.0206	-0.0528	-0.0527	-0.0445
Upper Bound	-0.0000895	0.0136	0.00389	0.00214	0.00203	-0.00686
Black Comp Effect				0.0481		
White Comp Effect				-0.0651		
Ν	129	129	129	129	129	129
Dep var mean	33.19	45.22			38.49	42.12
GM SD	29.56	29.56			29.56	29.56
R-squared	0.286	0.275			0.572	0.417

TABLE D12: COMPARISON OF RACE-SPECIFIC VS. POOLED EFFECTS OF GREAT MIGRATION ON UPWARD MOBILITY

Notes: This table compares the Migration's effect on upward mobility for Black individuals with low income parents; the Migration's effect on white individuals with low income parents; the average of these effects weighted by each racial group's sample average share of low income households in 2000, i.e., ignoring the "composition effect" or the Migration's causal effect on the latter; the coefficient in the previous column plus the composition effect; the effect on the locally weighted average of Black and white individuals from low income families; and the effect on mobility for all racial groups pooled together from Chetty et al. (2020a). The unit of observation is a commuting zone. Dependent variable is expected mean household income rank for individuals with parents at the 25th percentile of the parent income distribution. Income is measured from IRS tax returns for cohorts and parents of cohorts born between 1978 and 1983. Independent variable is the percentile of Black population increase during the Great Migration. The instrument for Black population increase is the percentile of predicted Black population increase, defined as the interaction between pre-1940 Black southern migration patterns and post-1940 outflows of migrants as predicted by southern economic factors alone. Baseline 1940 controls include share of urban population made up of 1935-1940 Black southern migrants, educational upward mobility, share of labor force in manufacturing, and census region fixed effects. Standard errors are in parentheses. Data sources: CCDB; IPUMS complete count 1940 US census; Boustan (2016a); Chetty et al. (2020a).

D.3.3 Heterogeneity by geography

This section explores geographic mediators of the Migration's effect on upward mobility. The deleterious effects of the Great Migration are more pronounced in urban areas of commuting zones than non-urban areas, as shown in Figure D3. The dependent variable on the left side of the panel in the figure is the population-weighted average of upward mobility for non-urban census tracts in a commuting zone only. The dependent variable on the right side of the panel in the figure is the population-weighted average of upward mobility for non-urban census tracts in a commuting zone only. I define a census tract as urban if 100% of the population in that census tract is part of an urban area.⁴⁶ The effect on non-urban upward mobility in the commuting zones in the sample is muted and statistically indistinguishable from zero while the effect on urban upward mobility in commuting zones is close to the baseline effect reported in Table 8 and is statistically significant at the 5% level.

At the same time, sorting into different types of neighborhoods by race within destination CZs does not fully explain the Migration's disparate impact on Black families. Figure D4 reports the effect of the Great Migration on the racial gap in upward mobility in commuting zones versus the average racial gap by census tract within commuting zones. The Great Migration exacerbated gaps in upward mobility between Black and white families in the same census tract. This suggests that Black families and Black boys in particular growing up in the same census tract as white boys and from the same family income experience an effectively different environment. This may be driven, for example, by increased interaction with the criminal justice system or differential access to networks and resources within census tracts that promote upward mobility.

 $^{^{46}}$ In the sample, about 72% are urban according to this definition. Data on urban and rural populations in census tracts are from NHGIS (Manson et al., 2021).

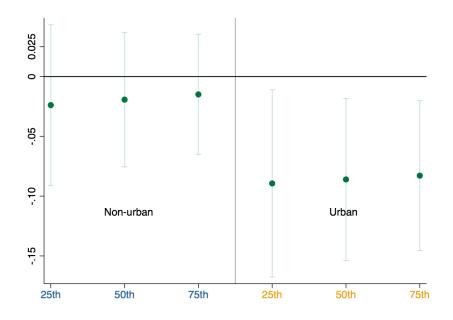


FIGURE D3: IMPACT OF GREAT MIGRATION ON URBAN VS. NON-URBAN UPWARD MOBILITY IN CZS

Notes: This figure plots the 2SLS coefficient on the Great Migration in separate regressions where the dependent variable is upward mobility for different geographies and parent income levels. The dependent variable is the mean household income rank pooling all racial groups and men and women. Income is measured from IRS tax returns for cohorts born between 1978 and 1983. The first three coefficients reflect the Great Migration's impact on individuals in non-urban census tracts with parents at the 25th, 50th, and 75th percentile, weighted by total number of individuals per census tract. The last three coefficients plot the analogous effects for individuals from urban census tracts. Independent variable is the percentile of Black population increase during the Great Migration. Baseline controls included. 95% confidence intervals indicated. *Data sources*: Chetty, Hendren, Jones, and Porter (2018); IPUMS Complete Count 1940 US Census; Boustan (2016a).

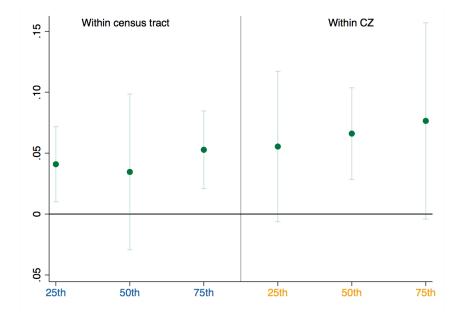


FIGURE D4: IMPACT OF GREAT MIGRATION ON WITHIN-CENSUS-TRACT RACIAL GAP

Notes: This figure plots the 2SLS coefficient on the Great Migration in separate regressions where the dependent variable is the racial gap in upward mobility for different geographies and parent income levels. The dependent variable is the difference in mean household income rank between Black and white individuals, pooling men and women. Income is measured from IRS tax returns for cohorts born between 1978 and 1983. The first three coefficients reflect the Great Migration's impact on the CZ-average within-census-tract racial gap for individuals with parents at the 25th, 50th, and 75th percentile, weighted by total Black plus white population per census tract whose tax returns were used to construct the estimates. The last three coefficients plot the CZ-level gap. Independent variable is the percentile of Black population increase during the Great Migration. Baseline controls included. 95% confidence intervals indicated. *Data sources*: Chetty, Hendren, Jones, and Porter (2018); IPUMS Complete Count 1940 US Census; Boustan (2016a).

D.4 Supplementary information on childhood exposure effects measures

This section provides background information on causal place effect estimates from Chetty and Hendren (2018b). To construct measures of the causal effect of childhood location on adult outcomes, the authors study families that moved across commuting zones as observed from address changes in US federal tax records. Exploiting variation in children's ages at the time different families moved, the authors estimate the effect of an additional year of childhood exposure to a location on children's adult income. Under the assumption that the age of a child at the time a family moved is orthogonal to unobserved family characteristics θ_i , estimating the effect of one additional year of childhood exposure to a location and multiplying this effect by number of years of childhood provides a direct estimate of $\mu_{p,CZ}$ from equation 8. This assumption may not be appropriate when comparing families with school age children to those with younger children as the former may systematically select into locations with better school quality. However, the assumption is far more plausible when making comparisons narrowly across one-year age differences, e.g., 8 year-olds versus 9-year-olds and 14 year-olds versus 15 year-olds.⁴⁷ Chetty and Hendren (2018b) use variation in age of child at time of family's move to purge place effect estimates of bias from sorting on family unobservables, θ_i :

$$y_i = \delta_c + \theta_i$$

$$\downarrow$$

$$\Delta y_i = \alpha_c \Delta t_i$$

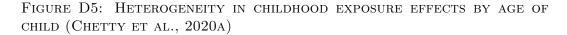
 α_c is an unbiased estimate of the effect of an additional year of childhood exposure to location c on adult outcome y_i .

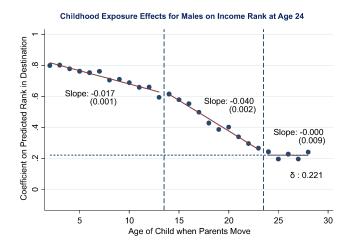
Scaling childhood exposure effects Assuming muted effects for early years according to Figure D5 from Chetty et al. (2020a), the effect of full childhood exposure for 23 years should be adjusted in the following manner⁴⁸:

$$Years = (23 - 13) + (17/40) * 13 = 15.53$$

⁴⁷The specifications the authors use to estimate place effects include origin-destination CZ-pair fixed effects, so comparisons are made across children moving to and from identical locations. Chetty and Hendren (2018a) also provide several checks of the identifying assumption stated above, including instrumenting for moves with displacement shocks to families and the inclusion of family fixed effects.

⁴⁸Deutscher (2020) replicates this finding using data from Australia, finding that on all outcomes, childhood location has its biggest effect during one's teen years.





Notes: This figure from Chetty et al. (2020a) depicts heterogeneity in childhood exposure effects by age of exposure. Early years of childhood exposure have more muted impacts compared to teen years of exposure.

An even more conservative scaling factor might take into account that observational upward mobility estimates reflect the adult outcomes of children born between 1980 and 1988 whose parents remained in the same commuting zone from 1996 to 2012, as measured in their tax records. Thus, the oldest children are age 16 when their parent's location is first measured while the youngest are age 8. I use the information from footnote 17 on p.12 in Chetty and Hendren (2018a) to make extreme, conservative assumptions regarding average years of exposure for the children in the sample. The footnote explains that among families who remain in the same location while their children are between the ages of 16 and 24, "81.5% of them lived in the same CZ when their children were age 8." I first compute average years of exposure at age 16 making the following two assumptions: of the 81.5% of 16 year-olds observed in the same CZ at age 8, none had exposure before age 8 and thus were exposed for a total of 16 years while the other 18.5% had zero years of exposure before the age of 16 and were only exposed from age 16 to 24, or 8 years. This implies an average exposure of 14.52 years for the oldest children in the sample. Using this multiplier for all children in the sample implies that a 1-s.d. increase in

Great Migration inflows is associated with a 3.1-percentile reduction in income rank via location channels alone, compared with a 3.6-percentile reduction in average income rank (including both location and selection channels). Thus, with this assumption, 87% of the Great Migration's impact is explained via location channels and the remaining portion by selection.

D.5 Net effect of the Great Migration

This appendix discusses the overall impact of the Great Migration on Black economic status over the 20th century, through the lens of intergenerational mobility. The main analysis in the paper poses the counterfactual of upward mobility for children in the northern US had they grown up in locations less affected by the Great Migration. This counterfactual does not consider the impact of the Migration on earlier generations, which affects the adult income of Black children today through their parents and grandparents, or on southern stayers, who may have been affected by Black *emigration* from South.

The Great Migration moved Black grandparents dramatically up in the national income distribution: estimates suggest that migrants could approximately double their earnings by moving North (Collins and Wanamaker, 2014; Boustan, 2016a). At the same time, racially segmented labor markets in the North led to increased competition between Black incumbents and new arrivals such that racial earnings convergence in the destinations slowed (Boustan, 2009). Evidence on the timing of changes in conditions in northern cities, presented in Appendix F, suggests that the cohorts growing up in the 1970s would have been exposed to negative environmental factors including extreme segregation, high crime rates, and spillovers from greater greater police presence. Nonetheless, the sharp increase in average grandparent income through migration likely outweighed the competition effect in the North, and potentially even the harsher environment faced by the second generation.

A final relevant factor for understanding the effect of the Migration is the

impact of Black emigration on southern locations. For the Great Migration to have had a net negative impact on Black economic status, it would be necessary to assume that in the absence of Black emigration, southern locations would have been better off. There are two key reasons why this is unlikely to be the case. First, emigration put direct pressure on southern jurisdictions to offer better amenities for Black workers. Boustan and Tabellini (2018) find that votes for segregationist policies decreased in places where Black migrants left in greater numbers. This echoes the "voting with one's feet" hypothesis explored by Margo (1991). Second, Calderon et al. (2019) find that the Great Migration may have played a role in bringing civil rights issues to the national stage and helped civil rights legislation get passed. The effects of civil rights legislation were felt more strongly in the South than in the North, so this suggests another mechanism through which the Migration may have improved southern conditions.

In a simplified counterfactual exercise, I explore the aggregate effect of the Great Migration while making several conservative assumptions. First, I assume a zero effect of the Great Migration on the South and that the net effect of the Great Migration on parent income (inclusive of the effects on grandparents) is reflected in the difference in average Black parent income rank in the North and the South today. In the absence of the Great Migration, 23% of Black grandchildren would experience the counterfactual northern intergenerational mobility curve had the Migration not occurred while 77% would remain on the southern curve. This exercise suggests a positive net effect of the Great Migration on Black income of 0.2 income percentiles.

I conclude that while the Migration eventually reduced the gains to parent income for Black children in the North, the large positive effect the Migration had on the income of earlier generations (moving Black children *up* the IGM curve) makes up for these losses. Any additional positive impacts on the South would only magnify a positive net effect of the Great Migration.

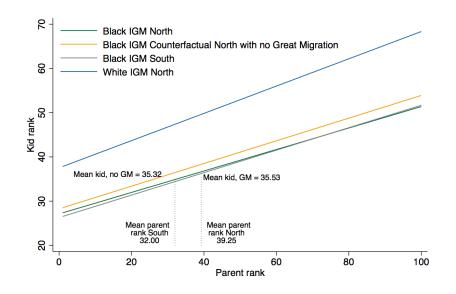


FIGURE D6: INTERGEN. MOBILITY BY RACE AND REGION

Notes: This figure plots intergenerational mobility curves by race and region. The y-axis plots the income rank of individuals from the 1978-1983 birth cohorts and the x-axis plots the income rank of their parents. Income is measured from IRS tax returns. The green line plots the intergenerational mobility curve for Black families in the North; the gray line plots the intergenerational mobility curve for Black families in the South; and the gold line plots a counterfactual intergenerational mobility curve for Black families in the North in the absence of the Great Migration. Average parent income rank in the North and South are indicated on the plot. The counterfactual line is plotted using estimates of the Migration's impact on Black men from the 25th, 50th, and 75th percentiles of the parent income distribution from regressions described in Section 4. Baseline 1940 controls include share of urban population made up of 1935-1940 Black southern migrants, educational upward mobility, share of labor force in manufacturing, and census region fixed effects. *Data sources*: CCDB; IPUMS complete count 1940 US census; Boustan (2016a); Chetty et al. (2020a).

D.6 Additional details on robustness checks in Section 5.3

This section provides further details on the robustness checks discussed in Section 5.3.

Deindustrialization As discussed in the main text, I control for the manufacturing share of the labor force in 1940 to assess deindustrialization as a potential confound for the Great Migration's impact on upward mobility. The 1940 manufacturing share is strongly predictive of the manufacturing share in following decades. Specifically, The correlation between 1950 share of the labor force in manufacturing and the baseline period share is 0.97. By 1970, this drops only slightly, to 0.85. To further address this concern, I control for a Bartik demand shock from 1940-1970. I construct this demand shock by interacting industry shares with national, leave-one-out changes in manufacturing employment between 1940 and 1970, using data from Ruggles et al. (2021) and Manson et al. (2021).

Other migrations To control for an exogenous measure of historical European migration and assess whether or not this confounds my estimates, I use data from Sequeira et al. (2019) that contains instruments for historical European migration based on whether a county was connected to railways during migration booms versus busts during the Age of Mass Migration. I also examine white southern migration over 1940-1970 as a potential confound by developing a shift share instrument for this type of migration, that combines 1935-1940 white southern migrant settlement patterns with white southern county outmigration. As discussed in the main text, white southern migration does not lead to changes in the gains from growing up in destination CZs today. Interestingly, white southern migration appears associated with lower outcomes for white men and women from lower income parents. The lack of an effect on childhood exposure effects suggests that the channel is the composition of the average white child as opposed to changes in local public goods or neighborhood quality in response to historical in-migration of white southerners.

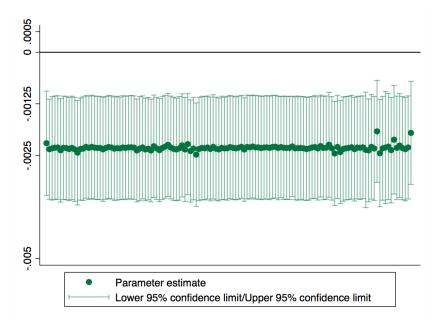
D.7 Additional robustness checks

This section reports additional robustness checks on the core result that the Great Migration reduced gains from growing up in northern commuting zones in the US. All of the results report the reduced form relationship between the instrument for the Great Migration and upward mobility. I briefly discuss each of the results in turn.

D.7.1 Robustness to dropping each CZ once

To insure that my results are not driven by any particular commuting zone, I rerun the analysis of the Great Migration's impact on childhood exposure effects for low income families dropping one CZ at a time. The results are shown in Figure D7. The coefficient on the Great Migration is highly consistent across all 130 regressions dropping one CZ from the sample each time, indicating that no single CZ drives the relationship between the Great Migration and upward mobility.

FIGURE D7: GREAT MIGRATION EFFECT ROBUST TO LEAVING OUT EACH CZ ONCE FROM SAMPLE



Notes: This figure plots the coefficient on percentile of predicted Black population change in 130 separate regressions where each CZ in the sample has been left out of the regression once. 95% confidence intervals indicated. The unit of observation is a commuting zone. Dependent variable is the estimated causal impact of one additional year of childhood in the commuting zone on adult household income rank for men and women with parents at the 25th percentile of the parent income distribution. Income is measured from IRS tax returns for cohorts and parents of cohorts born between 1980 and 1986. Baseline 1940 controls include share of urban population made up of 1935-1940 Black southern migrants, educational upward mobility, share of labor force in manufacturing, and census region fixed effects. *Data sources*: CCDB; IPUMS complete count 1940 US census; Boustan (2016a); Chetty and Hendren (2018b).

D.7.2 Controls for CZ urbanicity

The next set of robustness checks explore robustness to baseline population characteristics of the commuting zones in the sample, including the share of the population that was urban in 1940 and log population density. Figures D8a and D9b report the results. Controlling for the CZ's urban share in 1940 or log population density does not affect the results suggesting that mere urbanicity (and the trajectory of upward mobility in urban areas) is not driving the Great Migration's impact on upward mobility in the CZ.

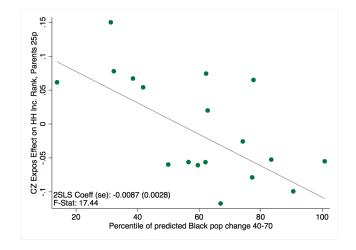
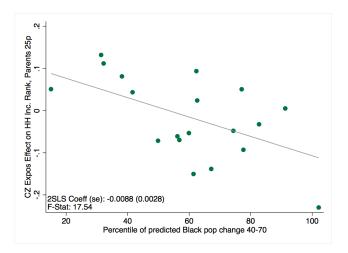


FIGURE D8: GREAT MIGRATION IMPACT ON CHILDHOOD EXPOSURE EFFECTS, POPULATION CONTROLS

(a) Controlling for CZ urban share in 1940





Notes: Panel (a) depicts a binned scatterplot of the reduced form relationship between the instrument for the Great Migration and CZ childhood exposure effects for individuals from low income families, controlling for the urban population share in the CZ in 1940. Panel (b) depicts the same but with log population density in 1940 in the CZ as a control. The right hand side variable is grouped into 20 bins (5 percentiles each). The unit of observation is a commuting zone. Dependent variable is the estimated causal impact of one additional year of childhood in the commuting zone on adult household income rank for men and women with parents at the 25th percentile of the parent income distribution. Income is measured from IRS tax returns for cohorts and parents of cohorts born between 1980 and 1986. The Great Migration instrument is the percentile of predicted Black population increase, defined as the interaction between pre-1940 Black southern migration patterns and post-1940 outflows of migrants as predicted by southern economic factors alone. Baseline 1940 controls include share of urban population made up of 1935-1940 Black southern migrants, educational upward mobility, share of labor force in manufacturing, and census region fixed effects. The 2SLS coefficient and the first-stage F-statistic are also reported. *Data sources*: CCDB; IPUMS complete count 1940 US census; Boustan (2016a); Chetty and Hendren (2018b).

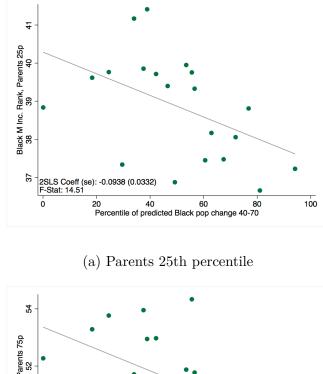
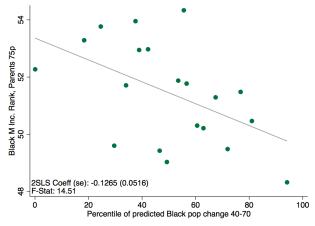


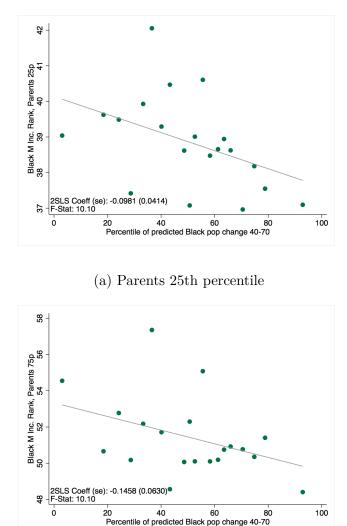
FIGURE D9: GREAT MIGRATION IMPACT ON BLACK MEN'S UPWARD MOBILITY, CONTROLS FOR URBAN SHARE



(b) Parents 75th percentile

Notes: Panel (a) depicts a binned scatterplot of the reduced form relationship between the instrument for the Great Migration and Black men's upward mobility (25th percentile of parent income distribution), controlling for the urban population share in the CZ in 1940. Panel (b) depicts the same for Black men from the 75th percentile of parent income distribution. The right hand side variable is grouped into 20 bins (5 percentiles each). The unit of observation is a commuting zone. Upward mobility is defined as mean individual or household income rank by childhood commuting zone where income is measured from IRS tax returns for cohorts born between 1978 and 1983. The Great Migration instrument is the percentile of predicted Black population increase, defined as the interaction between pre-1940 Black southern migration patterns and post-1940 outflows of migrants as predicted by southern economic factors alone. Baseline 1940 controls include share of urban population made up of 1935-1940 Black southern migrants, educational upward mobility, share of labor force in manufacturing, and census region fixed effects. The 2SLS coefficient and the first-stage F-statistic are also reported. *Data sources*: CCDB; IPUMS complete count 1940 US census; Boustan (2016a); Chetty et al. (2020a).

FIGURE D10: GREAT MIGRATION IMPACT ON BLACK MEN'S UPWARD MOBILITY, CONTROLS FOR 1940 LOG POPULATION DENSITY

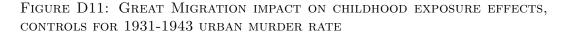


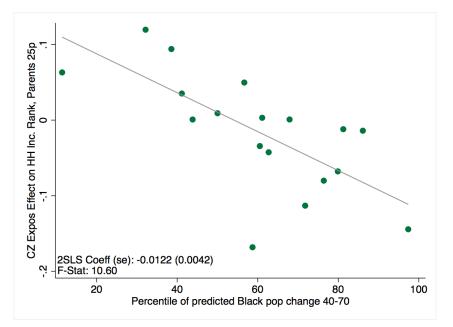
(b) Parents 75th percentile

Notes: Panel (a) depicts a binned scatterplot of the reduced form relationship between the instrument for the Great Migration and Black men's upward mobility (25th percentile of parent income distribution), controlling for log population density in 1940 in the CZ in 1940. Panel (b) depicts the same for Black men from the 75th percentile of parent income distribution. The right hand side variable is grouped into 20 bins (5 percentiles each). The unit of observation is a commuting zone. Upward mobility is defined as mean individual or household income rank by childhood commuting zone where income is measured from IRS tax returns for cohorts born between 1978 and 1983. The Great Migration instrument is the percentile of predicted Black population increase, defined as the interaction between pre-1940 Black southern migration patterns and post-1940 outflows of migrants as predicted by southern economic factors alone. Baseline 1940 controls include share of urban population made up of 1935-1940 Black southern migrants, educational upward mobility, share of labor force in manufacturing, and census region fixed effects. The 2SLS coefficient and the first-stage F-statistic are also reported. *Data sources*: CCDB; IPUMS complete count 1940 US census; Boustan (2016a); Chetty et al. (2020a).

D.7.3 Robustness to inclusion of pre-1940 murder rate control

Figure 9 shows that the Black in-migration into cities between 1940 and 1970 was correlated with the urban murder rate prior to 1940. However, controlling for pre-1940 murder rates does not alter the relationship between the Migration and post-period police spending, murder rates, or incarceration. In this section, I examine whether my main results on upward mobility are also robust to controlling for pre-1940 murder rates (average of 1931 and 1943 murder rates). Figures D11 and Figures D12a and D13b show binned scatter plots of the reduced form relationship between the Great Migration shock and childhood exposure effects of CZs for low income families and upward mobility for Black men from low and high income families. Results are robust to controlling for pre-1940 urban murder rates.





Notes: The figure depicts a binned scatterplot of the reduced form relationship between the instrument for the Great Migration and CZ childhood exposure effects for individuals from low income families, controlling for the average urban murder rate in the CZ from 1931 to 1943. Dependent variable is the estimated causal impact of one additional year of childhood in the commuting zone on adult household income rank for men and women with parents at the 25th percentile of the parent income distribution. Income is measured from IRS tax returns for cohorts and parents of cohorts born between 1980 and 1986. The right hand side variable is grouped into 20 bins (5 percentiles each). The unit of observation is a commuting zone. The Great Migration instrument is the percentile of predicted Black population increase, defined as the interaction between pre-1940 Black southern migration patterns and post-1940 outflows of migrants as predicted by southern economic factors alone. Baseline 1940 controls include share of urban population made up of 1935-1940 Black southern migrants, educational upward mobility, share of labor force in manufacturing, and census region fixed effects. The 2SLS coefficient and the first-stage F-statistic are also reported. *Data sources*: CCDB; IPUMS complete count 1940 US census; Boustan (2016a); see Appendix E for the full list of data sources on each of the mechanisms.

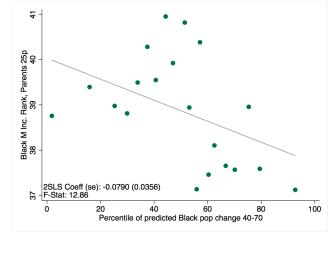
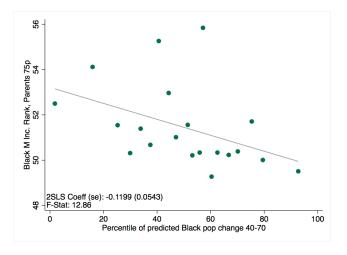


FIGURE D12: GREAT MIGRATION IMPACT ON BLACK MEN, CONTROLS FOR 1931-1943 URBAN MURDER RATE

(a) Parents 25th percentile



(b) Parents 75th percentile

Notes: Panel (a) depicts a binned scatterplot of the reduced form relationship between the instrument for the Great Migration and Black men's upward mobility (25th percentile of parent income distribution), controlling for the average urban murder rate in the CZ from 1931 to 1943. Panel (b) depicts the same for Black men from the 75th percentile of parent income distribution. The unit of observation is a commuting zone. Upward mobility is defined as mean individual or household income rank by childhood commuting zone where income is measured from IRS tax returns for cohorts born between 1978 and 1983. The right-hand side variable is the Great Migration instrument: the percentile of predicted Black population increase, defined as the interaction between pre-1940 Black southern migration patterns and post-1940 outflows of migrants as predicted by southern economic factors alone. The right hand side variable is grouped into 20 bins (5 percentiles each). Baseline 1940 controls include share of urban population made up of 1935-1940 Black southern migrants, educational upward mobility, share of labor force in manufacturing, and census region fixed effects. The 2SLS coefficient and the first-stage F-statistic are also reported. *Data sources:* Chetty, Hendren, Jones, and Porter (2018); IPUMS Complete Count 1940 US Census; Boustan (2016).

D.7.4 Placebo check using white southern migration

I next regress upward mobility (childhood exposure effects and Black men's upward mobility) on an instrument for white southern migration into nonsouthern cities as a placebo check.⁴⁹ The results are reported in Figures D13, D14a, and D15b. There is no relationship between white southern migration between 1940 and 1970 and the returns to growing up in northern locations. The relationship between historical white southern migration and Black men's upward mobility is statistically insignificant and slightly positive.

These results indicate that the Migration's impact is not simply reflecting increases in the southern migrant population more generally, but rather that racial composition shocks through Black migration are what altered northern locations and reduced upward mobility for low income families.

 $^{^{49}}$ Data on white southern migration come from Gardner and Cohen (1992) and Bowles et al. (2016).

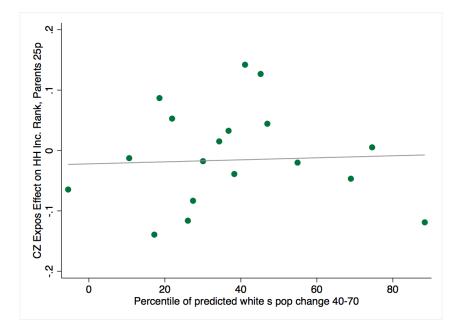


FIGURE D13: WHITE SOUTHERN MIGRATION IMPACT ON CHILDHOOD EXPOSURE EFFECTS

Notes: This figure depicts a binned scatterplot of the relationship between the percentile of predicted white southern in-migration and CZ childhood exposure effects for individuals from low income families. The unit of observation is a commuting zone. Dependent variable is the estimated causal impact of one additional year of childhood in the commuting zone on adult household income rank for men and women with parents at the 25th percentile of the parent income distribution. Income is measured from IRS tax returns for cohorts and parents of cohorts born between 1980 and 1986. The right-hand side variable is the instrument for white southern migration: the percentile of predicted white population increase, defined as the interaction between pre-1940 white southern migration patterns and post-1940 outflows of white migrants. The right hand side variable is grouped into 20 bins (5 percentiles each). Baseline 1940 controls include share of urban population made up of 1935-1940 Black southern migrants, educational upward mobility, share of labor force in manufacturing, and census region fixed effects. *Data sources*: Chetty, Hendren, Jones, and Porter (2018); IPUMS Complete Count 1940 US Census; Boustan (2016).

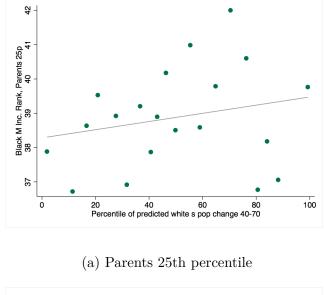


FIGURE D14: WHITE SOUTHERN MIGRATION IMPACT ON BLACK MEN'S UPWARD MOBILITY

56 Black M Inc. Rank, Parents 75p 50 52 54 48 100 60 ò 20 40 80 Percentile of predicted white s pop change 40-70

(b) Parents 75th percentile

Notes: Panel (a) depicts a binned scatterplot of the relationship between the percentile of predicted white southern in-migration and Black men's upward mobility (25th percentile of parent income distribution). Panel (b) depicts the same for Black men from the 75th percentile of parent income distribution. The unit of observation is a commuting zone. Units of outcome variables are standard deviations. Upward mobility is defined as mean individual or household income rank by childhood commuting zone where income is measured from IRS tax returns for cohorts born between 1978 and 1983. The right-hand side variable is the instrument for white southern migration: the percentile of predicted white population increase, defined as the interaction between pre-1940 white southern migration patterns and post-1940 outflows of white migrants. The right hand side variable is grouped into 20 bins (5 percentiles each). Baseline 1940 controls include share of urban population made up of 1935-1940 Black southern migrants, educational upward mobility, share of labor force in manufacturing, and census region fixed effects. Baseline 1940 controls include share of urban population made up of 1935-1940 Black southern migrants, educational upward mobility, share of labor force in manufacturing, and census region fixed effects. Data sources: Chetty, Hendren, Jones, and Porter (2018); IPUMS Complete Count 1940 US Census; Boustan (2016).

D.7.5 Impact of the Great Migration on change in Black men's upward mobility

Results on Black men's upward mobility are robust to examining the effect of the Migration on the long-run change in Black men's upward mobility within CZs, assuaging the concern that unobserved fixed characteristics of CZs confound the effect of the Great Migration on upward mobility.

To construct the long difference in Black men's upward mobility, I take the difference in the Z-score of Black men's income upward mobility in 2015 (for men from parents at the median of the national parent income distribution) and the Z-score of Black boys' educational upward mobility in 1940 (for boys whose parents had 5-8 years of schooling, the national median for adults). I then standardize this difference, so that the units of outcome variables are standard deviations.

Figure D15 shows a binned scatter plot of the reduced form relationship between the instrument for the Great Migration and the change in Black men's upward mobility. The 2SLS coefficient and first-stage F-statistic are also reported in the figure. Consistent with the baseline specifications, the results show the Migration is associated with reductions in Black men's upward mobility between 1940 and 2015.

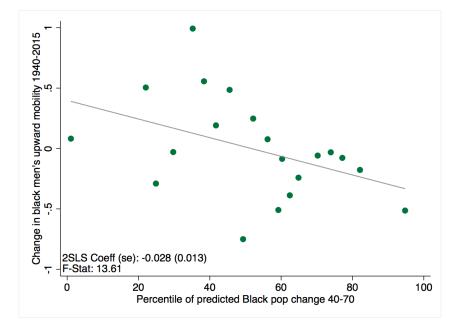


FIGURE D15: GREAT MIGRATION IMPACT ON CHANGE IN BLACK MEN'S UPWARD MOBILITY, 1940-2015

Notes: This figure depicts a binned scatterplot of the reduced form relationship between the instrument for the Great Migration and the change in Black men's upward mobility between 1940 and 2015. The unit of observation is a commuting zone. Upward mobility in 1940 is the fraction of children from median-educated households with more years of education than their parents. Upward mobility in the 2000s is defined as mean individual or household income rank by childhood commuting zone where income is measured from IRS tax returns for cohorts born between 1978 and 1983. Units of the outcome variable are standard deviations. The right-hand side variable is the Great Migration instrument: the percentile of predicted Black population increase, defined as the interaction between pre-1940 Black southern migration patterns and post-1940 outflows of migrants as predicted by southern economic factors alone. The right hand side variable is grouped into 20 bins (5 percentiles each). Baseline 1940 controls include share of urban population made up of 1935-1940 Black southern migrants, the share of labor force in manufacturing, and census region fixed effects. The 2SLS coefficient and the first-stage F-statistic are also reported. *Data sources*: Chetty, Hendren, Jones, and Porter (2018); IPUMS Complete Count 1940 US Census; Boustan (2016).

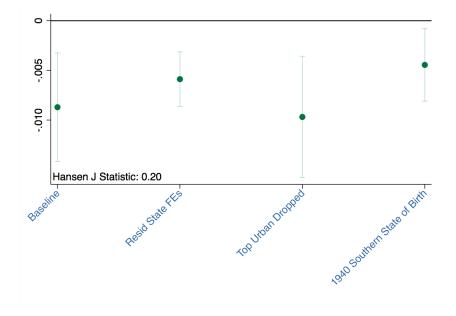
D.7.6 Robustness to using alternative instruments

Figure D16 shows robustness of the main result to the use of alternative instruments intended to address the concern of correlated shocks to southern origin counties and northern destination cities. The first coefficient shown in the figure is the coefficient on the baseline instrument for the Great Migration, the interaction between pre-1940 Black southern migration patterns and post-1940 southern county net-migration as predicted by local economic

factors alone. The second point estimate represents the the coefficient on an instrument formed using southern county net-migration first residualized on southern state fixed effects to address correlated shocks to northern destinations and southern states. The third point estimate reports the coefficient on the instrument formed by first dropping the 15 southern counties coded as central in MSAs containing a population of 1 million or more in $1990.^{50}$ This instrument takes into account potential correlated shocks to major urban areas in the South and northern cities. Version 4 of the instrument leverages an alternative set of shocks by using state of birth of southern-born Black individuals in the North as of 1940 to link southern shocks to northern cities. Predicted southern county outflows between 1940 and 1970 are aggregated to the state level and assigned to northern cities using the state origins of southern-born Black residents in northern cities. This last instrument, by exposing northern cities to an alternative set of shocks, takes into account potential correlated shocks to the original set of counties in the other forms of the instrument and northern destination cities. As the figure shows, results are extremely similar across the different types of instruments. A formal over-identification test yields a Hansen J statistic of 0.20; thus, I fail to reject the null that the estimated effects are statistically indistinguishable.

 $^{^{50}}$ Data on urbanicity of US counties come from CDC (2021).

FIGURE D16: ALTERNATIVE INSTRUMENTS FOR THE GREAT MIGRATION



Notes: This figure plots the 2SLS coefficient on the Great Migration using alternative instruments for historical Black population change, where the dependent variable is commuting zone childhood exposure effects in the 2000s for men and women with low income parents. The unit of observation is a commuting zone. Childhood exposure effects are the estimated causal impact of one additional year of childhood in the commuting zone on adult household income rank for men and women with parents at the 25th percentile of the parent income distribution. Income is measured from IRS tax returns for cohorts and parents of cohorts born between 1980 and 1986. The instrument is the percentile of predicted Black population increase, defined in versions 1-3 as the interaction between pre-1940 Black southern migration patterns and 1) post-1940 southern county net-migration as predicted by local economic factors alone; 2) southern county net-migration residualized on state fixed effects; and 3) southern county net-migration from less urban counties (dropping the counties coded as central in MSAs with populations of 1 million or more in 1990 - 15 counties total). In version 4, predicted southern county outflows between 1940 and 1970 are aggregated to the state level and assigned to northern cities according to the share of the Black population born in that southern state and living in the destination city in 1940. Baseline 1940 controls include share of urban population made up of 1935-1940 Black southern migrants, educational upward mobility, share of labor force in manufacturing, and census region fixed effects. Data sources: Chetty, Hendren, Jones, and Porter (2018); IPUMS Complete Count 1940 US Census; Boustan (2016).

D.7.7 Alternative inference on Great Migration shift-share instrument

Adão et al. (2019) show that standard inference procedures result in standard errors that may be too small in the case of shift-share instruments due to correlated errors across observations that are similarly exposed to the same set of shocks. To assess whether my results are driven by noise, I follow Adão et al. (2019) and interact shares with shocks drawn from a random normal distribution. For simplicity, I use the same mean and variance as Adão et al. (2019): 0 and 5, respectively. Because I use a rank transformation of Black migration, the ranking of migration inflows is invariant to the specific variance chosen.

I iterate this procedure 1,000 times and document the fraction of times results show significant effects at the 5% and 1% level. Figure D17 reports the results of this analysis. I show that the coefficients on the resulting placebo instruments are significant in either the positive or negative direction 16.1% of the time at the 5% level, compared to 55% of the time in the application discussed in Adão et al. (2019). Furthermore, the coefficients are significant in the negative direction just 6.1% of the time at the 1% level. The results from this placebo analysis suggest that while the standard errors likely warrant adjusting, the impact of the Great Migration on upward mobility is unlikely to be driven by noise and would remain highly significant.

The robust inference procedure that generates alternative p-values from Adão et al. (2019) applies only to linear shift-share instruments and is thus not applicable in the context of this paper, which uses a nonlinear transformation of the standard shift share (percentile ranks of historical Black in-migration). A new working paper by Borusyak and Hull (2020), however, provides a robust inference procedure that generalizes to the case of nonlinear shift-share instruments, and which is based on randomization inference. Following their procedure, I permute observed county-decade net-migration across southern counties and interact these permuted shocks with migration shares to generate counterfactual shocks to northern locations. I repeat this procedure 1,000 times and implement the two-sided significance test of Borusyak and Hull (2020). Using this procedure, I obtain a p-value of 0.054. Thus, with this alternative inference procedure, my results are robust at the 10% level.

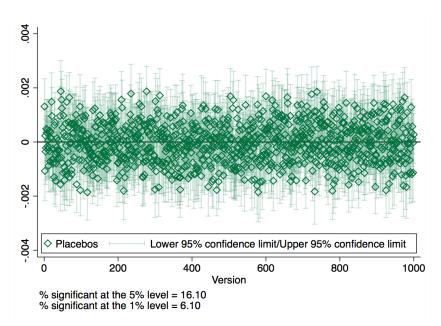


FIGURE D17: PLACEBO MIGRATION SHOCKS

Notes: This figure plots the coefficient on placebo shocks in 1,000 separate regressions, where the dependent variable is commuting zone childhood exposure effects in the 2000s for men and women with low income parents. The unit of observation is a commuting zone. Childhood exposure effects are the estimated causal impact of one additional year of childhood in the commuting zone on adult household income rank for men and women with parents at the 25th percentile of the parent income distribution. Income is measured from IRS tax returns for cohorts and parents of cohorts born between 1980 and 1986. The placebo shock is the percentile placebo increase in the Black population, defined as the interaction between pre-1940 Black southern migration patterns and a normally distributed random variable with mean 0 and variance 5. Baseline 1940 controls include share of urban population made up of 1935-1940 Black southern migrants, educational upward mobility, share of labor force in manufacturing, and census region fixed effects. *Data sources:* Chetty, Hendren, Jones, and Porter (2018); IPUMS Complete Count 1940 US Census; Boustan (2016).

Appendix E Public Finance and Neighborhoods Database, 1920-2015

This section explains the construction of the new database of local public finance and neighborhood quality measures I assembled for commuting zones spanning the years 1920-2015. The database covers statistics on schooling, demographics, racial tension and voting behavior, local government expenditures, incarceration, and crime, among other characteristics. I digitized and harmonized data from a variety of sources. Below I describe the data sources and construction of measures from each category of local public finance and neighborhood quality covered in the database.

Private school enrollment rates

Data on private school enrollments come from two different sources depending on the time period. For pre-1940 statistics on private school enrollment, I digitized tabulations on city school systems from the 1922 Biennial Survey of Education report (U.S. Office of Education, Department of the Interior, 1924). This report contains the total number of elementary and high school students enrolled in private schools in that city as well as total school enrollment in the city.

For 1970 onwards, I use county-level counts of private school enrollments from IPUMS NHGIS (Manson et al., 2021), which I aggregate up to the CZ level. Starting in 1970 through 2010, enrollment is also reported separately for elementary and high school students and separately by race from 1970 to 2000.

Incarceration rates

For 1920 and 1930, I use the complete count censuses (Ruggles et al., 2021) to construct the percent of the population in a county that is incarcerated in jails or local correctional institutions. I do not include inmates in federal or state prison in these estimates as it is not possible to allocate state and federal prisoners back to localities they came from. For 1940, I digitized data from a census report on the incarcerated population (U.S. Census Bureau, 1943). For 1960, I digitized data from the published 1960 US Decennial Census, which includes a table on the incarcerated population and reports the non-white and white incarcerated population by county separately (U.S. Census Bureau, 1963).

For the post-1970 period, I use a rich new dataset from the Vera Institute of Justice In Our Backyards Symposium (IOB), which provides counts of federal and state prisoners by their county-of-commitment to federal and state prison (Kang-Brown et al., 2020). These data begin in the year 1983. These figures are available separately by race. Due to reliability issues for the local jail population in these data, I focus on total jail rates rather than jail population breakdowns by race.

Crime rates

For crime rates, I focus on murder rates as these are less subject to reporting bias than other crime categories, such as property crime or non-fatal violent crimes. I digitize murder rates for cities with a population of 25,000 or more from the Uniform Crime Reports (UCR) series of the FBI in 1931, 1943, and 1950 (U.S. Department of Justice, FBI, 1950).⁵¹ For the years 1958 to 1969, I use city-level tabulations of murder rates from UCR available from ICPSR (ICPSR, 2005). Finally, for the post 1970 period, I use county-level tabulations of UCR murder rates available from the IOB database (Kang-Brown et al., 2020).

In addition to looking at crime rates as a measure of neighborhood quality, I also use data on the intensity and duration of race riots in major cities in the 1960s.⁵²

⁵¹Some large cities did not report to the FBI UCR series in these years. A notable case is New York City in 1931 and in 1950. For these cities in 1931, I supplement using data generously shared by Price Fishback (Fishback et al., 2010). I drop 1950 from the analysis due to missing data from New York City.

⁵²These data were generously shared by William Collins and Robert Margo (Collins and Margo, 2007a) and are based on the work of based on the work of Carter (1986).

Local government expenditures

Data on local government expenditures come from surveys of state and substate level governments conducted by the US Census Bureau.

The first full set of such data are available in the 1932 publication of *Finan*cial statistics of state and local governments (U.S. Census Bureau, Department of Commerce, 1933). I digitized county aggregate and individual local government expenditures from this report. Individual Government Finances data (U.S. Census Bureau, c); County Government Finances data (U.S. Census Bureau, b); and City Government Employment data (U.S. Census Bureau, a).

For post-migration years, I use individual government expenditure data in digital format for roughly 15,000 local governments across the United States from 1967 to 2012 (U.S. Census Bureau, c). I also include data on city government expenditures available for intermittent years from 1948 to 1975 from U.S. Census Bureau, Department of Commerce (2008) and county aggregates of expenditures on different categories from U.S. Census Bureau, Department of Commerce (2012) and U.S. Census Bureau (b). In the case of police expenditures, I supplement these two measures with counts of police officers per capita using the complete count censuses available from IPUMS for the years 1920, 1930, and 1940 and US Census Bureau data surveying public sector employment in cities from 1951-2007 (U.S. Census Bureau, a).

For each dataset, I construct commuting zone area aggregate expenditures by all local governments for the expenditure categories of interest. The advantage to this approach is that changes in which levels or types of government are responsible for providing a certain public good will not affect this measure of spending. I focus on expenditures per capita (or per student) and the share of total expenditures devoted to that expenditure category.

For example, for police spending, CZ-area local government expenditure

share is defined as

Pol. Exp. Share_{CZ} =
$$\frac{\text{Spent on Police by All Local Governments}_{CZ}}{\text{Spent by All Local Governments}_{CZ}}$$

and per capita expenditures at the CZ-area level are defined as

Per Cap Pol. Exp._{CZ} =
$$\frac{\text{\$Spent on Police by All Local Governments}}{\text{Population}_{CZ}}$$

Finally, I focus on categories of expenditures over which local governments have a large degree of discretion: police expenditures, education expenditures, and fire expenditures. Table E1 shows the the contribution of different levels of government (e.g., federal, state, county, etc.) to direct expenditures for each category of government spending.

Govt Type Fed	Revenue 18.93%	$\mathop{\rm Rev}\limits_{24.90\%} ({\rm Own})$	Dir Exp 22.95%	$\frac{\text{Elem} + \text{HS}}{0.00\%}$	Fire Prot 0.00%	$\begin{array}{l} {\rm Health} + {\rm Hosp} \\ 6.12\% \end{array}$	Highway 0.36%	Parks & Rec 3.24%	Police 3.73%	Pub. Welf 6.27%	Sanitation 0.82%
\mathbf{State}			34.17%	1.14%	0.00%	42.77%	60.37%	14.44%		79.42%	5.13%
County			9.51%	7.84%	13.83%	26.71%	14.76%	16.76%		10.33%	16.03%
Muni			13.90%	8.31%	67.90%	10.64%	19.41%	51.29%		3.78%	56.82%
Town			1.31%	2.24%	6.29%	0.25%	3.79%	3.75%		0.09%	5.40%
Spec. Dist.			4.79%	0.24%	11.99%	13.51%	1.32%	10.52%		0.11%	15.81%
School Dist.			13.38%	80.23%	0.00%	0.00%	0.00%	0.00%		0.00%	0.00%
Total		100.00%	100.00%	100.00%	100.00%	100.00%	100.00%	100.00%	100.00%	100.00%	100.00%
Notes: This table	shows the bread	kdown in spending	by government	type for different s	spending catego	Notes: This table shows the breakdown in spending by government type for different spending categories, averaged across all reporting years from 1967-2012. Column 1 lists the government types: federal, state,	reporting years	from 1967-2012. Col	lumn 1 lists th	e government typ	ss: federal, state,

TABLE E1: EXPENDITURE BY GOVERNMENT TYPE BY SPENDING CATEGORY

county, municipality, town, special district, and school district. Column 2 shows the total breakdown of government revenue by government type. Column 3 shows this number net of intergovernmental transfers. Column 4 shows total direct government expenditures by government type. Starting with Column 5, the categories of spending from left to right are education for elementary and high school districts; fire protection services: health and hospitals; parks and recreation; public welfare; and sanitation. Sanitation spending includes sewage and waste management. *Source*. US Census Bureau Individual Government Finances data (1967-2012).

Appendix F Additional results on local mechanisms

F.1 Impact on private schooling and residential segregation

In this section, I report additional results on private schooling and residential segregation. Figure F1 plots the coefficients on predicted Black population increases on standardized measures of private school enrollment rates separately for each year that data are available. The outcome variable is the share of elementary and high school students enrolled in private school. Beginning in 1970, these measures are available separately by race. I find no impact of the 1940-1970 Migration shock on private school enrollment rates in 1920. In 1970, the next year that data are available,⁵³ a 1-standard-deviation increase in the Great Migration shock is associated with approximately a 0.2 standard deviation increase in white private school enrollment rates by 1990. Individually, these results are not statistically significantly different from zero. However, the black-white gap in public school enrollment is significantly larger in Great Migration CZs.

 $^{^{53}}$ Starting in 1960, the Census began asking about the type of school households enrolled their children in; however, aggregate statistics for children attending high school as well as breakdowns by race are only available through NHGIS until 1970. See Appendix E for more details.

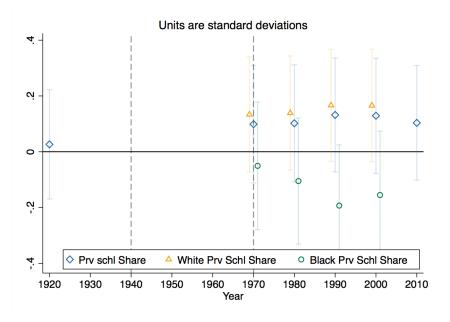


FIGURE F1: GREAT MIGRATION IMPACT ON PRIVATE SCHOOLING

Notes: This figure plots the coefficient on the instrument for Black population increases during the Great Migration, in approximately one standard deviation units, in separate regressions for each year where the dependent variable is private school enrollment rates. The unit of observation is a commuting zone. Units of outcome variables are standard deviations. The instrument is the percentile of predicted Black population increase, defined as the interaction between pre-1940 Black southern migration patterns and post-1940 outflows of migrants as predicted by southern economic factors alone. A one standard deviation increase is approximately 30 percentiles. Baseline 1940 controls include share of urban population made up of 1935-1940 Black southern migrants, educational upward mobility, share of labor force in manufacturing, and census region fixed effects. *Data sources*: Biennial Statistics of Education, 1920-1922; NHGIS county-level aggregates of elementary and high school enrollment by school type (public or private), 1970-2010; IPUMS complete count 1940 US census; Boustan (2016a).

Consistent with Boustan (2010) and Tabellini (2019), I find that Black population increases also predict large declines in the urban white share at the commuting zone level. These results are shown in Figure F2.

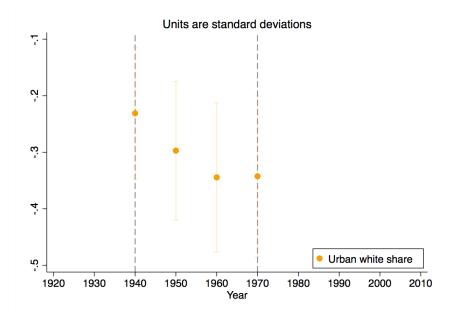


FIGURE F2: GREAT MIGRATION IMPACT ON URBAN WHITE SHARE

Notes: This figure plots the coefficient on a 1 s.d. increase in predicted Black population change in separate regressions for each year where the dependent variable is the urban white population share. The unit of observation is a commuting zone. Units of outcome variables are standard deviations. Baseline 1940 controls include share of urban population made up of 1935-1940 Black southern migrants, educational upward mobility, share of labor force in manufacturing, and census region fixed effects. *Data sources*: County Data Books 1947-1977.

F.2 Impact on local government expenditures

Next, I examine the impact of the Migration on the public spending patterns of local governments. I focus on categories of public expenditures over which sub-state governments have a large degree of discretion. Appendix Table E1 shows the contribution of different levels of government to each of several main categories of public expenditures. I focus on two categories in particular, police and school expenditures. Spending on police indicates levels of neighborhood safety and crime, but also may have direct effects on the outcomes of Black male youth in particular, which I discuss further in the main text (see Section 6). School spending has natural implications for the average outcomes of children in a given location. Figure F3 plots the coefficients on predicted Black population increases on standardized measures of police investments separately for each year that the data are available. The outcome variables are police expenditures per capita, the share of local government expenditures on police, and police officers per capita. As can be seen in the Figure, the Migration from 1940-1970 had no statistically significant or large effects on police investments from 1920-1940. Starting after 1940, the association between the Migration and police spending increases, peaking in the late 1970s and persisting for several decades after. At the peak of the association between the Migration and police investments, a 1 standard-deviation increase in the Migration shock increased the police expenditure share and police expenditure per capita by just over 0.2 standard deviations.

I then look at the impact of the Great Migration on educational investments in affected commuting zones. These investments include direct educational expenditures by school districts, both as a share of all local government expenditures in commuting zones and per pupil. Figures F4 and F5 report these results. I estimate a noisy negative association between the Migration on pre-1940 (1932) aggregate educational expenditure shares. In F5, I control for 1932 educational expenditure shares and estimate the Migration's impact on post-1970 educational investments. I find no impact of the Migration on aggregate education expenditures at the commuting zone level in the post-Migration period. I discuss the implications of these findings in the main text (see Section 6).

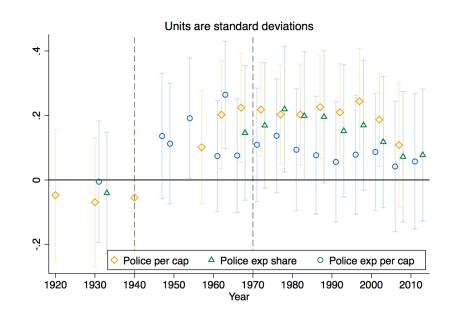


FIGURE F3: GREAT MIGRATION IMPACT ON POLICING INVESTMENTS

Notes: This figure plots the coefficient on a 1 s.d. increase in predicted Black population change in separate regressions for each year where the dependent variable is either the share of local government expenditures on policing, police expenditures per capita, or city police employees per 100k urban population. The unit of observation is a commuting zone. Units of outcome variables are standard deviations. The instrument is the percentile of predicted Black population increase, defined as the interaction between pre-1940 Black southern migration patterns and post-1940 outflows of migrants as predicted by southern economic factors alone. A one standard deviation increase is approximately 30 percentiles. Baseline 1940 controls include share of urban population made up of 1935-1940 Black southern migrants, educational upward mobility, share of labor force in manufacturing, and census region fixed effects. *Data sources*: Financial statistics of state and local governments, 1932; US Census Bureau Annual Survey of Local Governments (1967-2012);Census of Governments, 1952-1989; IPUMS complete count US censuses (1920-1940); Boustan (2016a).

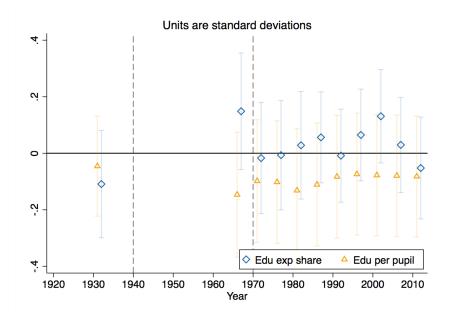


FIGURE F4: GREAT MIGRATION IMPACT ON SCHOOLING INVESTMENTS

Notes: This figure plots the coefficient on a 1 s.d. increase in predicted Black population change in separate regressions for each year where the dependent variable is either the share of local government expenditures on education or education expenditures per student. Education expenditure data are for elementary and high school districts. The unit of observation is a commuting zone. Units of outcome variables are standard deviations. Baseline 1940 controls include share of urban population made up of 1935-1940 Black southern migrants, educational upward mobility, share of labor force in manufacturing, and census region fixed effects. *Data sources*: Financial statistics of state and local governments, 1932; US Census Bureau Annual Survey of Local Governments (1967-2012); IPUMS complete count US censuses (1920-1940); Boustan (2016a).

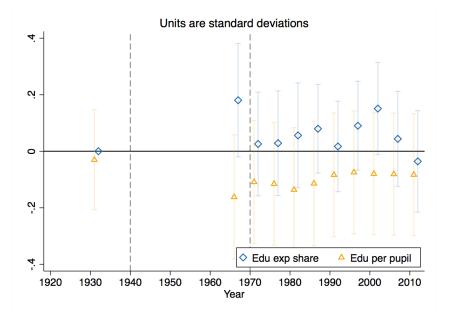


FIGURE F5: GREAT MIGRATION IMPACT ON SCHOOLING INVESTMENTS, WITH PRE-1940 CONTROL

Notes: This figure plots the coefficient on a 1 s.d. increase in predicted Black population change in separate regressions for each year where the dependent variable is either the share of local government expenditures on education or education expenditures per student. Education expenditure data are for elementary and high school districts. All regressions include controls for the 1932 share of local government expenditures on education. The unit of observation is a commuting zone. Units of outcome variables are standard deviations. Baseline 1940 controls include share of urban population made up of 1935-1940 Black southern migrants, educational upward mobility, share of labor force in manufacturing, and census region fixed effects. *Data sources*: Financial statistics of state and local governments, 1932; US Census Bureau Annual Survey of Local Governments (1967-2012); IPUMS complete count US censuses (1920-1940); Boustan (2016a).

To check whether the effect of the Migration on police expenditures is simply driven by increases in municipal spending in Great Migration destinations, I estimate the impact of the shock on fire fighting expenditures. Figure F6 reports these results. I find no impact of the Migration on fire-fighting expenditures. Higher police expenditures may be associated with higher crime and incarceration rates. I investigate these below.

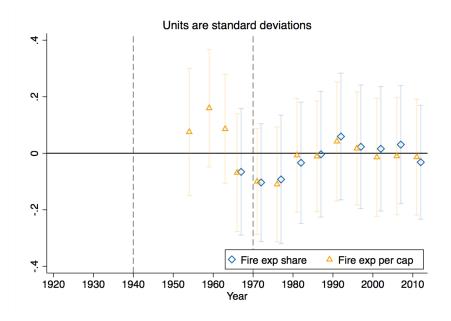


FIGURE F6: GREAT MIGRATION IMPACT ON FIRE-FIGHTING INVESTMENTS

Notes: This figure plots the coefficient on a 1 s.d. increase in predicted Black population change in separate regressions for each year where the dependent variable is either the share of local government expenditures on fire-fighting or fire-fighting expenditures per capita. The unit of observation is a commuting zone. Units of outcome variables are standard deviations. Baseline 1940 controls include share of urban population made up of 1935-1940 Black southern migrants, educational upward mobility, share of labor force in manufacturing, and census region fixed effects. *Data sources*: US Census Bureau Annual Survey of Local Governments (1967-2012); IPUMS complete count US censuses (1920-1940); Boustan (2016a).

F.3 Impact on incarceration rates

Figure F7 plots the coefficients on predicted Black population increases on standardized measures of incarceration separately for each year. The outcome variables are the local correctional institution population per 100,000, the non-white local correctional institution population per 100,000 of the non-white population, and the state and federal imprisoned population by commuting-zone-of-commitment per 100,000, for all individuals aged 15-64 and then separately for this group by race. As can be seen in the Figure, the Migration had no statistically significant effects on pre-1940 incarceration. The Migration is most strongly associated with incarceration in the 1980s and 1990s, during the rise of incarceration rates nationally.

In Figure F8, I report the impact of the Migration on the incarceration rate in levels. At the peak of the association between the Great Migration and Black incarceration rates, in 1992, a 1 standard-deviation increase in predicted Black population increases was associated with 300 more Black people per 100,000 being committed to federal and state prison. The impact for whites was an increase of approximately 30 per 100,000.

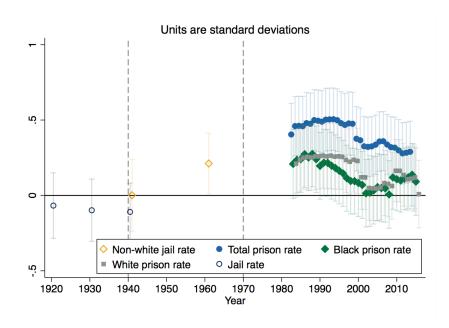


FIGURE F7: GREAT MIGRATION IMPACT ON INCARCERATION RATES

Notes: This figure plots the coefficient on a 1 s.d. increase in predicted Black population change in separate regressions for each year where the dependent variable is county jail population per 100,000 (1940 and 1960) or federal and state prison population by 100,000 by county-of-commitment from 1983-2015. Each jail or prison population group is normalized by the population for that group. Federal and state prison rates are for Black and white men aged 15-64. The unit of observation is a commuting zone. Units of outcome variables are standard deviations. The instrument is the percentile of predicted Black population increase, defined as the interaction between pre-1940 Black southern migration patterns and post-1940 outflows of migrants as predicted by southern economic factors alone. A one standard deviation increase is approximately 30 percentiles. Baseline 1940 controls include share of urban population made up of 1935-1940 Black southern migrants, educational upward mobility, share of labor force in manufacturing, and census region fixed effects. *Data sources*: 1960 US Census; Vera Institute of Justice In Our Backyards Database; IPUMS complete count 1940 US census; Boustan (2016a).

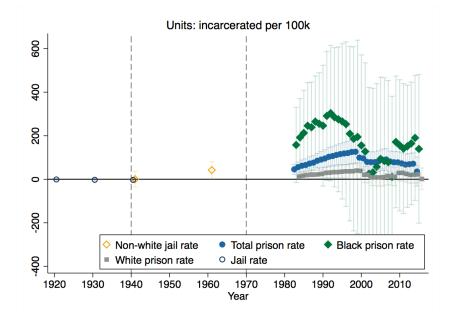


FIGURE F8: GREAT MIGRATION IMPACT ON INCARCERATION RATES, LEVELS

Notes: This figure plots the coefficient on a 1 s.d. increase in predicted Black population change in separate regressions for each year where the dependent variable is county jail population per 100,000 (1940 and 1960) or federal and state prison population by 100,000 by county-of-commitment from 1983-2015. Each jail or prison population group is normalized by the population for that group. Federal and state prison rates are for Black and white men aged 15-64. The unit of observation is a commuting zone. Baseline 1940 controls include share of urban population made up of 1935-1940 Black southern migrants, educational upward mobility, share of labor force in manufacturing, and census region fixed effects. *Data sources*: 1960 US Census; Vera Institute of Justice In Our Backyards Database; IPUMS complete count 1940 US census; Boustan (2016a).

F.4 Impact on murder rates

Figure F9 shows the impact of the migration on standardized measures of murder rates between 1931 and 2015. A 1 standard-deviation increase in the Great Migration shock is associated with just under 0.3 standard deviations higher murder rates in 1931, before the period of Black population change predicted by the shock, but is not associated with higher murder rates in 1936 or 1943. Murder rates are not significantly associated with the Migration again until the late 1960s. In the post-1970 period, a 1 standard-deviation increase in the migration shock is associated with a 0.2 standard deviation increase in murder rates. Controlling for the 1931 murder rate attenuates some of the impact of the Migration on post-1970 murder rates, but the effect on late 1960s murder rates remains positive and statistically significant.

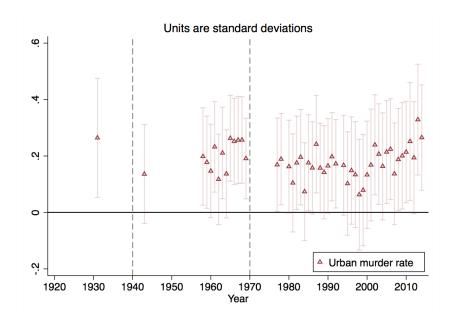


FIGURE F9: GREAT MIGRATION IMPACT ON MURDER RATES

Notes: This figure plots the coefficient on a 1 s.d. increase in predicted Black population change in separate regressions for each year where the dependent variable is urban murder rates per 100,000 in commuting zones. The unit of observation is a commuting zone. Units of outcome variables are standard deviations. The instrument is the percentile of predicted Black population increase, defined as the interaction between pre-1940 Black southern migration patterns and post-1940 outflows of migrants as predicted by southern economic factors alone. A one standard deviation increase is approximately 30 percentiles. Baseline 1940 controls include share of urban population made up of 1935-1940 Black southern migrants, educational upward mobility, share of labor force in manufacturing, and census region fixed effects. *Data sources*: Uniform Crime Reports; IPUMS complete count 1940 US census; Boustan (2016a).

F.5 Impact on racial animus

In a final set of results, I explore the effect of the Migration on racial attitudes, both at the end of the 1960s and in the 2000s. To proxy for racial animus in the 1960s, I use the the share of votes for segregationist presidential candidate George Wallace in the 1968 election from Clubb et al. (2006). White voters would still have made up the majority of the electorate in most cities in the sample at the time, thus votes for Wallace may reflect increased racial animus among white voters.⁵⁴ Table F1 reports the OLS, reduced form, and 2SLS results. Baseline controls are included. Focusing on the reduced form results, I find that a 1 standard-deviation increase in the instrument for the Great Migration increases George Wallace's vote share per 1,000 voters by 12.2 pp. Alternatively, normalizing votes by 1,000 of the white population, the effect is an increase of 4.8 pp. These are sizable relative to the sample mean of each variable, 58.5 votes per 1,000 voters and 24.2 votes per 1,000 white population.

The late 1960s were also marked by a series of race riots that erupted in urban areas across the US. As another measure of racial tension, I explore whether the Migration affected the intensity of these riots. Table F2 reports these results. I find that Great Migration destination cities experienced longer riots and that riots in these areas involved more deaths, injuries, and arrests than places with fewer Black migrant inflows. Focusing on the reduced form, a 1 standard-deviation increase in the Migration shock is associated with over 12 more arrests per 100,000 during the 1960s riots. Both of these events may have contributed to rising police investments during this period. Both the impact on police expenditures and incarceration rates appear to have persisted for several decades afterwards.

Next I provide suggestive evidence that the Migration is associated with greater levels of racial animus today. I examine the reduced form relationship between the instrument for the Great Migration and Google searches for racist terms between 2004 and 2007 from Stephens-Davidowitz (2014b). To construct

 $^{^{54}}$ Voter registration data come from ICPSR (1991).

a CZ-level index, I take the population-weighted average of the Racial Animus Index, which is available at the media market level. The regression is weighted using the CZ population in 2000, and baseline controls are included. The results show a positive relationship between the percentile of predicted Black population between 1940 and 1970 and racial animus in CZs today, suggesting potential long-lasting effects of the Migration on racial attitudes.

	Ordinary Least Squares				
	Wallace Vote	Wallace Votes			
	Per 1k Voters	Per 1k White Pop			
GM	12.41	4.812			
	(4.058)	(1.713)			
R-squared	0.518	0.514			
	Redu	uced Form			
\hat{GM}	12.20	4.840			
	(3.642)	(1.536)			
R-squared	0.525	0.521			
	Two-stage	e least squares			
GM	40.39	16.03			
	(13.87)	(5.771)			
none					
Ν	130	130			
Mean Dep Var	58.49	24.17			
SD Dep Var	44.87	18.85			
SD GM	28.98	28.98			

TABLE F1: GREAT MIGRATION AND VOTES FOR GEORGE WALLACE, 1968

Notes: This table reports the estimated impact of a 1 s.d. increase in Great Migration inflows on votes for George Wallace, pro-segregation former governor of Alabama and third-party presidential candidate in 1968. Dependent variable is votes for Wallace per 1000 voters in column 1 and votes for Wallace per 1,000 white population in column 2. The unit of observation is a commuting zone. \widehat{GM} is the instrument for the Great Migration, or predicted Black population increase through variation in Black southern migration alone. OLS, Reduced Form, and 2SLS estimates are reported. Baseline 1940 controls include share of urban population made up of 1935-1940 Black southern migrants, educational upward mobility, share of labor force in manufacturing, and census region fixed effects. Standard errors are in parentheses. *Data sources*: Clubb et al. (2006); CCDB; IPUMS complete count 1940 US census; Boustan (2016a).

	Ordinary Least Squares					
	Killed Per 100k	Arson Per 100k	Arrests Per 100k	Days of Riots Per 100k	Injured Per 100k	Riots Per 100k
GM	0.0403 (0.0175)	2.189 (0.970)	13.43 (4.824)	0.767 (0.158)	2.801 (1.018)	0.314 (0.0619)
R-squared	0.308	0.440	0.605	0.292	0.461	0.311
			Redi	iced Form		
\hat{GM}	0.0240 (0.0160)	1.424 (0.886)	12.38 (4.354)	0.291 (0.153)	2.264 (0.926)	0.120 (0.0606)
R-squared	0.291	0.429	0.606	0.179	0.455	0.191
			Two-stag	e least squares		
GM	0.0795 (0.0518)	4.715 (2.890)	41.00 (15.74)	0.962 (0.460)	7.496 (3.198)	0.396 (0.181)
none						
Ν	130	130	130	130	130	130
Mean Dep Var SD Dep Var	0.0589 0.162	4.697 9.952	24.91 58.87	0.950 1.437	4.474 10.65	0.413 0.572
SD GM	28.98	28.98	28.98	28.98	28.98	28.98

TABLE F2: GREAT MIGRATION CZS EXPERIENCED MORE SEVERE 1960S RIOTS

Notes: This table reports the estimated impact of a 1 s.d. increase in Great Migration inflows on 1960s race riots and riot severity. Dependent variables in columns 1-5 are individual measures of the severity of riots, including number of individuals killed, number of arson incidents, number of arrests, the duration of the riot in days, and the number of injuries; the final column is total number of riots. All outcomes are normalized by the total CZ population in 1960 and multiplied by 100,000, so they are in per 100,000 of the population units. The unit of observation is a commuting zone. GM is the instrument for the Great Migration, or predicted Black population increase through variation in Black southern migration alone. OLS, Reduced Form, and 2SLS estimates are reported. Baseline 1940 controls include share of urban population made up of 1935-1940 Black southern migrants, educational upward mobility, share of labor force in manufacturing, and census region fixed effects. Standard errors are in parentheses. *Data sources*: Collins and Margo (2007b); Carter (1986); CCDB; IPUMS complete count 1940 US census; Boustan (2016a).

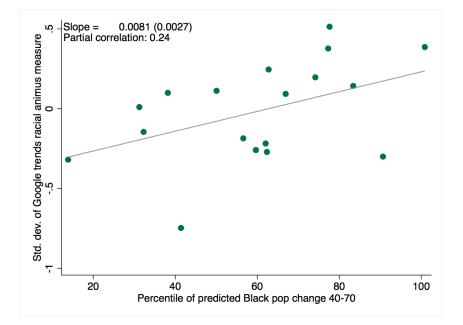


FIGURE F10: Association between Great Migration and Racial Animus in the 2000s

Notes: This binned scatterplot depicts the relationship between the Stephens-Davidowitz (2014a) Racial Animus Index based on Google searches for racist terms from 2004-2007 and the instrument for Black population increases during the Great Migration. The unit of observation is a CZ. The right hand side variable is grouped into 20 bins (5 percentiles each). A population-weighted average of the Racial Animus Index at the CZ level was taken and the measure standardized across the CZs in the sample. The regression is weighted by the CZ population in 2000. The instrument is the percentile of predicted Black population increase, defined as the interaction between pre-1940 Black southern migration patterns and post-1940 outflows of migrants as predicted by southern economic factors alone. Both the left hand and right hand side variables have been residualized on the set of baseline 1940 controls, including share of urban population made up of 1935-1940 Black southern migrants, educational upward mobility, share of labor force in manufacturing, and census region fixed effects. *Data sources*: IPUMS complete count 1940 US census; Boustan (2016a); Stephens-Davidowitz (2014a).

Appendix G Additional data sources

In addition to the data sources described in the preceding appendices, population data used in various measures were obtained from Haines (2010), data from Schpero (2016) were used in geographic crosswalks, and data from Fouka et al. (2018) were used in the analysis.

Appendix H Replication files

Replication files for this study are available at https://doi.org/10.3886/E147963V1.

References

- Adão, R., M. Kolesár, and E. Morales (2019). Shift-Share Designs: Theory and Inference. Quarterly Journal of Economics 134(4), 1949–2010.
- Autor, D. H. and D. Dorn (2013). Replication data for: The Growth of Low-Skill Service Jobs and the Polarization of the U.S. Labor Market. https://www.ddorn.net/data.htm. (Accessed on August 30, 2016).
- Belloni, A., V. Chernozhukov, and C. Hansen (2011). Lasso Methods for Gaussian Instrumental Variables Models. arxiv:[math.st].
- Borusyak, K. and P. Hull (2020). Non-Random Exposure to Exogenous Shocks: Theory and Applications. NBER Working Paper 27845.
- Borusyak, K., P. Hull, and X. Jaravel (2021). Quasi-Experimental Shift-Share Research Designs. *The Review of Economic Studies Forthcoming.*
- Boustan, L. and M. Tabellini (2018). Black Out-Migration and Southern Political Realignment. Working Paper, Harvard Business School.
- Boustan, L. P. (2009). Competition in the Promised land: Black Migration and Racial Wage Convergence in the North, 1940–1970. Journal of Economic History 69(3), 755–782.
- Boustan, L. P. (2010). Was Postwar Suburbanization "White Flight"? Evidence from the Black Migration. *Quarterly Journal of Economics* 125(1), 417–443.
- Boustan, L. P. (2016a). Competition in the Promised Land: Black Migrants in Northern Cities and Labor Markets. Princeton: Princeton University Press.
- Boustan, L. P. (2016b). Data for: Competition in the Promised Land: Black Migrants in Northern Cities and Labor Markets. Princeton University Press. https://scholar.princeton.edu/lboustan/data-books. (Accessed on March 4, 2017).
- Bowles, G. K., J. D. Tarver, C. L. Beale, and E. S. Lee (2016). Net Migration of the Population by Age, Sex, and Race, 1950-1970. Dataset, Inter-university Consortium for Political and Social Research [distributor]. https://doi.org/10.3886/ICPSR08493.v2.

- Calderon, A., V. Fouka, and M. Tabellini (2019). Racial Diversity, Electoral Preferences, and the Supply of Policy: the Great Migration and Civil Rights. IZA Discussion Paper 14312.
- Card, D., C. Domnisoru, and L. Taylor (2018). The Intergenerational Transmission of Human Capital: Evidence from the Golden Age of Upward Mobility. NBER Working Paper 25000.
- Carter, G. L. (1986). The 1960s Black Riots Revisited: City Level Explanations of Their Severity. Sociological Inquiry 56(2), 210–228.
- CDC (2021). CDC NCHS National Center for Health Statistics. https://www.cdc.gov/nchs/index.htm.
- Chernozhukov, V., D. Chetverikov, M. Demirer, E. Duflo, C. Hansen, W. Newey, and J. Robins (2018). Double/Debiased Machine Learning for Treatment and Structural Parameters. *Econometrics Journal* 21(1), C1– C68.
- Chetty, R., J. N. Friedman, N. Hendren, and M. R. Jones (2018). Replication data for: The Opportunity Atlas: Mapping the Childhood Roots of Social Mobility. https://opportunityinsights.org/data. (Accessed on October 1, 2018).
- Chetty, R., J. N. Friedman, N. Hendren, M. R. Jones, and S. R. Porter (2018). The Opportunity Atlas: Mapping the Childhood Roots of Social Mobility. Nber working paper 25147.
- Chetty, R. and N. Hendren (2018a). Replication data for: The Impacts of Neighborhoods on Intergenerational Mobility II: County-Level Estimates. https://opportunityinsights.org/data. (Accessed on August 24, 2017).
- Chetty, R. and N. Hendren (2018b). The Impacts of Neighborhoods on Intergenerational Mobility I: Childhood Exposure Effects. *Quarterly Journal of Economics* 133(3), 1107–1162.
- Chetty, R. and N. Hendren (2018c). The Impacts of Neighborhoods on Intergenerational Mobility II: County-Level Estimates. *Quarterly Journal of Economics* 133(3), 1163–1228.
- Chetty, R., N. Hendren, M. R. Jones, and S. R. Porter. Replication data for: Race and Economic Opportunity in the United States: An intergenerational perspective. https://opportunityinsights.org/data. (Accessed on October 1, 2018).

- Chetty, R., N. Hendren, M. R. Jones, and S. R. Porter (2020). Race and economic opportunity in the united states: an intergenerational perspective. *Quarterly Journal of Economics* 135(2), 711–783. https://doi.org/10.1093/qje/qjz042.
- Chetty, R., N. Hendren, and L. F. Katz (2016). The Effects of Exposure to Better Neighborhoods on Children: New Evidence from the Moving to Opportunity Experiment. American Economic Review 106(4), 855–902.
- Chetty, R., N. Hendren, P. Kline, E. Saez, and N. Turner (2014). Is the United States Still a Land of Opportunity? Recent Trends in Intergenerational Mobility. *American Economic Review* 104(5), 141–47.
- Clubb, J. M., W. H. Flanigan, and N. H. Zingale (2006). Electoral Data for Counties in the United States: Presidential and Congressional Races, 1840-1972. Dataset, Inter-university Consortium for Political and Social Research [distributor]. https://doi.org/10.3886/ICPSR08611.v1.
- Collins, W. J. and R. A. Margo (2007a). Data for: The Economic Aftermath of the 1960s Riots in American Cities: Evidence from Property Values. Unpublished data. (Received on July 12, 2018 via email from William Collins.).
- Collins, W. J. and R. A. Margo (2007b). The Economic Aftermath of the 1960s Riots: Evidence from Property Values. *Journal of Economic History* 67(4), 849–883.
- Collins, W. J. and M. H. Wanamaker (2014). Selection and Economic Gains in the Great Migration of African Americans: New Evidence from Linked Census Data. American Economic Journal: Applied Economics 6(1), 220– 252.
- Deutscher, N. (2020). Place, Peers, and the Teenage Years: Long-Run Neighborhood Effects in Australia. American Economic Journal: Applied Economics 12(2), 220–49.
- Fishback, P. V., R. S. Johnson, and S. Kantor (2010). Data for: Striking at the Roots of Crime: The Impact of Welfare Spending on Crime During the Great Depression. Unpublished data. (Received on June 7, 2018 via email from James Feigenbaum, with permission from Price Fishback.).
- Fouka, V., S. Mazumder, and M. Tabellini (2018). Data for: From Immigrants to Americans: Race, Status, and Assimilation During the Great Migration. Unpublished data. (Received on June 15, 2018 via email from Vasiliki Fouka.).

- Fouka, V., S. Mazumder, and M. Tabellini (2021). From Immigrants to Americans: Race and Assimilation during the Great Migration. The Review of Economic Studies Forthcoming.
- Gardner, J. and W. Cohen (1992). Demographic Characteristics of the Population of the United States, 1930-1950: County-Level. Dataset, Inter-university Consortium for Political and Social Research [distributor]. http://dx.doi.org/10.3886/ICPSR08493.v2.
- Haines, M. R. (2010). Historical, Demographic, Economic, and Social Data: The United States, 1790-2002. Dataset, Interuniversity Consortium for Political and Social Research [distributor]. http://dx.doi.org/10.3886/ICPSR02896.v3.
- ICPSR (1991). Voter Registration in the United States, 1968-1988. Dataset, Inter-university Consortium for Political and Social Research [distributor]. http://dx.doi.org/10.3886/ICPSR09405.V1.
- ICPSR (2005). Uniform Crime Reports, 1958-1969, and County and City Data Books, 1962, 1967, 1972: Merged Data. Dataset, Inter-University Consortium for Political and Social Research [distributor]. http://dx.doi.org/10.3886/ICPSR07715.v1.
- Kang-Brown, J., O. Hinds, E. Schattner-Elmaleh, and J. Wallace-Lee (2020). Incarceration Trends Project: Data and Methods for Historical Jail Populations in U.S. Counties, 1970-2018. Dataset, Vera Institute of Justice. http://trends.vera.org/incarceration-rates?data=pretrial.
- Manson, S., J. Schroeder, D. Van Riper, T. Kugler, and S. Ruggles (2021). IPUMS National Historical Geographic Information System: Version 16.0. Dataset, Minneapolis, MN: IPUMS. http://doi.org/10.18128/D050.V16.0.
- Margo, R. A. (1990). Race and Schooling in the South, 1880-1950: An Economic History. Chicago: University of Chicago Press.
- Margo, R. A. (1991). Segregated Schools and the Mobility Hypothesis: A Model of Local Government Discrimination. Quarterly Journal of Economics 106(1), 61–73.
- Ruggles, S., S. Flood, R. Goeken, J. Grover, E. Meyer, J. Pacas, and M. Sobek (2020). IPUMS Restricted Complete Count Data: Version 11.0. Dataset, Minneapolis, MN: IPUMS. https://doi.org/10.18128/D010.V11.0.

- Saavedra, M. and T. Twinam (2020a). A Machine Learning Approach to Improving Occupational Income Scores. *Explorations in Economic History* 75.
- Saavedra, M. and T. Twinam (2020b). Data for: A Machine Learning Approach to Improving Occupational Income Scores. *Explorations in Economic History 75*. https://www2.oberlin.edu/faculty/msaavedr/lido.html. (Accessed on December 16, 2020).
- Schpero, W. L. (2016, July). STATASTATES: Stata module to add US state identifiers to dataset. Statistical Software Components, Boston College Department of Economics.
- Sequeira, S., N. Nunn, and N. Qian (2019). Data for: Immigrants and the Making of America. https://scholar.harvard.edu/nunn/pages/data-0. (Accessed on December 22, 2018).
- Sequeira, S., N. Nunn, and N. Qian (2020). Immigrants and the Making of America. *Review of Economic Studies* 87(1), 382–419.
- Stephens-Davidowitz, S. (2014a). Racial Animus (Google Search Data): Racial Animus, by DMA [1 csv file]. http://sethsd.com/research. (Accessed on December 16, 2020).
- Stephens-Davidowitz, S. (2014b). The Cost of Racial Animus on a Black Candidate: Evidence using Google Search Data. Journal of Public Economics 118, 26–40.
- Tabellini, M. (2019). Racial Heterogeneity and Local Government Finances: Evidence from the Great Migration. Harvard Business School Working Paper 19-006.
- U.S. Census Bureau. City Government Employment data, 1957-2007. https://web.archive.org/web/20170419040247/https://www2.census.gov/ pub/outgoing/govs/special60/City_Govt_Emp.zip. (Accessed on April 13, 2017).
- U.S. Census Bureau. Data Base on Historical Finances of County Governments: County Area Finances, Fiscal Years 1957-2006. https://web.archive.org/web/20170419040247/https://www2.census.gov/ pub/outgoing/govs/special60/County_Area_Fin.ziphttps://web.archive.org /web/20170419040247/https://www2.census.gov/pub/outgoing/govs/special60/ County_Area_Fin.zip. (Accessed on April 13, 2017).

- U.S. Census Bureau. Individual Government Finances data, 1967-2012. https://www2.census.gov/pub/outgoing/govs/special60/_IndFin_1967-2012.zip. (Accessed on April 13, 2017).
- 1940 Census of U.S. Census Bureau (1943). Population: Special Report on Institutional Population 14Years and Over. Penal Characteristics of Inmates in Institutions, and in Inthe Delinquent, Defective. Dependent. stitutions for and https://www.census.gov/library/publications/1943/dec/populationinstitutional.html.
- U.S. Census Bureau (1963).1960 Census: Population, Subject Reports, Inmates of Institutions: Social and Eco-Data for Inmates by Area and Type of Institution. nomic https://www.census.gov/library/publications/1963/dec/population-pc-2-8a.html.
- U.S. Census Bureau (1998). New England County Metropolitan Areas and Components, 1990, with FIPS Codes. https://www2.census.gov/programssurveys/metro-micro/geographies/reference-files/1990/historicaldelineation-files/90nfips.txt. (Accessed on March 26, 2018).
- U.S. Census Bureau (2001). Metropolitan Areas and Components, 1990, with FIPS Codes. https://www2.census.gov/programssurveys/metro-micro/geographies/reference-files/1990/historicaldelineation-files/90mfips.txt. (Accessed on March 26, 2018).
- U.S. Census Bureau, Department of Commerce (1933). Financial Statistics of State and Local Governments: 1931[-1932]. Technical report, Hathi Trust [distributor]. http://hdl.handle.net/2027/uiug.30112106770735.
- U.S. Census Bureau, Department of Commerce (2008). County and City Data Book United States Consolidated File: City Data, 1944-1977. Dataset, Inter-university Consortium for Political and Social Research [distributor]. http://dx.doi.org/10.3886/ICPSR07735.v1.
- U.S. Census Bureau, Department of Commerce (2012). County and City Data Book Consolidated File: County Data, 1947-1977. Dataset, Interuniversity Consortium for Political and Social Research [distributor]. https://doi.org/10.3886/ICPSR07736.v2.
- U.S. Department of Justice, FBI (1931-1950). Uniform Crime Reports for the United States. Report, Hathi Trust

[distributor]. http://hdl.handle.net/2027/mdp.39015039714384, http://hdl.handle.net/2027/mdp.3901503971440, http://hdl.handle.net/2027/mdp.39015010691601, http://hdl.handle.net/2027/mdp.39015010691627.

U.S. Office of Education, Department of the Interior (1924). Biennial Survey of Education in the United States. Report, Hathi Trust [distributor]. http://hdl.handle.net/2027/umn.31951d01976418l.