# Online Appendix 

# "Cell Phone Access and Election Fraud: Evidence from a Spatial Regression Discontinuity Design in Afghanistan" 

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## Appendix A The Audit and Recount Process

In its final report of the 2009 election (Electoral Complaints Commission, 2010), the ECC reported that after receiving an increasing number of complaints on ballot stuffing and other irregularities, they decided on September 8, 2009 to conduct an audit of polling stations nationwide. To this end, they ordered the IEC to conduct an audit and recount of stations satisfying the following criteria:

- A1: Stations in which 600 or more valid votes were cast
- B1: Stations with more than 100 votes in which one candidate received 95 percent or more of the total votes cast
- C1: Stations satisfying both A1 and B1

After the initial samples were drawn, however, three additional categories were created due to a misunderstanding of the ECC orders by the IEC ${ }^{1}$. The three new categories expanded the scope of the audit. The categories were:

- A2: Stations with 600 or more votes cast (Excluding those in A1)
- B2: Stations in which a candidate received 95 percent or more of the total valid votes cast (Excluding those in B1)
- C2: Stations satisfying both A2 and B2

After the audit process, the IEC reported that 3,376 stations classified in at least one of these categories. Out of this sample, the ECC and IEC investigated 10 percent of the ballots within each category. Some of the physical indicators used to determine fraud were whether the ballot box was tampered, all required materials were included, visual inspection of the ballots, reviews of the tally results and the actual ballot counts, among others. For the purpose of this study I aggregate the six categories described above into three broader categories:

- Category A: Stations with 600 or more votes cast. Defined as $\mathrm{A} 1+\mathrm{A} 2+\mathrm{C} 1+\mathrm{C} 2$ from the categories above.
- Category B: Stations in which one candidate received 95 percent or more of the total votes cast. Defined as $\mathrm{B} 1+\mathrm{B} 2+\mathrm{C} 1+\mathrm{C} 2$

[^1]- Category C: Stations satisfying Categories A and B above. Defined as C1+C2.

The number of polling stations within each category is 1,706 in category A ( 545 from A1 +299 from $\mathrm{A} 2+741$ from $\mathrm{C} 1+121$ from C2), 2,532 in category $\mathrm{B}(1269$ from B1 +401 from B2 +741 from $\mathrm{C} 1+121$ from C 2$)$, and 862 in category $\mathrm{C}(741$ from $\mathrm{C} 1+121$ from C 2$)$.

## Appendix B Additional Figures and Tables



Expected voters


Center type (Other)


Total votes


Polling stations


Voter turnout


Center type (Mosque)


Road access


Share (Karzai)


Center type (School)


Limited access


Other access

Figure B1: Binned Averages for various Covariates (Covariate RD Plots) - Continues


Female stations


Slope (percent)


Distance (district hospital)


Distance (secondary river)


Kuchis stations


Distance (primary road)


Distance (basic clinic)


Distance (seasonal river)


Elevation (m)


Distance (secondary road)


Distance (primary river)


District/Provincial capital

Figure B1: Binned Averages for various Covariates (Covariate RD Plots) - Continues


Figure B1: Binned Averages for various Covariates (Covariate RD Plots) - Continued
Notes: Solid dots give the average value of the specified covariate for polling centers falling within 4000-meter distance bins. Hollow dots give the average value of the specified covariate for polling centers falling within 2000-meter distance bins. "Distance to boundary" refers to the distance between a polling center and the closest point in the cell phone coverage boundary. "Negative" values of distance give the distance of polling centers/villages in non-coverage areas. The solid line trends give the predicted values from a regression of the outcome variable on a second degree polynomial in distance to the boundary that uses a triangular kernel and a bandwidth of 20,000 meters.


Figure B2: Histograms and densities of the forcing variable
Notes: "Distance to boundary" refers to the distance between a polling center (Panels a and b) and villages (Panels c and d) to the closest point in the coverage boundary. Distance is measured in meters. Bin width of 160 meters. The distance to boundary is normalized so that "negative" values of distance give the distance of polling centers/villages in non-coverage areas. Panels (b) and (d) provide results from Panel (b) uses the test for breaks in the density of the forcing variable proposed in Cattaneo et al. (2019) and uses the code discussed in Cattaneo et al. (2018). P-value for test presented in figure caption.


Figure B3: Binned Averages for Category C fraud (RD plot)
Notes: South/East region (Panels A and B). North/West region (Panels C and D). Solid dots give the average share of votes classifying in Category C fraud for polling centers falling within 4000-meter distance bins. Hollow dots give the average share of votes classifying in Category C fraud for polling centers falling within 2000-meter distance bins. Refer to section II in the text for a detailed description of Category C fraud. "Distance to boundary" refers to the distance between a polling center and the closest point in the cell phone coverage boundary. "Negative" values of distance give the distance of polling centers/villages in non-coverage areas. The solid line trends give the predicted values from a regression of the outcome variable on a second degree polynomial in distance to the boundary that uses a triangular kernel and a bandwidth of 20,000 meters.


Figure B4: Binned Averages for Category C fraud (RD plot)
Notes: South/East region (Panels A and B). North/West region (Panels C and D). Solid dots give the average share of votes classifying in Category C fraud for polling centers falling within 4000-meter distance bins. Hollow dots give the average share of votes classifying in Category C fraud for polling centers falling within 2000-meter distance bins. Refer to section II in the text for a detailed description of Category C fraud. "Distance to boundary" refers to the distance between a polling center and the closest point in the cell phone coverage boundary. "Negative" values of distance give the distance of polling centers/villages in non-coverage areas. The solid line trends give the predicted values from a regression of the outcome variable on a second degree polynomial in distance to the boundary that uses a triangular kernel and a bandwidth of 20,000 meters.


Figure B5: Afghan provinces and regions
Notes: Regions of Afghanistan. Darker shade indicates the Southeastern provinces. Regions defined using International Security Assistance Forces (ISAF) regional command center definitions. Lines demarcate the provinces of Afghanistan. Blue dots indicate the location of Afghan villages (MISTI, 2013). Map overlaid on USGS topographic map.


Figure B6: Sensitivity of Results to Bandwidth Choice and Polynomial order (Category C fraud)
Notes: Panels (a) and (b): Each dot indicates the RD estimate using the specified bandwidth. Range spikes indicate $95 \%$ confidence intervals of the estimates. Panels (c) and (d): Each dot indicates the RD estimate using the specified order in the RD polynomial. Range spikes indicate $95 \%$ confidence intervals of the estimates.


Figure B7: Mobile towers in 2016 but not in 2009
Notes: Blue dots indicate the location of towers in Q1 of 2016. Data obtained from the Afghan Telecommunications Regulatory Authority (ATRA). Gray shaded areas indicate coverage in 2009. Light blue shaded areas indicate coverage areas post 2009 and used in the falsification exercise in section 3.4.1.


Figure B8: Insurgent Violence up to Election day (Jan. 1, 2009 - Aug. 20, 2009)
Notes: Dots indicate the location of the violent incident. Color scale of shaded districts gives the rate of violence as the number of incidents per 1,000 inhabitants in the district where the incident took place. Data are significant actions collected by Afghan forces and ISAF. Refer to Condra et al. (2018) for detailed description of the data.


Figure B9: Example of Boundary Segments with Corresponding polling centers
Notes: Example of two boundary segments highlighted in red and blue with polling center (highlighted in corresponding color) belonging to the neighborhood of each segment. Neighborhoods determined by all polling centers closest to a specific segment of the boundary.

Table B1: Sample and Imputations

|  | Full sample |  | Within 10 km of boundary |  | Within 6 km of boundary |  | Within 4 km of boundary |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Freq. | Percent | Freq. | Percent | Freq. | Percent | Freq. | Percent |
| Panel A. Unrestricted sample |  |  |  |  |  |  |  |  |
| Not imputed | 5,904 | 95.84 | 3,329 | 95.77 | 2,388 | 94.99 | 1,845 | 94.81 |
| Imputed based on: |  |  |  |  |  |  |  |  |
| Settlement | 169 | 2.74 | 103 | 2.96 | 89 | 3.54 | 72 | 3.70 |
| Nearest center | 81 | 1.31 | 42 | 1.21 | 35 | 1.39 | 27 | 1.39 |
| District capital | 6 | 0.10 | 2 | 0.06 | 2 | 0.08 | 2 | 0.10 |
| Total | 6,160 | 100 | 3,476 | 100 | 2,514 | 100 | 1,946 | 100 |
| Panel B. Restricted sample |  |  |  |  |  |  |  |  |
| Not imputed | 2,331 | 96.04 | 1,377 | 95.49 | 1,106 | 95.18 | 912 | 95.10 |
| Imputed based on: |  |  |  |  |  |  |  |  |
| Settlement | 67 | 2.76 | 49 | 3.40 | 41 | 3.53 | 36 | 3.75 |
| Nearest center | 26 | 1.07 | 15 | 1.04 | 14 | 1.20 | 10 | 1.04 |
| District capital | 3 | 0.12 | 1 | 0.07 | 1 | 0.09 | 1 | 0.10 |
| Total | 2,427 | 100 | 1,442 | 100 | 1,162 | 100 | 959 | 100 |

Notes: "Not imputed" refers to centers for which data were available after the merging of 2009 fraud data and 2010 geographic coordinate data. Imputations based on settlement give the polling center the coordinates of the village or settlement center where the polling center is located. Imputations based on nearest center give the polling center the coordinates of the polling center that, within the district, has the closest ID code to it. This is done because the assignment of ID codes followed a spatial order for the most part. Imputations based on district center simply give the polling center the coordinates of the district's capital where the center is located. Restricted sample refers to sample where at least one polling center is located on each side of a defined neighborhood. Refer to section 3.3 for the definition of neighborhood. The restricted sample constitutes the main estimation sample.

Table B2: Mean comparison for various polling center characteristics: Southeast Region

|  | Within 10 km of boundary |  |  | Within 5 km of boundary |  |  | RD estimates |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Coverage <br> (1) | No Coverage <br> (2) | S.E. <br> (3) | Coverage <br> (4) | No Coverage <br> (5) | S.E. <br> (6) | RD coeff. <br> (7) | S.E. <br> (8) |
| Electoral outcomes |  |  |  |  |  |  |  |  |
| No. of stations | 4.26 | 4.043 | (0.26) | 4.12 | 4.048 | (0.28) | 0.13 | (0.276) |
| No. of expected voters | 2325 | 2152 | (145.6) | 2255 | 2146 | (157.3) | 131.93 | (158.9) |
| Total votes | 850.6 | 878.4 | (98.36) | 868.8 | 860.1 | (109.2) | 24.67 | (104.5) |
| Voter turnout | 0.40 | 0.461 | (0.04)* | 0.43 | 0.455 | (0.04) | -0.01 | (0.039) |
| Vote share: |  |  |  |  |  |  |  |  |
| Karzai | 0.60 | 0.608 | (0.05) | 0.60 | 0.598 | (0.05) | 0.02 | (0.047) |
| Abdullah | 0.23 | 0.15 | (0.05)* | 0.23 | 0.162 | (0.05) | 0.05 | (0.044) |
| Polling center characteristics |  |  |  |  |  |  |  |  |
| Polling center type: |  |  |  |  |  |  |  |  |
| Mosque | 0.24 | 0.278 | (0.04) | 0.24 | 0.285 | (0.04) | -0.04 | (0.039) |
| School | 0.40 | 0.281 | $(0.04)^{* * *}$ | 0.38 | 0.27 | $(0.04)^{* * *}$ | 0.10 | $(0.042)^{* *}$ |
| Other type | 0.36 | 0.44 | (0.05)* | 0.38 | 0.444 | (0.05) | -0.06 | (0.047) |
| Polling center access (2010): |  |  |  |  |  |  |  |  |
| Road access | 0.62 | 0.652 | (0.06) | 0.59 | 0.644 | (0.06) | -0.07 | (0.059) |
| Limited access | 0.11 | 0.096 | (0.03) | 0.13 | 0.104 | (0.03) | 0.04 | (0.036) |
| Other access | 0.27 | 0.252 | (0.05) | 0.29 | 0.252 | (0.05) | 0.03 | (0.052) |
| Share female stations | 0.43 | 0.447 | (0.02) | 0.43 | 0.45 | (0.02) | -0.02 | (0.020) |
| Share Kuchis stations | 0.05 | 0.0363 | (0.01) | 0.03 | 0.0406 | (0.01) | 0.00 | (0.013) |
| Geographic characteristics |  |  |  |  |  |  |  |  |
| Elevation (meters) | 1789.00 | 1961 | $(73.54)^{* *}$ | 1812.00 | 1937 | $(70.94)^{*}$ | -142.20 | $(66.975)^{* *}$ |
| Slope (percent) | 5.67 | 7.277 | $(0.76)^{* *}$ | 6.30 | 7.226 | (0.77) | -0.90 | (0.796) |

Table B2: Mean comparison for various polling center characteristics: Southeast Region - Continued

|  | Within 10 km of boundary |  |  | Within 5 km of boundary |  |  | RD estimates |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Coverage <br> (1) | No Coverage <br> (2) | S.E. <br> (3) | Coverage <br> (4) | No Coverage (5) | S.E. <br> (6) | RD coeff. <br> (7) | S.E. <br> (8) |
| Economic development characteristics |  |  |  |  |  |  |  |  |
| Distance (km) to: |  |  |  |  |  |  |  |  |
| Primary road (2005) | 32.30 | 45.99 | (3.33)*** | 40.11 | 44.75 | (3.21) | -4.12 | (3.022) |
| Secondary road (2005) | 36.03 | 42.83 | (4.15) | 43.73 | 41.37 | (4.22) | 3.00 | (3.880) |
| District hospital (2005) | 25.22 | 31.5 | $(2.60)^{* *}$ | 28.17 | 30.99 | (2.32) | -2.14 | (2.279) |
| Basic health center (2005) | 20.01 | 22.75 | (3.02) | 21.90 | 21.83 | (2.81) | 1.20 | (2.942) |
| Primary river | 16.89 | 18.15 | (1.86) | 18.30 | 18.1 | (1.80) | 1.11 | (1.905) |
| Seasonal river | 6.27 | 5.988 | (1.24) | 7.24 | 5.315 | (1.41) | 1.48 | (1.136) |
| Distance to seasonal river | 15.71 | 12.63 | (2.50) | 15.98 | 12.25 | (2.95) | 3.09 | (2.696) |
| District/Province capital | 0.05 | 0.0397 | (0.02) | 0.04 | 0.0444 | (0.02) | -0.01 | (0.017) |
| Demographic characteristics (of closest settlement) |  |  |  |  |  |  |  |  |
| Population (2012-2013) | 1207 | 1185 | (240.0) | 1009 | 1249 | (214.3) | -246.5 | (207.9) |
| Language spoken (2012-2013): |  |  |  |  |  |  |  |  |
| Dari | 0.25 | 0.205 | (0.05) | 0.23 | 0.211 | (0.05) | 0.02 | (0.044) |
| Pashto | 0.67 | 0.685 | (0.06) | 0.66 | 0.693 | (0.05) | -0.02 | (0.052) |
| Other | 0.08 | 0.109 | (0.03) | 0.11 | 0.0963 | (0.04) | 0.01 | (0.034) |
| Observations |  |  |  |  |  |  |  |  |

Notes: Columns (1), (2), (4), and (5) give the means of the corresponding variable. Columns (3) and (6) give the clustered standard errors for the difference in means in parenthesis. Sample restricted to neighborhoods with at least one observation on each side of the boundary. ${ }^{*}$, **, and ${ }^{* * *}$ indicate 10,5 , and 1 percent significance respectively. No. of expected voters is the number of voters predicted by the IEC prior to election day. Total votes cast is the actual number of votes tallied at the center. Voter turnout is defined as the number of votes cast at the center divided by the expected number of voters. Vote share is the share of votes received by the each of the two main candidates divided by total votes. I report Total votes, Voter turnout, and Vote share for centers without evidence of fraud. The remaining variables are defined in section II.A of the text. Year values in parenthesis indicate the year the data was collected. Variables without year indication were collected in 2009 or are time-invariant variables. Refer to section II.B for a description of the RD model used in columns (7) and (8).

Table B3: Mean comparison for various polling center characteristics (Conley S.E.)

|  | Within 10 km of boundary |  |  | Within 5 km of boundary |  |  | RD estimates |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Coverage <br> (1) | No Coverage <br> (2) | S.E. <br> (3) | Coverage <br> (4) | No Coverage <br> (5) | S.E. <br> (6) | RD coeff. <br> (7) | S.E. <br> (8) |
| Fraud outcomes (Category C fraud) |  |  |  |  |  |  |  |  |
| All regions | 0.08 | 0.11 | $(0.02)^{* *}$ | 0.08 | 0.12 | $(0.02)^{* *}$ | -0.04 | (0.019)** |
|  |  |  | [0.02]** |  |  | [0.02]** |  | [0.020]** |
| East and South | 0.14 | 0.20 | (0.03)** | 0.13 | 0.20 | $(0.03)^{* *}$ | -0.08 | (0.032)** |
|  |  |  | [0.03]** |  |  | [0.03]** |  | [0.033]** |
| North and West | 0.01 | 0.01 | (0.01) | 0.01 | 0.01 | (0.01) | 0.00 | (0.010) |
|  |  |  | [0.01] |  |  | [0.01] |  | [0.010] |
| Electoral outcomes |  |  |  |  |  |  |  |  |
| No. of stations | 4.09 | 3.73 | $(0.16)^{* *}$ | 3.86 | 3.78 | (0.18) | 0.09 | (0.175) |
|  |  |  | [0.16]** |  |  | [0.17] |  | [0.173] |
| No. of expected voters | 2194.00 | 1944.00 | (94.04)*** | 2069.00 | 1979.00 | (100.00) | 94.33 | (101.198) |
|  |  |  | [92.99]*** |  |  | [99.35] |  | [100.161] |
| Total votes | 871.80 | 866.60 | (56.34) | 835.00 | 863.70 | (62.41) | -28.22 | (59.992) |
|  |  |  | [55.72] |  |  | [61.72] |  | [59.076] |
| Voter turnout | 0.43 | 0.50 |  | 0.45 | 0.49 | $(0.02)$ | -0.03 | (0.023) |
|  |  |  | $[0.02]^{* * *}$ |  |  | [0.02] |  | [0.023] |
| Vote share: |  |  |  |  |  |  |  |  |
| Karzai | 0.50 | 0.49 | (0.03) | 0.50 | 0.49 | (0.03) | 0.01 | (0.030) |
|  |  |  | [0.03] |  |  | [0.03] |  | [0.030] |
| Abdullah | 0.34 | 0.33 | (0.03) | 0.35 | 0.33 | (0.03) | 0.01 | (0.031) |
|  |  |  | [0.03] |  |  | [0.03] |  | [0.030] |

Table B3: Mean comparison for various polling center characteristics (Conley S.E.) - Continued


Table B3: Mean comparison for various polling center characteristics (Conley S.E.) - Continued

|  | Within 10 km of boundary |  |  | Within 5 km of boundary |  |  | RD estimates |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Coverage <br> (1) | No Coverage <br> (2) | S.E. <br> (3) | Coverage <br> (4) | No Coverage <br> (5) | S.E. <br> (6) | RD coeff. <br> (7) | S.E. <br> (8) |
| Geographic characteristics |  |  |  |  |  |  |  |  |
| Elevation (meters) | 1570.00 | 1782.00 | $(58.06)^{* * *}$ | 1617.00 | 1756.00 | $(50.71)^{* * *}$ | -128.72 | $(50.617)^{* *}$ |
|  |  |  | [57.59]*** |  |  | [51.89]*** |  | [51.586]** |
| Slope (percent) | 5.72 | 7.57 | (0.53)*** | 6.46 | 7.66 | $(0.58) * *$ | -1.01 | (0.602)* |
|  |  |  | $[0.53]^{* * *}$ |  |  | [0.58]** |  | [0.593]* |
| Economic development characteristics |  |  |  |  |  |  |  |  |
| Distance (km) to: |  |  |  |  |  |  |  |  |
| Primary road (2005) | 35.30 | 48.62 | (2.43)*** | 40.72 | 47.07 | (2.31)*** | -5.70 | (2.290)** |
|  |  |  | [2.51]*** |  |  | [2.41]*** |  | [2.370]** |
| Secondary road (2005) | 44.65 | 52.60 | (3.37)** | 50.08 | 49.84 | (3.12) | 0.49 | (3.056) |
|  |  |  | [3.41]** |  |  | [3.19] |  | [3.110] |
| District hospital (2005) | 37.56 | 45.42 | $(2.76)^{* * *}$ | 40.25 | 42.68 | (2.73) | -2.34 | (2.739) |
|  |  |  | [2.86]*** |  |  | [2.78] |  | [2.789] |
| Basic health center (2005) | 20.12 | 24.31 | (1.87)** | 21.78 | 22.68 | (1.74) | -0.38 | (1.837) |
|  |  |  | [1.83]** |  |  | [1.75] |  | [1.836] |
| Primary river | 17.45 | 18.38 | (1.40) | 18.00 | 17.90 | (1.29) | 0.19 | (1.319) |
|  |  |  | [1.45] |  |  | [1.34] |  | [1.366] |
| Secondary river | 8.11 | 8.71 | (1.12) | 8.96 | 8.03 | (1.10) | 0.61 | (1.040) |
|  |  |  | [1.13] |  |  | [1.13] |  | [1.053] |
| Seasonal river | 12.66 | 11.05 | (1.54) | 13.16 | 10.30 | (1.77) | 2.70 | (1.669) |
|  |  |  | [1.62] |  |  | [1.83] |  | [1.715] |
| District/Province capital | 0.05 | 0.03 | $(0.01)^{* *}$ | 0.04 | 0.03 | (0.01) | 0.01 | (0.013) |
|  |  |  | [0.01]** |  |  | [0.01] |  | [0.013] |

Table B3: Mean comparison for various polling center characteristics (Conley S.E.) - Continued


Notes: Columns (1), (2), (4), and (5) give the means of the corresponding variable. Columns (3) and (6) give the clustered standard errors for the difference in means in parenthesis and Conley (1999) standard errors in brackets for the difference in means. Estimation of standard errors uses code in Hsiang (2010). Conley (1999) standard errors use a distance cutoff of 50 kilometers and a Bartlett spatial weighting kernel. Sample restricted to neighborhoods with at least one observation on each side of the boundary. ${ }^{*},{ }^{* *}$, and ${ }^{* * *}$ indicate 10,5 , and 1 percent significance respectively. "Coverage" refers to cell phone coverage. No. of expected voters is the number of voters predicted by the IEC prior to election day. Total votes cast is the actual number of votes tallied at the center. Voter turnout is defined as the number of votes cast at the center divided by the expected number of voters. Vote share is the share of votes received by the each of the two main candidates divided by total votes. I report Total votes, Voter turnout, and Vote share for centers without evidence of fraud. The remaining variables are defined in the "Data and Variables" section of the text. Year values in parenthesis indicate the year the data was collected. Variables without year indication were collected in 2009 or are time-invariant variables. Refer to section 3.2.2 for a description of the RD model used in columns (7) and (8).

Table B4: Sensitivity of Results to the Addition of Baseline Covariates

|  | All regions |  | South and East region |  | Northwest region |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Optimal Bandwidth (1) | Polynomial in Distance (2) | Optimal Bandwidth (3) | Polynomial in Distance (4) | Optimal Bandwidth (5) | Polynomial in Distance (6) |
| Panel A. At least one station with Category $C$ fraud |  |  |  |  |  |  |
| Inside coverage | $\begin{gathered} -0.081^{* *} \\ (0.032) \end{gathered}$ | $\begin{gathered} -0.080^{* * *} \\ (0.028) \end{gathered}$ | $\begin{gathered} -0.176^{* * *} \\ (0.056) \end{gathered}$ | $\begin{gathered} -0.167^{* * *} \\ (0.051) \end{gathered}$ | $\begin{gathered} 0.025 \\ (0.025) \end{gathered}$ | $\begin{gathered} 0.010 \\ (0.023) \end{gathered}$ |
| Observations | 1074 | 2039 | 532 | 1087 | 527 | 952 |
| Mean Outside coverage | 0.183 | 0.141 | 0.311 | 0.285 | 0.0350 | 0.0402 |
| Bandwidth (km) | 7.278 | - | 6.100 | - | 7.675 | - |
| Neighborhoods | 230 | 237 | 95 | 101 | 133 | 137 |
| Panel B. Share of votes under Category C fraud |  |  |  |  |  |  |
| Inside coverage | $\begin{gathered} -0.043^{* *} \\ (0.020) \end{gathered}$ | $\begin{gathered} -0.039^{* *} \\ (0.020) \end{gathered}$ | $\begin{gathered} -0.064 \\ (0.039) \end{gathered}$ | $\begin{gathered} -0.080^{* *} \\ (0.035) \end{gathered}$ | $\begin{gathered} 0.016 \\ (0.013) \end{gathered}$ | $\begin{gathered} 0.006 \\ (0.014) \end{gathered}$ |
| Observations | 1064 | 2039 | 528 | 1087 | 503 | 952 |
| Mean Outside coverage | 0.124 | 0.0933 | 0.221 | 0.198 | 0.0176 | 0.0193 |
| Bandwidth (km) | 7.152 | - | 5.963 | - | 7.030 | - |
| Neighborhoods | 228 | 237 | 95 | 101 | 132 | 137 |

Notes: Results use equation (3) and add a set of baseline covariates. Refer to section 3.2.1 for a description. Covariates used are: the number of stations, number of female stations, elevation, slope, population, distance to primary and secondary roads, distance to health facility, and distance to a primary river. Optimal bandwidth chosen as in Calonico et al. (2014). Polynomial order determined using Akaike's criterion as suggested in Black et al. (2007). All specifications use neighborhood fixed effects and standard errors clustered at the neighborhood level. Refer to section 3.3 .2 for a description of how boundary neighborhoods are created. ${ }^{* * *},{ }^{* *},{ }^{*}$ indicate 10,5 , and 1 percent significance, respectively.

Table B5: Effect of Mobile Coverage on Category C Fraud, Polling Center-level Clustering

|  | All regions |  | Southeast region |  | Northwest region |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Optimal Bandwidth (1) | Polynomial in Distance (2) | Optimal Bandwidth (3) | Polynomial in Distance <br> (4) | Optimal Bandwidth (5) | Polynomial in Distance (6) |
| Panel A. At least one station with Category C fraud |  |  |  |  |  |  |
| Inside coverage | $\begin{gathered} -0.077^{* * *} \\ (0.029) \end{gathered}$ | $\begin{gathered} -0.082^{* * *} \\ (0.024) \end{gathered}$ | $\begin{gathered} -0.160^{* * *} \\ (0.051) \end{gathered}$ | $\begin{gathered} -0.171^{* * *} \\ (0.042) \end{gathered}$ | $\begin{gathered} 0.027 \\ (0.023) \end{gathered}$ | $\begin{gathered} 0.011 \\ (0.022) \end{gathered}$ |
| Observations | 1074 | 2039 | 532 | 1087 | 527 | 952 |
| Mean Outside coverage | 0.183 | 0.141 | 0.311 | 0.285 | 0.035 | 0.040 |
| Bandwidth (km) | 7.278 | - | 6.100 | - | 7.675 | - |
| Panel B. Share of votes under Category C fraud |  |  |  |  |  |  |
| Inside coverage | $\begin{gathered} -0.039^{*} \\ (0.022) \end{gathered}$ | $\begin{gathered} -0.041^{* *} \\ (0.018) \end{gathered}$ | $\begin{aligned} & -0.067^{*} \\ & (0.040) \end{aligned}$ | $\begin{gathered} -0.094^{* * *} \\ (0.033) \end{gathered}$ | $\begin{gathered} 0.019 \\ (0.014) \end{gathered}$ | $\begin{gathered} 0.006 \\ (0.012) \end{gathered}$ |
| Observations | 1064 | 2039 | 528 | 1087 | 503 | 952 |
| Mean Outside coverage | 0.124 | 0.093 | 0.221 | 0.198 | 0.018 | 0.019 |
| Bandwidth (km) | 7.152 | - | 5.963 | - | 7.030 | - |

Notes: Results use equation (1). Refer to section II.B for a description. Optimal bandwidth chosen as in Calonico et al. (2014). Columns (2), (4), and (6) use a third degree polynomial in distance to boundary. Polynomial order determined using Akaike's criterion as suggested in Black et al. (2007). All specifications use neighborhood fixed effects and standard errors clustered at the polling center level. Refer to section II.C for a description of how boundary neighborhoods are created. ${ }^{* * *},{ }^{* *}, *$ indicate 10,5 , and 1 percent significance, respectively.

Table B6: Effect of Mobile Coverage on Category C Fraud, District Fixed Effects and Clustering

|  | All regions |  | Southeast region |  | Northwest region |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Optimal Bandwidth <br> (1) | Polynomial in Distance <br> (2) | Optimal Bandwidth (3) | Polynomial in Distance <br> (4) | Optimal Bandwidth (5) | Polynomial in Distance (6) |
| Panel A. At least one station with Category $C$ fraud |  |  |  |  |  |  |
| Inside coverage | $\begin{gathered} -0.042 \\ (0.034) \end{gathered}$ | $\begin{aligned} & -0.052 \\ & (0.032) \end{aligned}$ | $\begin{gathered} -0.116^{* *} \\ (0.057) \end{gathered}$ | $\begin{gathered} -0.144^{* *} \\ (0.061) \end{gathered}$ | $\begin{gathered} 0.029 \\ (0.023) \end{gathered}$ | $\begin{gathered} 0.020 \\ (0.022) \end{gathered}$ |
| Observations | 1083 | 2039 | 532 | 1087 | 538 | 952 |
| Mean Outside coverage | 0.182 | 0.141 | 0.311 | 0.285 | 0.035 | 0.040 |
| Bandwidth (km) | 7.373 | - | 6.139 | - | 8.101 | - |
| Districts | 188 | 239 | 87 | 119 | 95 | 120 |
| Panel B. Share of votes under Category C fraud |  |  |  |  |  |  |
| Inside coverage | $\begin{aligned} & -0.022 \\ & (0.021) \end{aligned}$ | $\begin{aligned} & -0.027 \\ & (0.019) \end{aligned}$ | $\begin{gathered} -0.059^{*} \\ (0.035) \end{gathered}$ | $\begin{gathered} -0.087^{* *} \\ (0.040) \end{gathered}$ | $\begin{gathered} 0.020 \\ (0.013) \end{gathered}$ | $\begin{gathered} 0.010 \\ (0.013) \end{gathered}$ |
| Observations | 1073 | 2039 | 529 | 1087 | 510 | 952 |
| Mean Outside coverage | 0.125 | 0.093 | 0.221 | 0.198 | 0.017 | 0.019 |
| Bandwidth (km) | 7.235 | - | 5.993 | - | 7.243 | - |
| Districts | 187 | 239 | 87 | 119 | 94 | 120 |

Notes: Results use equation (1). Refer to section II.B for a description. Optimal bandwidth chosen as in Calonico et al. (2014). Columns (2), (4), and (6) use a third degree polynomial in distance to boundary. Polynomial order determined using Akaike's criterion as suggested in Black et al. (2007). All specifications use district fixed effects and standard errors clustered at the district level. Refer to section II.C for a description of how boundary neighborhoods are created. ${ }^{* * *},{ }^{* *}$, * indicate 10,5 , and 1 percent significance, respectively.

Table B7: Effect of Mobile Coverage on Alternative Measures of Fraud

|  | All regions |  | Southeast region |  | Northwest region |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Optimal Bandwidth (1) | Polynomial in Distance <br> (2) | Optimal Bandwidth (3) | Polynomial in Distance <br> (4) | Optimal Bandwidth (5) | Polynomial in Distance <br> (6) |
| Panel A. At least one disqualified station within polling center |  |  |  |  |  |  |
| Inside coverage | $\begin{aligned} & -0.023 \\ & (0.017) \end{aligned}$ | $\begin{aligned} & -0.012 \\ & (0.014) \end{aligned}$ | $\begin{gathered} -0.060^{* *} \\ (0.027) \end{gathered}$ | $\begin{gathered} -0.049^{*} \\ (0.025) \end{gathered}$ | $\begin{gathered} 0.020 \\ (0.019) \end{gathered}$ | $\begin{gathered} 0.025 \\ (0.019) \end{gathered}$ |
| Observations | 1082 | 2039 | 545 | 1087 | 573 | 952 |
| Mean Outside coverage | 0.045 | 0.037 | 0.067 | 0.061 | 0.018 | 0.018 |
| Bandwidth (km) | 7.357 | - | 6.624 | - | 9.982 | - |
| Panel B. At least one station with more than 600 votes cast |  |  |  |  |  |  |
| Inside coverage | $\begin{aligned} & -0.062^{*} \\ & (0.036) \end{aligned}$ | $\begin{gathered} -0.050 \\ (0.036) \end{gathered}$ | $\begin{gathered} -0.096 \\ (0.063) \end{gathered}$ | $\begin{gathered} -0.106^{*} \\ (0.063) \end{gathered}$ | $\begin{gathered} 0.011 \\ (0.037) \end{gathered}$ | $\begin{aligned} & -0.025 \\ & (0.031) \end{aligned}$ |
| Observations | 1081 | 2039 | 484 | 1087 | 542 | 952 |
| Mean Outside coverage | 0.255 | 0.234 | 0.399 | 0.363 | 0.093 | 0.143 |
| Bandwidth (km) | 7.340 | - | 4.803 | - | 8.404 | - |

Notes: Results use equation (1). Refer to section II.B for a description. Panel title indicated outcome used. Optimal bandwidth chosen as in Calonico et al. (2014). Columns (2), (4), and (6) use a third degree polynomial in distance to boundary. Polynomial order determined using Akaike's criterion as suggested in Black et al. (2007). All specifications use neighborhood fixed effects and standard errors clustered at the polling center level. Refer to section II.C for a description of how boundary neighborhoods are created. ${ }^{* * *},{ }^{* *}, *$ indicate 10,5 , and 1 percent significance, respectively.

Table B8: Summary Statistics for Violence Data

|  | All regions combined |  | Southeast region |  | Northwest region |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Mean | S.D. | Mean | S.D. | Mean | S.D. |
| Panel A: Likelihood of at least one incident within 1 km of polling center |  |  |  |  |  |  |
| Attack | 0.213 | 0.41 | 0.316 | 0.46 | 0.068 | 0.25 |
| IED | 0.148 | 0.35 | 0.218 | 0.41 | 0.045 | 0.21 |
| Direct Fire | 0.103 | 0.30 | 0.128 | 0.33 | 0.030 | 0.17 |
| Indirect Fire | 0.065 | 0.25 | 0.100 | 0.30 | 0.011 | 0.10 |
| Panel B: Number of incidents within 1 km of polling center |  |  |  |  |  |  |
| Attack | 1.000 | 4.35 | 1.235 | 4.42 | 0.151 | 0.81 |
| IED | 0.418 | 1.63 | 0.493 | 1.47 | 0.081 | 0.44 |
| Direct Fire | 0.362 | 2.52 | 0.384 | 2.66 | 0.045 | 0.29 |
| Indirect Fire | 0.220 | 1.65 | 0.358 | 2.28 | 0.025 | 0.28 |
| Observations | 6,160 |  | 2,642 |  | 1,913 |  |

Notes: Data are significant actions (SIGACTs) collected by Afghan forces and ISAF. Refer to Condra et al. (2018) for detailed description of the data. Insurgent attacks data or significant actions (SIGACTs) are obtained from time-stamped and georeferenced records collected by the International Security Assistance Forces (ISAF) and Afghan forces in Afghanistan. The three main outcomes: the total number of attacks (Attack), the number of IEDs (IED), and direct fire attacks (Direct Fire) within a 1-kilometer radius of polling centers and from January 1, 2009 to August 20, 2009. This encompasses all violence occurring on the election year and up to election day. Direct fire refers to attacks using small arms and rocket-propelled grenades.

Table B9: Summary Statistics for Complaints Data

|  | Mean <br> (1) | S.D. <br> (2) | Coverage |  | No coverage |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Mean (3) | S.D. <br> (4) | Mean (5) | S.D. <br> (6) |
| Number of complaints | 0.54 | 1.30 | 0.62 | 1.37 | 0.33 | 1.03 |
| Share of complaints submitted by individuals | 0.91 | 0.26 | 0.91 | 0.26 | 0.91 | 0.25 |
| Share of complaints submitted by females | 0.13 | 0.28 | 0.13 | 0.29 | 0.08 | 0.23 |
| Observations | 2785 |  | 2071 |  | 714 |  |

Notes: Individuals refers to citizens, candidates, or candidates' representatives and excludes government officials (polling officials, IEC or ECC officials, and other government organizations). Refer to section 4.3 for more details.

Table B10: Tribes and Tribal Confederations of Southeast Afghanistan

|  | Number | Share of <br> Total | Share of <br> ethnic <br> group | Share of <br> confederacy |
| :--- | :---: | :---: | :---: | :---: |
| Panel A: Pashtun Confederations and Tribes |  |  |  |  |
| Durrani confederation | 4,343 | 0.24 | 0.37 | 1.00 |
| Durrani Pashtun | 1,608 | 0.09 | 0.14 | 0.37 |
| Barakzai Pashtun | 902 | 0.05 | 0.08 | 0.21 |
| Noorzai Pashtun | 502 | 0.03 | 0.04 | 0.12 |
| Panjpai Durrani Pashtun | 401 | 0.02 | 0.03 | 0.09 |
| Alizai Pashtun | 328 | 0.02 | 0.03 | 0.08 |
| Mixed Durrani | 208 | 0.01 | 0.02 | 0.05 |
| Popalzai Pashtun | 200 | 0.01 | 0.02 | 0.05 |
| Ashakzai Pashtun | 194 | 0.01 | 0.02 | 0.04 |
| Ghilzai confederation | 4,469 | 0.25 | 0.38 | 1.00 |
| Ibrahim Ghilzai | 3,325 | 0.18 | 0.28 | 0.74 |
| Turan Pashtun | 874 | 0.05 | 0.07 | 0.20 |
| Baezai Mohamand Powindah | 110 | 0.01 | 0.01 | 0.02 |
| Kudi Lodi Powindah | 82 | 0.00 | 0.01 | 0.02 |
| Mian Khel Powindah | 39 | 0.00 | 0.00 | 0.01 |
| Other Ghilzai | 16 | 0.00 | 0.00 | 0.00 |
| Miani Powindah | 14 | 0.00 | 0.00 | 0.00 |
| Kukozai Mohamand Powindah | 9 | 0.00 | 0.00 | 0.00 |
| Other Pashtun | 3,052 | 0.17 | 0.26 | 1.00 |
| Safi Pashtun | 1,261 | 0.07 | 0.11 | 0.41 |
| Shinwari Pashtun | 509 | 0.03 | 0.04 | 0.17 |
| Pashtun | 279 | 0.02 | 0.02 | 0.09 |
| Wardak Pashtun | 235 | 0.01 | 0.02 | 0.08 |
| Salarzai | 113 | 0.01 | 0.01 | 0.04 |
| Kom | 107 | 0.01 | 0.01 | 0.04 |
| Jadran Pashtun | 87 | 0.00 | 0.01 | 0.03 |
| Other | 461 | 0.02 | 0.04 | 0.14 |
| All Pashtuns | 11,864 | 0.66 | 1.00 | . |
|  |  |  |  | $($ Continues) |
|  |  |  |  |  |

Table B10: Tribes and Tribal Confederations of Southeast Afghanistan-Continued

|  | Number | Share of <br> Total | Share of <br> ethnic <br> group | Share of <br> confederacy |
| :--- | :---: | :---: | :---: | :---: |
| Panel B: Hazara Tribes | 1,226 | 0.07 |  |  |
| Besud Hazara | 896 | 0.05 | 0.29 | . |
| Other Hazara | 615 | 0.03 | 0.14 | . |
| Dai Chopan Hazara | 309 | 0.02 | 0.07 | . |
| Dai Khitai Hazara | 222 | 0.01 | 0.05 | . |
| Muhammad Kwaja Hazara | 215 | 0.01 | 0.05 | . |
| Khatai Hazara | 187 | 0.01 | 0.04 | . |
| Jaghuri Hazara | 154 | 0.01 | 0.04 | . |
| Dai Kundi Hazara | 134 | 0.01 | 0.03 | . |
| Chahar Dasta Hazara | 134 | 0.01 | 0.03 | . |
| Polada Hazara | 94 | 0.01 | 0.02 | . |
| Jaghatus Hazara | 55 | 0.00 | 0.01 | . |
| Faoladi Hazara | 21 | 0.00 | 0.00 | . |
| Dai Zangi Hazara | 13 | 0.00 | 0.00 | . |
| Uruzgani Hazara | 4,275 | 0.24 | 1.00 | . |
| All Hazara |  |  |  |  |
| Panel C: Other Ethnic Groups | 1,117 | 0.06 | 1.00 | . |
| Tajik | 174 | 0.01 | 1.00 | . |
| Uzbek | 157 | 0.01 | 1.00 | . |
| Baluch | 0.01 | 1.00 | . |  |
| Mixed | 0.01 | 1.00 | . |  |
| Other (non-Pashtun) | 246 |  |  |  |

Notes: Tribal confederations created using definitions based on Tribal Hierarchy and Dictionary of Afghanistan (2007). Tribal maps obtained via the Culture and Conflict Studies.

Table B11: Effect of Mobile Coverage on Number of Insurgent Attacks near Polling Centers: Excluding Election Day

|  | All regions |  | Southeast region |  |
| :---: | :---: | :---: | :---: | :---: |
|  | Optimal Bandwidth <br> (1) | Polynomial in Distance <br> (2) | Optimal Bandwidth <br> (3) | Polynomial in Distance <br> (4) |
| Panel A. Number of violent incidents within 1 km radius of poling center |  |  |  |  |
| Inside coverage | 0.575 | 0.724 | 0.517 | 1.064 |
|  | (0.599) | (0.478) | (1.029) | (0.883) |
| Observations | 1278 | 2118 | 682 | 1163 |
| Mean Outside coverage | 0.81 | 0.66 | 1.49 | 1.50 |
| Bandwidth (km) | 10.2 | - | 9.53 | - |
| Neighborhoods | 236 | 239 | 101 | 102 |
| Panel B. Number of IEDs within 1 km radius of polling center |  |  |  |  |
| Inside coverage | 0.228 | 0.178 | 0.367 | 0.276 |
|  | (0.196) | (0.161) | (0.313) | (0.299) |
| Observations | 866 | 2118 | 528 | 1163 |
| Mean Outside coverage | 0.14 | 0.076 | 0.23 | 0.17 |
| Bandwidth (km) | 4.20 | - | 4.86 | - |
| Neighborhoods | 222 | 239 | 96 | 102 |
| Panel C. Number of direct fire incidents within 1 km radius of polling center |  |  |  |  |
| Inside coverage | 0.345 | 0.412 | 0.368 | 0.590 |
|  | (0.313) | (0.268) | (0.549) | (0.494) |
| Observations | 1473 | 2118 | 817 | 1163 |
| Mean Outside coverage | 0.43 | 0.36 | 0.83 | 0.82 |
| Bandwidth (km) | 13.6 | - | 13.4 | - |
| Neighborhoods | 238 | 239 | 102 | 102 |

Notes: Results use equation (1) and outcome variables denoted in panel names. Refer to section III.B for more details on the outcome variables used. Optimal bandwidth chosen as in Calonico et al. (2014). Columns (2) and (4) use a second degree polynomial in distance to boundary. Polynomial order determined using Akaike's criterion as suggested in Black et al. (2007). All specifications use neighborhood fixed effects and standard errors clustered at the neighborhood level. Refer to section II.C for a description of how boundary neighborhoods are created. ${ }^{* * *},{ }^{* *}$, $*$ indicate 10,5 , and 1 percent significance, respectively.

Table B12: Effect of Mobile Coverage on Number of Insurgent Attacks near Polling Centers: Entire time period (January 1, 2006-August 20, 2009 (Election Day))

|  | All regions |  | Southeast region |  |
| :---: | :---: | :---: | :---: | :---: |
|  | Optimal Bandwidth (1) | Polynomial in Distance <br> (2) | Optimal Bandwidth (3) | Polynomial in Distance <br> (4) |
| Panel A. Number of violent incidents within 1 km radius of poling center |  |  |  |  |
| Inside coverage | 3.565 | 3.482 | 5.024 | 5.607 |
|  | (3.171) | (2.784) | (5.492) | (5.240) |
| Observations | 1339 | 2118 | 712 | 1163 |
| Mean Outside coverage | 2.46 | 1.93 | 4.59 | 4.35 |
| Bandwidth (km) | 11.2 | - | 10.6 | - |
| Neighborhoods | 236 | 239 | 101 | 102 |
| Panel B. Number of IEDs within 1 km radius of polling center |  |  |  |  |
| Inside coverage | 0.721 | 0.326 | 1.056 | 0.402 |
|  | (0.522) | (0.485) | (0.877) | (0.899) |
| Observations | 900 | 2118 | 540 | 1163 |
| Mean Outside coverage | 0.62 | 0.36 | 1.05 | 0.81 |
| Bandwidth (km) | 4.55 | - | 5.14 | - |
| Neighborhoods | 222 | 239 | 97 | 102 |
| Panel C. Number of direct fire incidents within 1 km radius of polling center |  |  |  |  |
| Inside coverage | 2.161 | 2.127 | 3.008 | 3.400 |
|  | (1.669) | (1.534) | (2.998) | (2.878) |
| Observations | 1533 | 2118 | 774 | 1163 |
| Mean Outside coverage | 1.11 | 0.92 | 2.18 | 2.06 |
| Bandwidth (km) | 14.5 | - | 12.5 | - |
| Neighborhoods | 238 | 239 | 102 | 102 |

Notes: Results use equation (1) and outcome variables denoted in panel names. Refer to section III.B for more details on the outcome variables used. Optimal bandwidth chosen as in Calonico et al. (2014). Columns (2) and (4) use a second degree polynomial in distance to boundary. Polynomial order determined using Akaike's criterion as suggested in Black et al. (2007). All specifications use neighborhood fixed effects and standard errors clustered at the neighborhood level. Refer to section II.C for a description of how boundary neighborhoods are created. ${ }^{* * *},{ }^{* *}$, $*$ indicate 10,5 , and 1 percent significance, respectively.

Table B13: Effect of Mobile Coverage on Number of Insurgent Attacks near Polling Centers: 5KM Radius

|  | All regions |  | Southeast region |  |
| :---: | :---: | :---: | :---: | :---: |
|  | Optimal Bandwidth <br> (1) | Polynomial in Distance <br> (2) | Optimal Bandwidth (3) | Polynomial in Distance <br> (4) |
| Panel A. Number of violent incidents within 5 km radius of poling center |  |  |  |  |
| Inside coverage | 1.914 | 2.426 | 3.193 | 5.287 |
|  | (2.696) | (2.228) | (4.635) | (4.150) |
| Observations | 1284 | 2118 | 756 | 1163 |
| Mean Outside coverage | 7.99 | 6.09 | 14.6 | 13.6 |
| Bandwidth (km) | 10.3 | - | 11.8 | - |
| Neighborhoods | 236 | 239 | 101 | 102 |
| Panel B. Number of IEDs within 5km radius of polling center |  |  |  |  |
| Inside coverage | 1.796 | 0.719 | 3.197* | 1.996 |
|  | (1.108) | (1.099) | (1.860) | (2.004) |
| Observations | 844 | 2118 | 519 | 1163 |
| Mean Outside coverage | 1.90 | 1.12 | 3.23 | 2.53 |
| Bandwidth (km) | 3.98 | - | 4.69 | - |
| Neighborhoods | 222 | 239 | 96 | 102 |
| Panel C. Number of direct fire incidents within 5km radius of polling center |  |  |  |  |
| Inside coverage | 0.322 | 1.343 | 0.172 | 2.186 |
|  | (1.340) | (0.930) | (2.510) | (1.847) |
| Observations | 1433 | 2118 | 860 | 1163 |
| Mean Outside coverage | 4.49 | 3.57 | 8.39 | 7.96 |
| Bandwidth (km) | 13.1 | - | 14.3 | - |
| Neighborhoods | 238 | 239 | 102 | 102 |

Notes: Results use equation (1) and outcome variables denoted in panel names. Refer to section III.B for more details on the outcome variables used. Optimal bandwidth chosen as in Calonico et al. (2014). Columns (2) and (4) use a second degree polynomial in distance to boundary. Polynomial order determined using Akaike's criterion as suggested in Black et al. (2007). All specifications use neighborhood fixed effects and standard errors clustered at the neighborhood level. Refer to section II.C for a description of how boundary neighborhoods are created. ${ }^{* * *},{ }^{* *}$, $*$ indicate 10,5 , and 1 percent significance, respectively.

Table B14: Summary of News Articles Addressing Tower Shutdowns

|  | Full Sample <br> (Feb. 2008-Dec. 2019) |  |  | Up to 2009 Election <br> (Feb. 2008-Dec. 2009) |  |
| :--- | :---: | :---: | :--- | :---: | :---: |
|  | Number | Percent |  | Number | percent |
| Panel A: References to Timing of Shutdowns |  |  |  |  |  |
| Night | 29 | 93.5 |  | 10 | 90.9 |
| Day | 1 | 3.2 |  | 1 | 9.1 |
| Complete shutdown | 1 | 3.2 |  | 0 | 0.0 |
| Total articles (timing reference) | 31 | 100 |  | 11 | 100 |
| Timing not specified | 10 | - |  | 0 | - |
| Panel B: References to Location of Shutdowns |  |  |  |  |  |
| Helmand | 9 | 26.5 |  | 1 | 14.3 |
| Kandahar | 3 | 8.8 |  | 1 | 14.3 |
| Ghazni | 5 | 14.7 |  | 1 | 14.3 |
| Southern Afghanistan | 8 | 23.5 |  | 4 | 57.1 |
| Other | 9 | 26.5 |  | 0 | 0.0 |
| Total articles (location reference) | 34 | 100 |  | 7 | 100 |
| Location not specified | 7 | - |  | 4 | - |
| Total articles |  | 41 |  |  | 11 |

Notes: Night/day refers to whether the article specified if the shutdowns were during the night or during the day. References to locations refers to whether the article mentioned the location of the shutdowns. Total is the number of articles. Shares calculated out of the total articles making reference to timing or location of the shutdowns. Other provinces referenced in the location of shutdowns where: Balkh, Farah, Zabul, Paktika, Ghor, Wardak, Herat, Kunduz,Baghlan, Nangharar, Takhar, Khost, Paktya, Logar, and Kunar.

## Appendix C Model Extensions

## C. 1 The Transmitter's Problem

Consider a polling center serving $n$ voters. Furthermore assume that there exists a nationwide phone hotline to report electoral fraud. Given the widespread use of cell phones in the developing world and Afghanistan in particular, suppose individual $i$ uses a cell phone if he decides to report fraud. Reporting fraud carries a physical cost $c(D)$ where $D$ indicates the accessibility of the medium (cell phones) used to report fraud. In the context of this study $D$ is an indicator for whether the polling center is located in an area with cell phone coverage. ${ }^{2}$ Specifically, let the cost of reporting fraud equal $\underline{c}$ if the center is on an area with cell phone coverage (i.e., $D=1$ ) and $\bar{c}$ otherwise with $\bar{c}>\underline{c}^{3}{ }^{3}$

Furthermore, assume that reporting fraud gives $i$ a utility gain $\lambda_{i}$ that can be interpreted as a "warm glow" parameter or $i$ 's satisfaction from his pro-social behavior. The individual's net payoff from reporting fraud is therefore given by $\lambda_{i}-c(D)$. He will then decide to report fraud if:

$$
\begin{equation*}
\lambda_{i} \geq c(D) \tag{1}
\end{equation*}
$$

Assuming $\lambda_{i}$ is distributed among voters at the center with probability function $G(\lambda)$ then the probability of an individual making a report given coverage status $D$ is given by $\rho(D)=1-G(D) .{ }^{4}$

## C. 2 Nonlinear Reporting Cost Function

Since reporting fraud via cell phones increases in cost as polling centers are farther away from coverage areas, I consider the case where this function is nonlinear. Specifically, assume reporting fraud carries a physical cost $c(d)$ where $d$ defines how accessible the medium (cell phones) used to report fraud is. In the spatial context of this study $d$ can be interpreted as the shortest distance from the polling center where fraud takes place to a geographic boundary that defines whether

[^2]there is cell phone coverage. Using this interpretation and assuming that $d<0$ (i.e., a negative distance) arbitrarily defines a polling center located in the non-coverage side of the boundary one can specify a function $D=\mathbb{1}\{d \geq 0\}$ that indicates coverage such that the cost of reporting fraud is given by:
\[

$$
\begin{equation*}
c(d)=D \underline{c}+(1-D) h(d) \tag{2}
\end{equation*}
$$

\]

where $\underline{c}$ is the marginal cost of making a call when the center is in the coverage side. I assume this cost is equal for everyone in the coverage side. $h(d)$ is a smooth cost function faced by individuals on the non-coverage side with $h(d)>\underline{c}$ for all $d<0$ and $h(0)=\bar{c}$ with $\bar{c}>\underline{c} .^{5}$

## C. 3 Free-riding and Fraud Reporting

I extend the transmitter's problem presented in section C. 1 by allowing for the possibility of "freeriding" in the reporting process. Free-riding can be a concern in this context if individuals assess that the probability that a fraudulent official is punished conditional on his report is trivial. If that is the case, then the probability of making a report does not change regardless of accessibility to the reporting medium, in this context, coverage status. I show that individuals have an incentive to report fraud, despite the free-riding problem, as long as there is some utility gain from the reporting process itself (i.e., the warm glow parameter specified in section C.1)

More specifically, assume that since reporting fraud is costly, individual $i$ assesses the likelihood that the center is actually audited (and hence the fraudulent candidate is penalized) as a result of his report. ${ }^{6}$ More specifically, let $\phi_{1}$ and $\phi_{0}$, with $\phi_{1} \geq \phi_{0}$ denote $i$ 's subjective assessment of the probability that the candidate will be punished given that he reports him and does not report him respectively. When the candidate is punished $i$ gets a net utility value $\xi_{i}$ that can be interpreted as a utility gain from the fraudulent candidate being punished net of any affinity or benefits that the individual might receive from non-punishment. ${ }^{7}$ Additionally, assume as in section C.1, that the act of reporting fraud gives $i$ a utility gain $\lambda_{i}$.

The individual's net payoffs from reporting fraud are therefore given by $\xi_{i}+\lambda_{i}-c(d)$ when the fraudulent candidate is punished and $\lambda_{i}-c(d)$ otherwise. Lastly, if $i$ decides not to report, he simply obtains $\xi_{i}$ when the candidate is punished and zero otherwise. Assuming linear utility, he

[^3]will then decide to report fraud if:
\[

$$
\begin{align*}
\phi_{1}\left[\xi_{i}+\lambda_{i}-c(d)\right]+\left(1-\phi_{1}\right)\left[\lambda_{i}-c(d)\right] & \geq \phi_{0} \xi_{i} \\
\lambda_{i}+\left(\phi_{1}-\phi_{0}\right) \xi_{i} & \geq c(d) \tag{3}
\end{align*}
$$
\]

Notice that even when there is a "free-riding" problem (i.e., $i$ believes that his report does not affect the probability that the candidate is punished $\left(\phi_{1}=\phi_{0}\right)$ and hence the decision rule above reduces to: Report if $\lambda_{i} \geq c(d)$ ) an individual $i$ might still have an incentive to report fraud as he derives utility from doing this alone. Therefore "free-riding" will lower the willingness to report fraud but not eliminate it completely.

## C. 4 The Voter's Problem

An individual $i$ considers a campaign promise and the possibility of a violent outcome on election day when deciding whether to vote. In the spirit of Dekel et al. (2008), assume each voter is characterized by parameter $a_{i}$ representing the net utility the individual obtains from simply tendering his vote to the candidate. This can be interpreted as $i$ 's affinity towards the candidate. The candidate offers individual $i$ a campaign promise $p_{l, i}$ regardless of the election result but conditional on $i$ tendering the vote to him. Assuming the individual is an expected income maximizer then $i$ 's expected payoff from voting is given by $V_{i}^{1}=a_{i}+p_{l, i}$ while $i$ 's expected payoff from not voting is simply $V_{i}^{0}=0$. Since elections in conflict zones are often characterized by violence, the individual takes into account an exogenous probability $\delta_{j}$ that a violent event takes place at polling center $j$ and as a result receives a negative payoff $P$. This consideration is particularly important in the Afghan context as the Taliban issued several warnings targeting polling centers and voters on election day (Gall (2009), Filkins (2009)). The individual therefore decides to vote if $\delta\left(V_{i}^{1}-P\right)+(1-\delta) V_{i}^{1} \geq V_{i}^{0} .{ }^{8}$

Given $i$ 's payoffs, the minimum price per legal vote (i.e., the campaign promise) that guarantees $i$ 's vote is therefore given by:

$$
\begin{equation*}
p_{l, i}=\delta_{j} P-a_{i} \tag{4}
\end{equation*}
$$

where $p_{l, i}=0$ if the affinity parameter $a_{i}$ is sufficiently large as to offset the negative payoff of violence (i.e., $\left.a_{i}>\delta_{j} P\right) .{ }^{9}$

[^4]
## C. 5 Two Candidates

This section introduces the possibility of two candidates in the voter's problem. A key distinction from the model presented in section C. 4 is that the voter not only has to decide whether to vote but also for whom to vote taking into account each candidate's campaign promise. In terms of the candidate's problem, the level of fraud in equilibrium changes slightly when considering a second candidate. In essence, this introduces an additional channel of fraud, namely what I refer to as a challenger effect. Broadly speaking, in order for a candidate to entice voters to vote for him, he has to pay a legal price for their vote that matches the highest value between the expected net payoff from violence and the opposing candidates' campaign offer to voters.

The Voter's Problem Suppose voters have to decide between two candidates indexed by $k$. As in Dekel et al. (2008), assume that each voter $i$ is characterized by parameters $U_{i}^{k}$ and $a_{i}^{k}$ that represent the utility the individual obtains from $k$ 's victory and from simply tendering his vote to $k$, respectively. Candidate $k$ offers individual $i$ a campaign promise $p_{l, i}^{k}$ regardless of the election result but conditional on $i$ tendering his vote to $k^{10}$. Letting $\psi_{k \mid l}$ denote the probability that candidate $k$ wins given that $i$ tendered his vote to candidate $l \in\{X, Y\}$ and letting the individual be an expected income maximizer then $i$ 's expected payoff from voting for $k$ is given by:

$$
\begin{equation*}
V_{i}^{k}=\psi_{k \mid k}\left(U_{i}^{k}+a_{i}^{k}+p_{l, i}^{k}\right)+\psi_{\sim k \mid k}\left(U_{i}^{\sim k}+a_{i}^{k}+p_{l, i}^{k}\right) \tag{5}
\end{equation*}
$$

where $\sim k$ denotes "not" $k$. Similarly, letting $\psi_{k \mid 0}$ denote the probability that candidate $k$ wins given that $i$ did not vote, then $i$ 's expected payoff from not voting is simply:

$$
\begin{equation*}
V_{i}^{0}=\psi_{k \mid 0} U_{i}^{k}+\psi_{\sim k \mid 0} U_{i}^{\sim k} \tag{6}
\end{equation*}
$$

Since elections in conflict zones are often characterized by violence, the individual takes into account an exogenous probability $\delta_{j}$ that a violent event takes place at the polling center and as a result receives a very negative payoff $P$. The individual therefore decides to vote if:

$$
\begin{equation*}
\delta\left(V_{i}^{k}-P\right)+(1-\delta) V_{i}^{k} \geq V_{i}^{0} \tag{7}
\end{equation*}
$$

[^5]Given expressions (5), (6) and (7), the individual will decide to vote and vote for $k$ if the following two conditions hold.

$$
\begin{align*}
p_{l, i}^{k}+a_{i}^{k}+\psi_{k \mid k} U_{i}^{k}+\psi_{\sim k \mid k} U^{\sim k} & \geq \delta P+\psi_{k \mid 0} U_{i}^{k}+\psi_{\sim k \mid 0} U^{\sim k}  \tag{8}\\
p_{l, i}^{k}+a_{i}^{k}+\psi_{k \mid k} U_{i}^{k}+\psi_{\sim k \mid k} U^{\sim k} & \geq p_{l, i}^{\sim k}+\psi_{k \mid \sim k} U_{i}^{k}+\psi_{\sim k \mid \sim k} U^{\sim k} \tag{9}
\end{align*}
$$

Assuming, for simplicity, that the individual believes that his vote is "non-pivotal" 11 , then the two expressions above simply reduce to:

$$
\begin{align*}
p_{l, i}^{k}+a_{i}^{k} & \geq \delta P  \tag{10}\\
p_{l, i}^{k}+a_{i}^{k} & \geq p_{l, i}^{\sim k}+a_{i}^{\sim k} \tag{11}
\end{align*}
$$

The minimum price per legal vote (i.e., the campaign promise) that candidate $k$ must pay is therefore given by:

$$
\begin{equation*}
p_{l, i}^{k}=\max \left\{0, \delta P-a_{i}^{k}, p_{l, i}^{\sim k}+a_{i}^{\sim k}-a_{i}^{k}\right\} \tag{12}
\end{equation*}
$$

The Candidate's Problem Candidate $k$ must decide how many votes (both legal and fraudulent) to buy from each center $j$. Assume that the auditing agency can differentiate between fraudulent and legal votes so that, once audited, any fraudulent votes are dropped and the candidate only receives legal votes $v_{l}^{k}$. In case where the center is not audited the candidate simply keeps all votes $v_{l}^{k}+v_{f}^{k}$. Given that the assessed probability of an audit is given by $\pi$ and assuming that the candidate has quasilinear preferences over votes, then the maximization problem of the candidate is given by:

$$
\begin{array}{ll}
\max _{v_{l}^{k}, v_{f}^{k}} & \pi v_{l}^{k}+(1-\pi)\left[v_{l}^{k}+\left(v_{f}^{k}\right)^{\alpha}\right] \\
\text { subject to } & p_{f} v_{f}^{k}+\sum_{i=1}^{v_{l}^{k}} p_{l, i}^{k} \leq E^{k}
\end{array}
$$

where fraudulent votes enter non-linearly (with $\alpha \leq 1$ ) to capture the possibility that fraudulent and legal votes are not perfect substitutes as specified in Callen and Weidmann (2013) and $E^{k}$ is some campaign endowment of candidate $k$. To simplify the analysis, assume that the candidate does not observe the affinity parameters $a_{i}^{k}$ but knows their distribution among the voters in center $j$ so

[^6]that $\sum_{i=1}^{v_{l}^{k}} p_{l, i}^{k}$ is simply given by $p_{l}^{k} v_{l}^{k}$ where $p_{l}^{k}$ uses the expected value of these affinity parameters. The solution to the problem above provides an optimal relationship between fraudulent votes and their price $p_{f}$ that is given by:
\[

$$
\begin{equation*}
v_{f}^{k}=\left[\frac{\alpha(1-\pi) p_{l}^{k}}{p_{f}}\right]^{\frac{1}{1-\alpha}} \tag{13}
\end{equation*}
$$

\]

Substituting the expressions for prices $p_{f}$ and $p_{l}^{k}$ in order to obtain the equilibrium level of fraud gives:

$$
\begin{equation*}
v_{f}^{k}=\left[\frac{\alpha(1-\pi) \cdot \max \left\{0, \delta P-a^{k}, p_{l}^{\sim k}+a^{\sim k}-a^{k}\right\}}{\pi F}\right]^{\frac{1}{1-\alpha}} \tag{14}
\end{equation*}
$$

From this expression it is clear that the probability $\pi$ that the center is audited (i.e., the level of social monitoring) decreases the equilibrium fraud level in the center, however, notice also that the "social monitoring effect" is one among others that explain fraud. To see this more clearly, I rewrite expression (14) by separating the different components of fraud:

$$
v_{f}^{k}=[\alpha \cdot \underbrace{\frac{1-\pi}{\pi}}_{\begin{array}{c}
\text { Social }  \tag{15}\\
\text { monitoring } \\
\text { effect }
\end{array}} \cdot \underbrace{\frac{1}{F}}_{\begin{array}{c}
\text { Fine } \\
\text { effect }
\end{array}} \cdot \max \{0, \underbrace{\delta P-a^{k}}_{\begin{array}{c}
\text { Violence } \\
\text { effect }
\end{array}}, \underbrace{p_{l}^{\sim k}+a^{\sim k}-a^{k}}_{\begin{array}{c}
\text { Challenger } \\
\text { effect }
\end{array}}\}]^{\frac{1}{1-\alpha}}
$$

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[^1]:    ${ }^{1}$ The misunderstanding was mainly due to the definition used to classify votes as "valid".

[^2]:    ${ }^{2}$ In reality $D$ should indicate whether there is coverage in the area where the individual decides to report fraud (the polling center, his house, etc.). This information is unavailable hence I only consider coverage at the polling center. This implicitly assumes that the call to report fraud is made right at the center. However, since centers were located within settlements, it is likely that any calls are made within the "catchment area" of the center so that using the center as a reference in determining coverage should not greatly affect the analysis of the model.
    ${ }^{3}$ Without loss of generality, I assume that the reporting cost on the non-coverage side $\bar{c}$ is constant, however, this cost might increase as polling center are further away from the coverage boundary. Refer to Appendix C for a discussion of an alternative specification of the reporting cost function that uses a smooth, non-linear function on the non-coverage side.
    ${ }^{4} G($.$) is actually a function of c(D)$, which in turn, is a function of $D$. Refer to Appendix C for an extension of Equation 1 that considers the possibility of free-riding when reporting fraud.

[^3]:    ${ }^{5}$ A possible parameterization for $c(d)$ could be $c(d)=\bar{c} \cdot \exp (-\beta d)$
    ${ }^{6}$ I assume that once a fraudulent center is audited, the candidate and polling center manager are penalized. Therefore I do not consider any "concealment technology" as in Cremer and Gahvari (1994)
    ${ }^{7}$ The idea is that the individual might obtain some "justice has been served" satisfaction while at the same time punishment to a candidate of his liking might bring some disutility.

[^4]:    ${ }^{8}$ For simplicity, this specification of the model considers only the problem of whether to vote but not the problem of for whom to vote. Refer to Appendix C for an extension of the "Voter's Problem" that considers two candidates and hence the decision becomes whether to vote and for whom to vote.
    ${ }^{9}$ Alternatively, one can specify the reservation campaign promise $p_{l, i}$ as equal to the $\max \left\{0, \delta_{j} P-a_{i}\right\}$

[^5]:    ${ }^{10}$ This campaign promise can be interpreted in either of two ways: Voting is not secret so that candidate $k$ knows which individuals voted for him and hence he pays them the campaign promise $p_{l, i}^{k}$, or voting is secret but once a voter commits a priori to tender the vote to $k$ he does not change his vote the day of the election.

[^6]:    ${ }^{11}$ This assumption simply states that the individual believes that tendering the vote to $k$ will not alter the probability that $k$ wins. More specifically, $\psi_{k \mid k}=\psi_{k \mid \sim k}=\psi_{k \mid 0}$ for $k \in\{X, Y\}$.

